

Point Process Theory for Bilinear and Stochastic Volatility Models

by

Richard A. Davis, CSU

Sidney Resnick, Cornell

Jay Breidt, Iowa State U

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II. Point process theory for stochastic volatility models

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I. Point process theory for bilinear models

The Model

Linear model: $X_t = \sum_{k=0}^{\infty} c_k Z_{t-k}$, where $\{Z_t\}$ is an IID sequence.

Simple first-order bilinear model:

$$X_t = cZ_{t-1}X_{t-1} + Z_t$$

where $\{Z_t\}$ is an IID sequence.

Note:

$$\begin{aligned} X_t &= cZ_{t-1}X_{t-1} + Z_t \\ &= Z_t + cZ_{t-1}^2 + c^2Z_{t-1}Z_{t-2}X_{t-2} \\ &= Z_t + cZ_{t-1}^2 + c^2Z_{t-1}Z_{t-2}^2 + c^3Z_{t-1}Z_{t-2}Z_{t-3}X_{t-3} \\ &= \sum_{k=0}^{\infty} c^k Y_t^{(k)} \end{aligned}$$

where

$$Y_t^{(0)} = Z_t$$

$$Y_t^{(1)} = Z_{t-1}^2$$

$$Y_t^{(k)} = \left(\prod_{i=1}^{k-1} Z_{t-i} \right) Z_{t-k}^2, \text{ for } k > 0.$$

Tail Behavior of X_1

Assume $\{Z_t\}$ is an IID sequence of non-negative rv's with cdf

$$1-F(x) = x^{-\alpha}L(x)$$

where $L(x)$ is a slowly varying function at infinity.

Proposition:

$$\frac{P[Y_t^{(k)} > x]}{P[Z_1^2 > x]} \rightarrow \left(EZ_1^{\alpha/2} \right)^{k-1}$$

and

$$\frac{P[Y_t^{(k)} > x, Y_t^{(j)} > x]}{P[Z_1^2 > x]} \rightarrow 0 \quad \text{as } x \rightarrow \infty.$$

Corollary: If $c^{\alpha/2} \mathbf{E} Z_1^{\alpha/2} < 1$, then

$$\mathbf{P}\left[\sum_{k=0}^m c^k Y_t^{(k)} > x\right] \approx \sum_{k=1}^m c^{k\alpha/2} \left(\mathbf{E} Z_1^{\alpha/2}\right)^{k-1} \mathbf{P}[Z_1^2 > x]$$

and

$$\begin{aligned} \lim_{x \rightarrow \infty} \frac{\mathbf{P}\left[\sum_{k=0}^{\infty} c^k Y_t^{(k)} > x\right]}{\mathbf{P}[Z_1^2 > x]} &= \sum_{k=1}^{\infty} c^{k\alpha/2} \left(\mathbf{E} Z_1^{\alpha/2}\right)^{k-1} \\ &= \frac{c^{\alpha/2}}{1 - c^{\alpha/2} \mathbf{E} Z_1^{\alpha/2}}. \end{aligned}$$

Point process convergence

Assumptions: $X_t = cZ_{t-1}X_{t-1} + Z_t$, $\{Z_t\}$ is IID

- $P[|Z_t| > x] = x^{-\alpha}L(x)$, $P[Z_t > x] / P[|Z_t| > x] \rightarrow p$.

$$X_t = \sum_{k=0}^{\infty} c^k Y_t^{(k)}, \quad Y_t^{(1)} = Z_{t-1}, \quad Y_t^{(k)} = \left(\prod_{i=1}^{k-1} Z_{t-i} \right) Z_{t-k}$$

- $b_n = \inf\{x: P[|Z_t| > x] < n^{-1}\}$
- Define the sequence of point processes

$$N_n(\cdot) = \sum_{t=1}^n \mathcal{E}_{b_n^{-2}X_t}(\cdot)$$

Theorem:

$$(i) \quad N_n(\cdot) \xrightarrow{d} \sum_{k=1}^{\infty} \sum_{s=1}^{\infty} \varepsilon_{j_s^2 c^k W_{s,k}}(\cdot),$$

where $\{j_s\}$ are points of a Poisson process with intensity measure $\mu(dx) = \alpha(p x^{-\alpha-1} 1[x > 0] + (1-p)(-x)^{-\alpha-1} 1[x < 0])dx$, and

$$W_{s,k} = \begin{cases} \prod_{i=1}^{k-1} U_{s,i}, & \text{if } k > 1, \\ 1, & \text{if } k = 1, \\ 0, & \text{if } k < 1, \end{cases}$$

with $\{U_{s,k}\}$ IID (F).

$$(ii) \sum_{t=1}^n \mathcal{E}_{b_n^{-2}}(X_t, X_{t-1}, \dots, X_{t-h}) (\cdot)$$

$$\xrightarrow{d} \sum_{k=1}^{\infty} \sum_{s=1}^{\infty} \mathcal{E}_{j_s^2}(c^k W_{s,k}, c^{k-1} W_{s,k-1}, \dots, c^{k-h} W_{s,k-h}) (\cdot).$$

Applications

Extremes: Define $M_n = \max\{X_1, \dots, X_n\}$ and observe that

$$\{b_n^{-2}M_n \leq x\} = \{N_n(x, \infty] = 0\}$$

where N_n is the point process $\sum_{k=1}^n \varepsilon_{b_n^{-2}X_k}$. Thus,

$$P[b_n^{-2}M_n \leq x] = P[N_n(x, \infty] = 0]$$

$$\rightarrow P[N(x, \infty] = 0]$$

$$= \begin{cases} 0, & \text{if } x \leq 0, \\ \exp\{E(V_1^{\alpha/2})x^{-\alpha}\}, & \text{if } x > 0, \end{cases}$$

and $V_1 = \max\{c^k W_{1,k}; k > 0\}$.

Partial sums: Assume $\alpha \in (0,4)$ and define

$$S_n = \sum_{k=1}^n X_k.$$

Then the partial sums, normalized by b_n^{-2} , are asymptotically stable. For the case $\alpha \in (0,2)$, it follows directly from part (i) of the theorem that

$$\begin{aligned} b_n^{-2} S_n &\xrightarrow{d} S := \sum_{k=1}^{\infty} \sum_{s=1}^{\infty} j_s^2 c^k W_{s,k} \\ &= \sum_{s=1}^{\infty} j_s^2 A_s, \end{aligned}$$

which has a stable distribution. (See Davis and Hsing '95 for form of characteristic function.)

Sample ACF: Define the sample autocorrelation function w/o mean correction by

$$\hat{\rho}(h) = \frac{\sum_{t=1}^{n-h} X_t X_{t+h}}{\sum_{t=1}^n X_t^2}.$$

For heavy-tailed linear processes, $\hat{\rho}(h)$ was shown to be *consistent* by Davis and Resnick '85. (Limit distribution was also derived.) For the bilinear process, $\hat{\rho}(h)$ has a **nondegenerate** limit distribution w/o any normalization. Proof follows directly from part (ii) of the theorem.

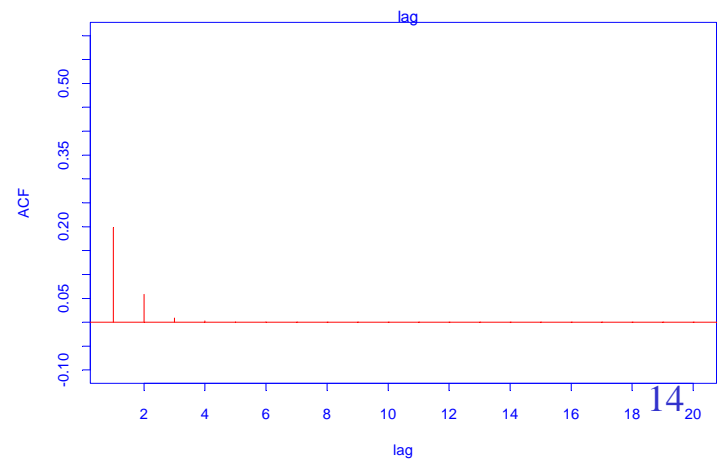
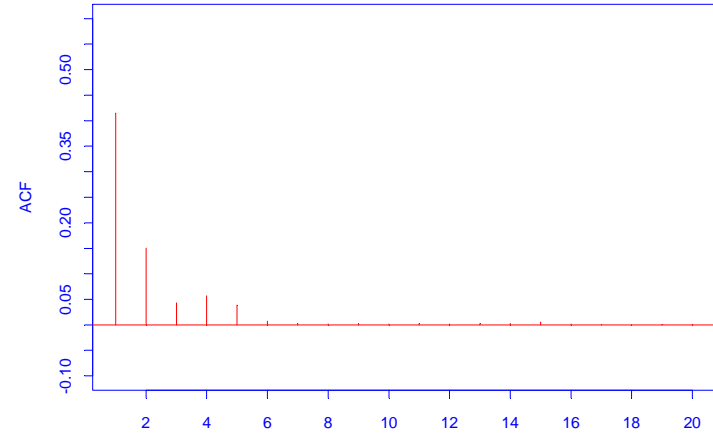
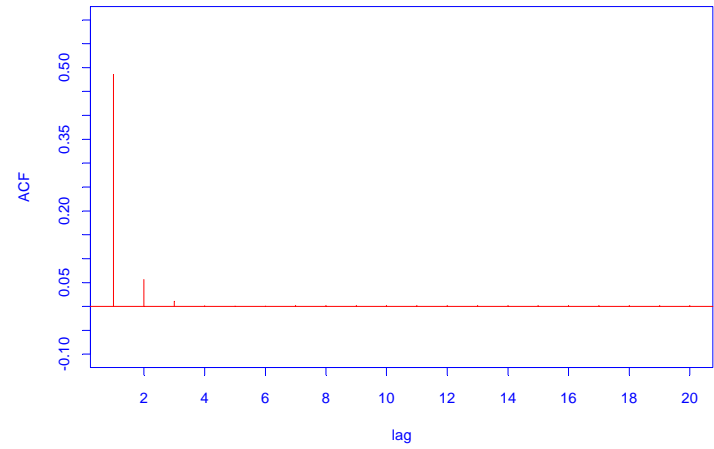
In particular, if $\alpha \in (0,4)$, then

$$\hat{\rho}(h) \xrightarrow{d} \frac{\sum_{k=1}^{\infty} \sum_{s=1}^{\infty} j_s^4 c^{2k-h} W_{s,k} W_{s,k-h}}{\sum_{k=1}^{\infty} \sum_{s=1}^{\infty} j_s^4 c^{2k} W_{s,k}^2}.$$

Examples: Sample ACF from 3 realizations, $n=5000$ from the model:

$$X_t = 0.1Z_{t-1}X_{t-1} + Z_t,$$

$\{Z_t\}$ IID, $P[Z_t > x] = 1/x, x \geq 1$.



II. Point process theory for stochastic volatility models

The Stochastic Volatility Model

$$Y_t = \exp(\alpha_t / 2) \xi_t, \quad \alpha_t = \sum_{j=0}^{\infty} c_j Z_{t-j},$$

where $\{\xi_t\}$ is IID $N(0,1)$, $\{Z_t\}$ is IID $N(0, \sigma^2)$ independent of $\{\xi_t\}$ and the $\{c_j\}$ are square summable.

Note:

(i) $\{Y_t\}$ is a stationary martingale difference sequence.

(ii) $X_t = \ln Y_t^2 = \alpha_t + \ln \xi_t^2$ is a Gaussian linear process plus an IID $\log\text{-}\chi^2$ process.

Tail Behavior of X_1

Using a Tauberian argument as in Feigin and Yashchin (1983) and Davis and Resnick (1991), we have

$$\begin{aligned} \mathbb{P}[X_1 > x] \approx & \frac{\sigma_\alpha}{\sqrt{\pi}} \exp \left\{ -\frac{x^2}{2\sigma_\alpha^2} + \frac{x \ln x}{\sigma_\alpha^2} + \frac{(k-1)x}{\sigma_\alpha^2} \right. \\ & \left. - \frac{(k + \sigma_\alpha^2) \ln x}{2\sigma_\alpha^2} - \frac{\ln^2 x}{2\sigma_\alpha^2} - \frac{k^2}{2\sigma_\alpha^2} + O\left(\frac{\ln^2 x}{x}\right) \right\}, \end{aligned}$$

as $x \rightarrow \infty$, where $k = \ln(2 / \sigma_\alpha^2)$, and $\sigma_\alpha^2 = \text{Var}(\alpha_t)$.

Point Process Convergence

Choose a_n and b_n such that

$$nP[a_n(X_1 - b_n) > x] \longrightarrow e^{-x},$$

i.e. can take

$$a_n = (2 / \sigma_\alpha^2)^{1/2} d_n, \quad d_n = (\ln n)^{1/2},$$

and

$$b_n = c_1 d_n + c_2 \ln d_n + c_3 + c_4 \frac{\ln d_n}{d_n} + c_5 \frac{1}{d_n}.$$

Then, if $\text{Cor}(\alpha_t, \alpha_{t+h}) \ln h \rightarrow 0$, as $h \rightarrow \infty$,

$$P[a_n(M_n - b_n) \leq x] \longrightarrow \exp\{-e^{-x}\},$$

$(M_n = \max\{X_1, \dots, X_n\})$.

More generally, we have the point process convergence,

$$N_n(\cdot) = \sum_{t=1}^{\infty} \mathcal{E}_{(t/n, a_n(X_t - b_n))}(\cdot) \xrightarrow{d} N(\cdot),$$

where, $N(\cdot)$ is a Poisson process on $[0, \infty) \times (-\infty, \infty]$ with intensity measure $dt \times \exp(-x)dx$

Remarks:

- Scaling constants a_n are the same as in the IID Gaussian case. Location constants b_n differ from those in the IID case (call these b_n^*). Then,

$$b_n - b_n^* = c_2 \ln d_n + c_3 + c \frac{1}{d_n},$$

so that b_n is ‘slightly larger’ than b_n^* .

- The condition, $\text{Cor}(\alpha_t, \alpha_{t+h}) \ln h \rightarrow 0$, as $h \rightarrow \infty$, is satisfied by both **short memory** ARMA and **long memory** fractionally integrated ARMA stochastic volatility models.
- Limit approximation is not as good when $\text{Cor}(\alpha_t, \alpha_{t+h})$ decays slowly or when σ_α^2 is large.