

# Regular Variation and Financial Time Series Models

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# Outline

- ☞ Characteristics of some financial time series
  - IBM returns
  - Multiplicative models for log-returns (GARCH, SV)
- ☞ Regular variation
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  - multivariate case
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- ☞ Applications of regular variation
  - Stochastic recurrence equations (GARCH)
  - Point process convergence
  - Extremes and extremal index
  - Limit behavior of sample correlations
- ☞ Wrap-up

## Characteristics of some financial time series

Define  $X_t = \ln(P_t) - \ln(P_{t-1})$  (log returns)

- heavy tailed

$$P(|X_1| > x) \sim C x^{-\alpha}, \quad 0 < \alpha < 4.$$

- uncorrelated

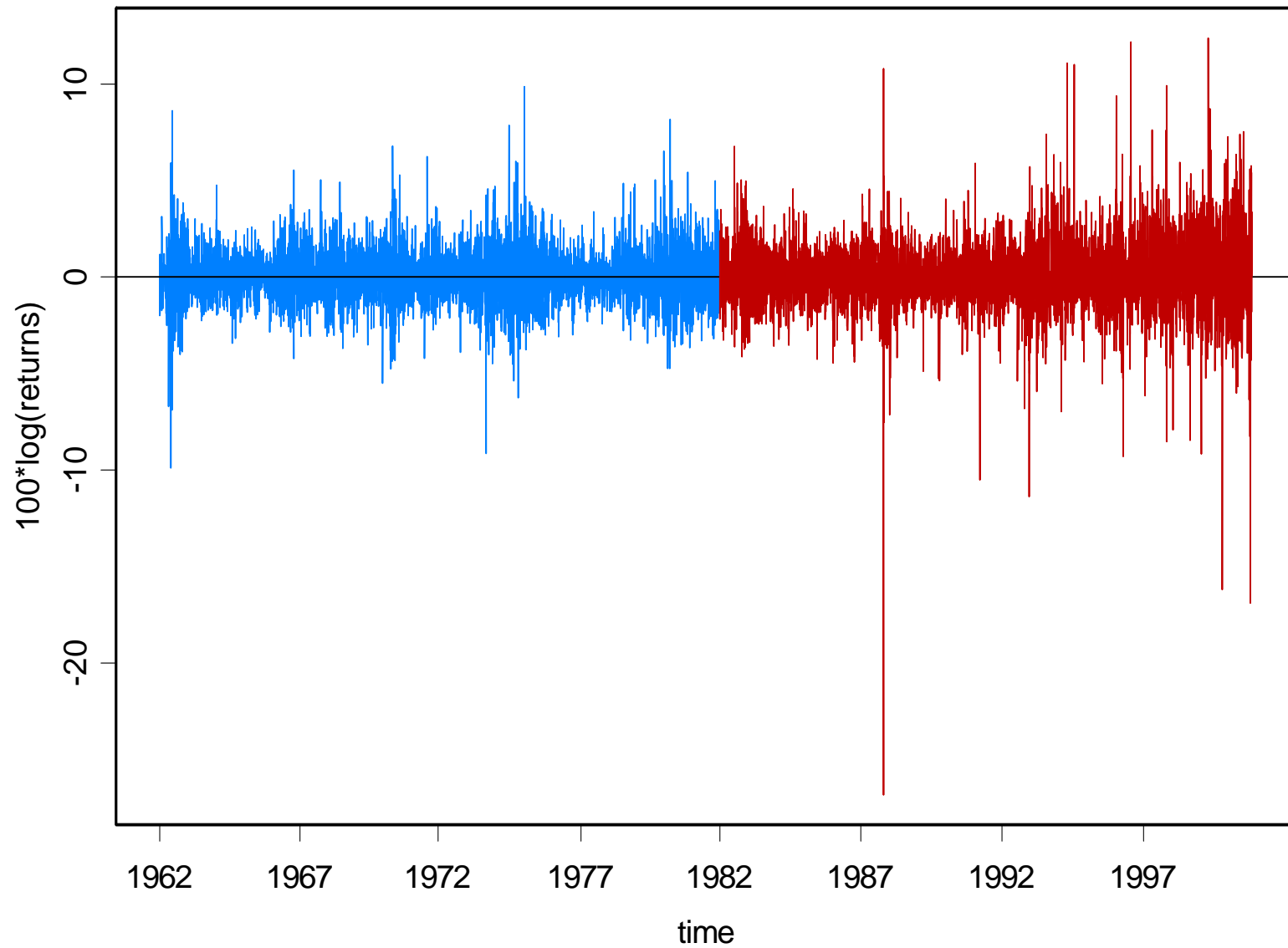
$\hat{\rho}_x(h)$  near 0 for all lags  $h > 0$  (MGD sequence)

- $|X_t|$  and  $X_t^2$  have slowly decaying autocorrelations

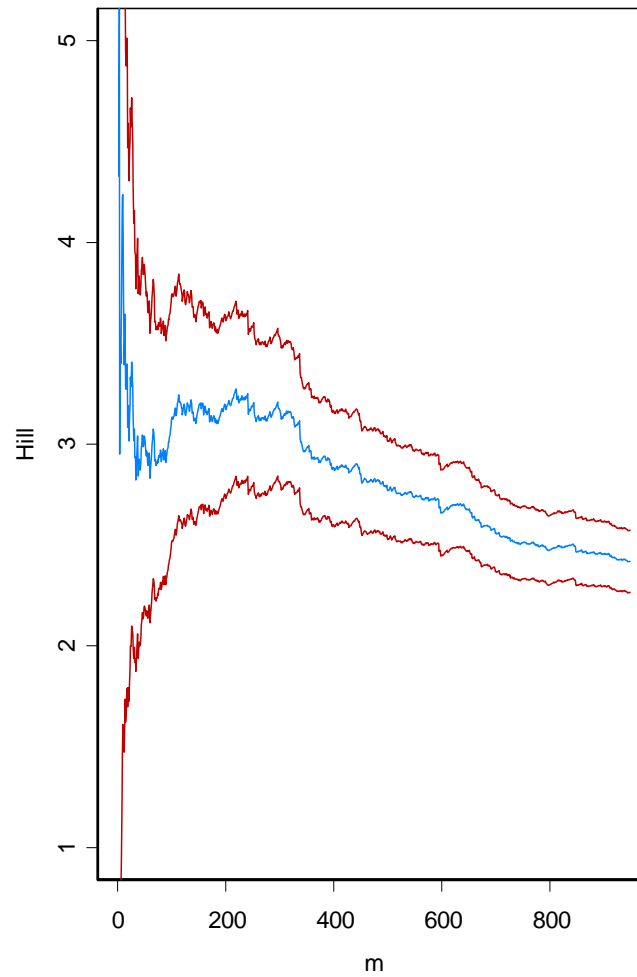
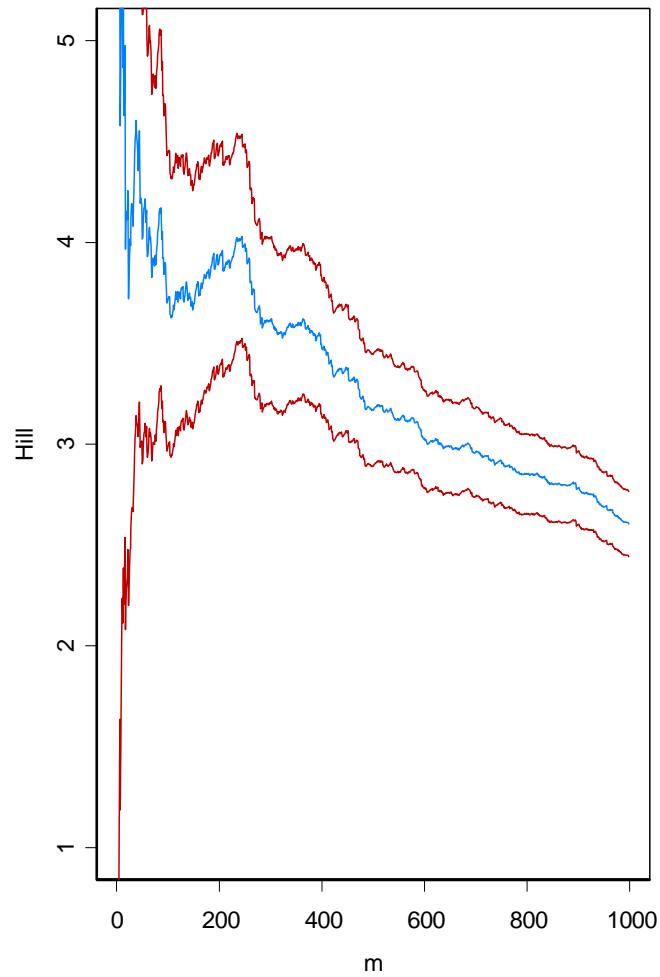
$\hat{\rho}_{|X|}(h)$  and  $\hat{\rho}_{X^2}(h)$  converge to 0 *slowly* as  $h$  increases.

- process exhibits 'volatility clustering'.

Log returns for IBM 1/3/62-11/3/00 (blue=1961-1981)



## Hill's plot of tail index for IBM (1962-1981, 1982-2000)



## Multiplicative models for log(returns)

### Basic model

$$\begin{aligned} X_t &= \ln(P_t) - \ln(P_{t-1}) \quad (\text{log returns}) \\ &= \sigma_t Z_t, \end{aligned}$$

where

- $\{Z_t\}$  is IID with mean 0, variance 1 (if exists). (e.g.  $N(0,1)$  or a  $t$ -distribution with  $\nu$  df.)
- $\{\sigma_t\}$  is the volatility process
- $\sigma_t$  and  $Z_t$  are independent.

### Properties:

- $EX_t = 0$ ,  $\text{Cov}(X_t, X_{t+h}) = 0$ ,  $h > 0$  (uncorrelated if  $\text{Var}(X_t) < \infty$ )
- conditional heteroscedastic (condition on  $\sigma_t$ ).

## Multiplicative models for log(returns)-cont

$$X_t = \sigma_t Z_t \quad (\text{observation eqn in state-space formulation})$$

Two classes of models for volatility:

(i) GARCH(p,q) process (General AutoRegressive Conditional Heteroscedastic-observation-driven specification)

$$\sigma_t^2 = \alpha_0 + \alpha_1 X_{t-1}^2 + \cdots + \alpha_p X_{t-p}^2 + \beta_1 \sigma_{t-1}^2 + \cdots + \beta_q \sigma_{t-q}^2 .$$

Special case: ARCH(1):

$$\begin{aligned} X_t^2 &= (\alpha_0 + \alpha_1 X_{t-1}^2) Z_t^2 \\ &= \alpha_1 Z_t^2 X_{t-1}^2 + \alpha_0 Z_t^2 \\ &= A_t X_{t-1}^2 + B_t \end{aligned} \quad (\text{stochastic recurrence eqn})$$

$$\rho_{X^2}(h) = \alpha_1^h, \text{ if } \alpha_1^2 < 1/3.$$

## Multiplicative models for log(returns)-cont

$$X_t = \sigma_t Z_t \text{ (observation eqn in state-space formulation)}$$

(ii) stochastic volatility process (parameter-driven specification)

$$\log \sigma_t^2 = \sum_{j=-\infty}^{\infty} \psi_j \varepsilon_{t-j}, \quad \sum_{j=-\infty}^{\infty} \psi_j^2 < \infty, \quad \{\varepsilon_t\} \sim \text{IIDN}(0, \sigma^2)$$

$$\rho_{X^2}(h) = \text{Cor}(\sigma_t^2, \sigma_{t+h}^2) / EZ_1^4$$

Question:

- Joint distributions of process regularly varying if distr of  $Z_1$  is regularly varying?



## Regular variation — multivariate case

Multivariate regular variation of  $\mathbf{X}=(X_1, \dots, X_m)$ : There exists a random vector  $\theta \in \mathbf{S}^{m-1}$  such that

$$P(|\mathbf{X}| > t x, \mathbf{X}/|\mathbf{X}| \in \bullet) / P(|\mathbf{X}| > t) \rightarrow_v x^{-\alpha} P(\theta \in \bullet)$$

( $\rightarrow_v$  vague convergence on  $\mathbf{S}^{m-1}$ , unit sphere in  $\mathbf{R}^m$ ).

- $P(\theta \in \bullet)$  is called the **spectral measure**
- $\alpha$  is the **index of  $\mathbf{X}$** .

Equivalence:

$$\frac{P(\mathbf{X} \in t\bullet)}{P(|\mathbf{X}| > t)} \rightarrow_v \mu(\bullet)$$

$\mu$  is a measure on  $\mathbf{R}^m$  which satisfies for  $x > 0$  and  $A$  bounded away from 0,

$$\mu(xB) = x^{-\alpha} \mu(xA).$$

## Regular variation — multivariate case (cont)

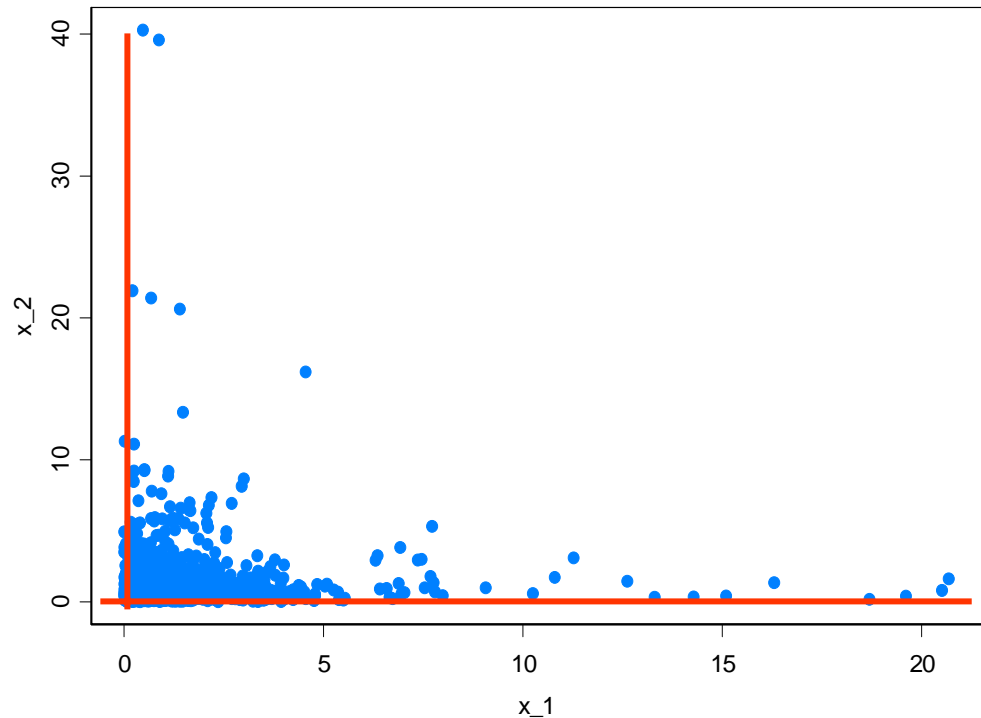
Examples:

1. If  $X_1 > 0$  and  $X_2 > 0$  are iid  $RV(\alpha)$ , then  $\mathbf{X} = (X_1, X_2)$  is multivariate regularly varying with index  $\alpha$  and *spectral distribution*

$$P(\theta = (0,1)) = P(\theta = (1,0)) = .5 \quad (\text{mass on axes}).$$

**Interpretation:** Unlikely that  $X_1$  and  $X_2$  are very large at the same time.

**Figure:** plot of  $(X_{t1}, X_{t2})$  for realization of 10,000.



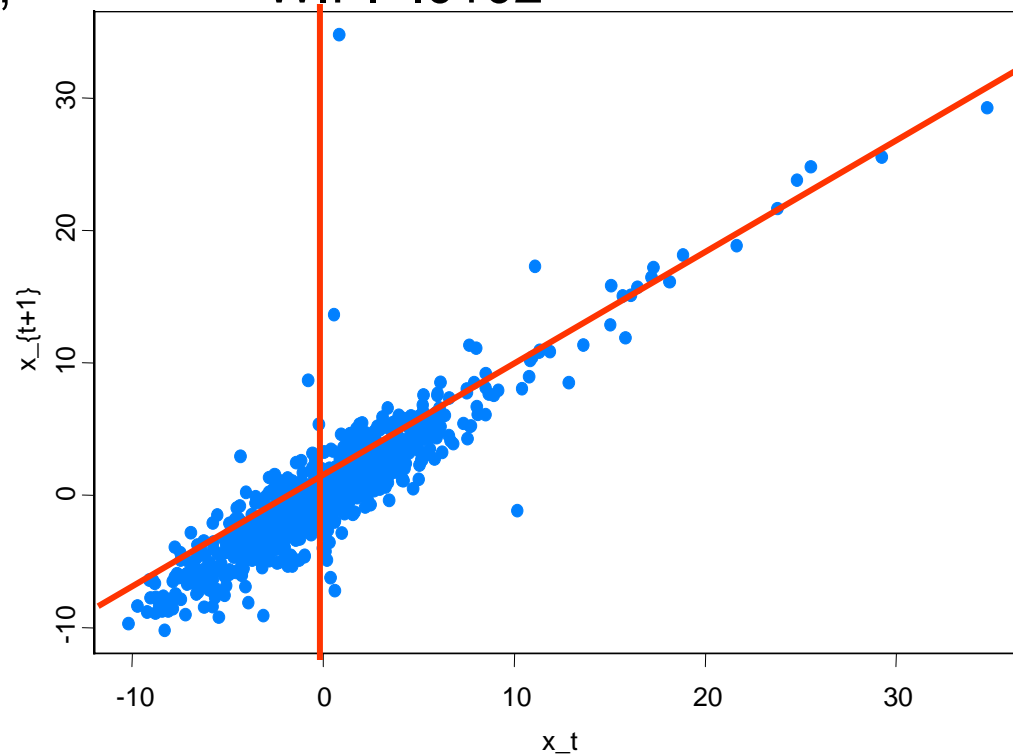
2. If  $X_1 = X_2 > 0$ , then  $\mathbf{X} = (X_1, X_2)$  is multivariate regularly varying with index  $\alpha$  and *spectral distribution*

$$P(\theta = (1/\sqrt{2}, 1/\sqrt{2})) = 1.$$

3. AR(1):  $X_t = .9 X_{t-1} + Z_t$ ,  $\{Z_t\} \sim \text{IID symmetric stable (1.8)}$

Distr of  $\theta$ :  $\begin{cases} \pm(1,.9)/\text{sqrt}(1.81), \text{ W.P. } .9898 \\ \pm(0,1), \\ \text{ W.P. } .0102 \end{cases}$

**Figure:** plot of  $(X_t, X_{t+1})$  for realization of 10,000.



## Applications of multivariate regular variation

- Domain of attraction for *sums of iid random vectors* (Rvaceva, 1962). That is, when does the partial sum

$$a_n^{-1} \sum_{t=1}^n \mathbf{X}_t$$

converge for some constants  $a_n$ ?

- *Spectral measure* of multivariate stable vectors.
- *Domain of attraction* for componentwise maxima of iid random vectors (Resnick, 1987). Limit behavior of

$$a_n^{-1} \bigvee_{t=1}^n \mathbf{X}_t$$

- Weak convergence of *point processes* with iid points.
- Solution to *stochastic recurrence equations*,  $\mathbf{Y}_t = \mathbf{A}_t \mathbf{Y}_{t-1} + \mathbf{B}_t$
- Weak convergence of *sample autocovariances*.

## Applications of multivariate regular variation (cont)

### Linear combinations:

$\mathbf{X} \sim \text{RV}(\alpha) \Rightarrow$  all linear combinations of  $\mathbf{X}$  are regularly varying

i.e., there exist  $\alpha$  and slowly varying fcn  $L(\cdot)$ , s.t.

$$P(\mathbf{c}^T \mathbf{X} > t) / (t^\alpha L(t)) \rightarrow w(\mathbf{c}), \text{ exists for all real-valued } \mathbf{c},$$

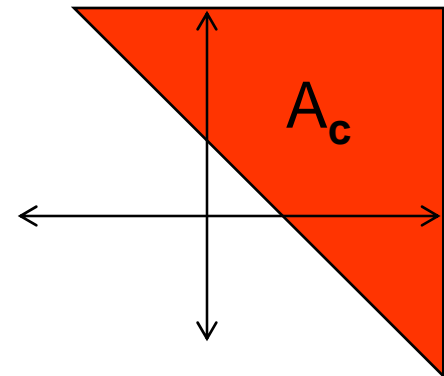
where

$$w(t\mathbf{c}) = t^{-\alpha} w(\mathbf{c}).$$

Use vague convergence with  $A_{\mathbf{c}} = \{\mathbf{y} : \mathbf{c}^T \mathbf{y} > 1\}$ , i.e.,

$$\frac{P(\mathbf{X} \in tA_{\mathbf{c}})}{t^{-\alpha} L(t)} = \frac{P(\mathbf{c}^T \mathbf{X} > t)}{P(|\mathbf{X}| > t)} \rightarrow \mu(A_{\mathbf{c}}) =: w(\mathbf{c}),$$

where  $t^\alpha L(t) = P(|\mathbf{X}| > t)$ .



## Applications of multivariate regular variation (cont)

### Converse?

$\mathbf{X} \sim \text{RV}(\alpha) \iff$  all linear combinations of  $\mathbf{X}$  are regularly varying?

There exist  $\alpha$  and slowly varying fcn  $L(\cdot)$ , s.t.

(LC)  $P(\mathbf{c}^T \mathbf{X} > t) / (t^\alpha L(t)) \rightarrow w(\mathbf{c})$ , exists for all real-valued  $\mathbf{c}$ .

Theorem (Basrak, Davis, Mikosch, '02). Let  $\mathbf{X}$  be a random vector.

1. If  $\mathbf{X}$  satisfies (LC) with  $\alpha$  non-integer, then  $\mathbf{X}$  is  $\text{RV}(\alpha)$ .
2. If  $\mathbf{X} > 0$  satisfies (LC) for non-negative  $\mathbf{c}$  and  $\alpha$  is non-integer, then  $\mathbf{X}$  is  $\text{RV}(\alpha)$ .
3. If  $\mathbf{X} > 0$  satisfies (LC) with  $\alpha$  an odd integer, then  $\mathbf{X}$  is  $\text{RV}(\alpha)$ .

## Applications of theorem

1. **Kesten (1973)**. Under general conditions, **(LC)** holds with  $L(t)=1$  for stochastic recurrence equations of the form

$$\mathbf{Y}_t = \mathbf{A}_t \mathbf{Y}_{t-1} + \mathbf{B}_t, \quad (\mathbf{A}_t, \mathbf{B}_t) \sim \text{IID},$$

$\mathbf{A}_t$   $d \times d$  random matrices,  $\mathbf{B}_t$  random  $d$ -vectors.

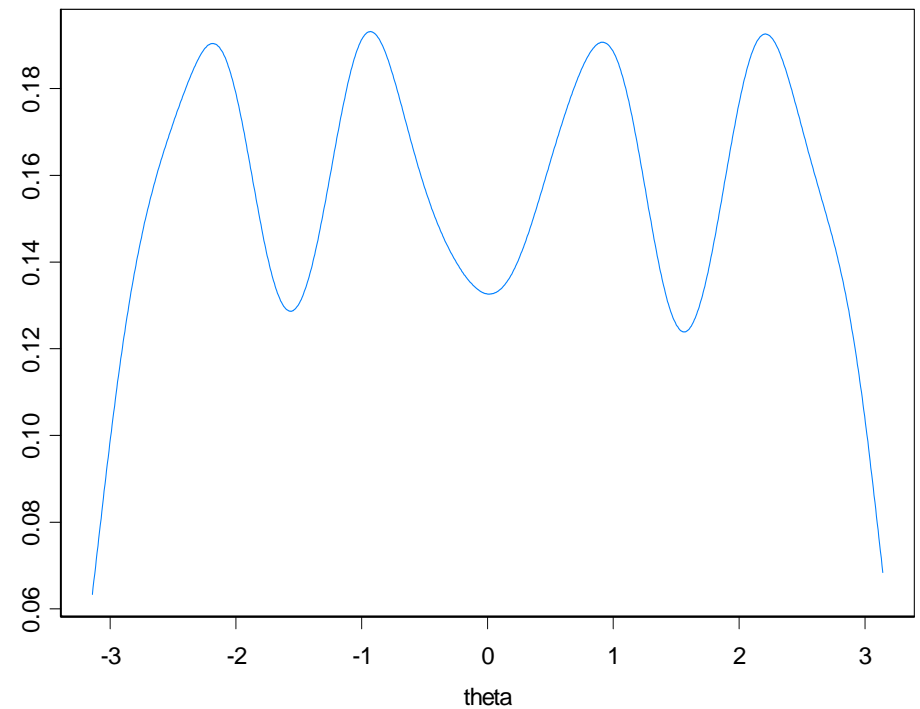
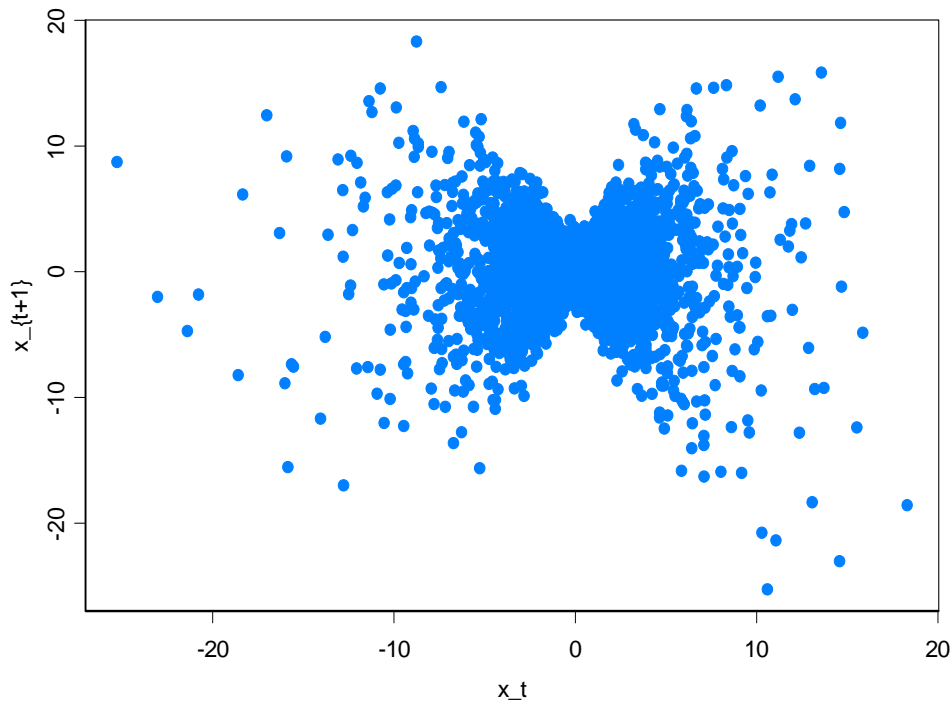
It follows that the distributions of  $\mathbf{Y}_t$ , and in fact all of the finite dim'l distrs of  $\mathbf{Y}_t$  are regularly varying (if  $\alpha$  is non-even).

2. **GARCH processes**. Since squares of a GARCH process can be embedded in a SRE, the *finite dimensional distributions* of a **GARCH** are regularly varying.

## Examples (cont)

Example of ARCH(1):  $\alpha_0=1, \alpha_1=1, \alpha=2, X_t=(\alpha_0+\alpha_1 X_{t-1}^2)^{1/2}Z_t, \{Z_t\}\sim\text{IID}$

Figures: plots of  $(X_t, X_{t+1})$  and estimated distribution of  $\theta$  for realization of 10,000.





## Applications of theorem (cont)

Example: SV model  $X_t = \sigma_t Z_t$

Suppose  $Z_t \sim \text{RV}(\alpha)$  and

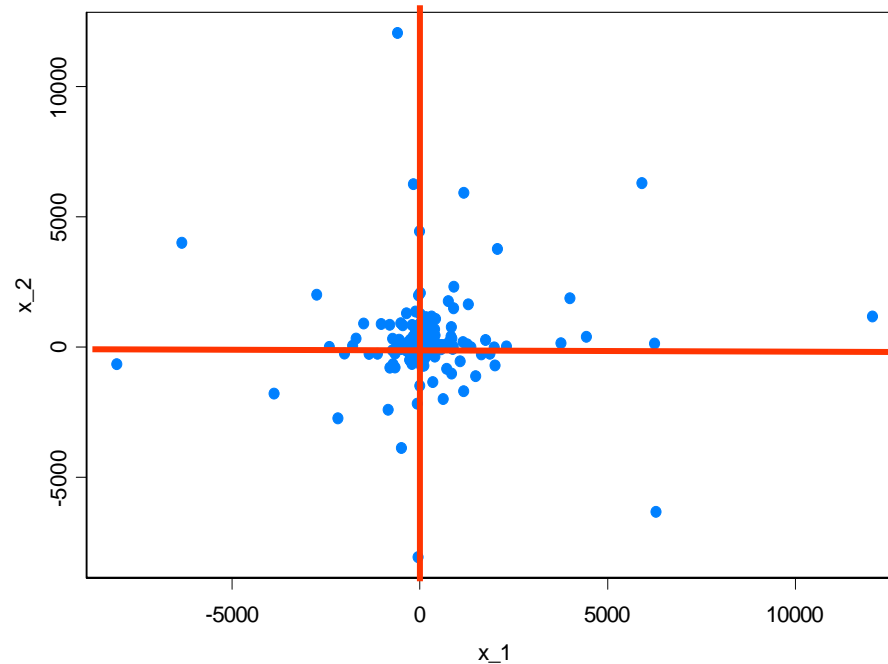
$$\log \sigma_t^2 = \sum_{j=-\infty}^{\infty} \psi_j \varepsilon_{t-j}, \quad \sum_{j=-\infty}^{\infty} \psi_j^2 < \infty, \quad \{\varepsilon_t\} \sim \text{IIDN}(0, \sigma^2).$$

Then  $\mathbf{Z}_n = (Z_1, \dots, Z_n)'$  is regular varying with index  $\alpha$  and so is

$$\mathbf{X}_n = (X_1, \dots, X_n)' = \text{diag}(\sigma_1, \dots, \sigma_n) \mathbf{Z}_n$$

with spectral distribution concentrated on  $(\pm 1, 0), (0, \pm 1)$ .

Figure: plot of  
 $(X_t, X_{t+1})$  for  
realization of 10,000.



## Extremes for GARCH and SV processes

### Setup

- $X_t = \sigma_t Z_t$ ,  $\{Z_t\} \sim \text{IID}(0,1)$
- $X_t$  is RV ( $\alpha$ )
- Choose  $\{b_n\}$  s.t.  $nP(X_t > b_n) \rightarrow 1$

Then

$$P^n(b_n^{-1} X_1 \leq x) \rightarrow \exp\{-x^{-\alpha}\}.$$

Then, with  $M_n = \max\{X_1, \dots, X_n\}$ ,

(i) GARCH:

$$P(b_n^{-1} M_n \leq x) \rightarrow \exp\{-\gamma x^{-\alpha}\},$$

$\gamma$  is extremal index ( $0 < \gamma < 1$ ).

(ii) SV model:

$$P(b_n^{-1} M_n \leq x) \rightarrow \exp\{-x^{-\alpha}\},$$

extremal index  $\gamma = 1$  no clustering.

## Extremes for GARCH and SV processes (cont)

(i) GARCH:  $P(b_n^{-1}M_n \leq x) \rightarrow \exp\{-\gamma x^{-\alpha}\}$

(ii) SV model:  $P(b_n^{-1}M_n \leq x) \rightarrow \exp\{-x^{-\alpha}\}$

### Remarks about extremal index.

(i)  $\gamma < 1$  implies clustering of exceedances

(ii) Numerical example. Suppose  $c$  is a threshold such that

$$P^n(b_n^{-1}X_1 \leq c) \sim .95$$

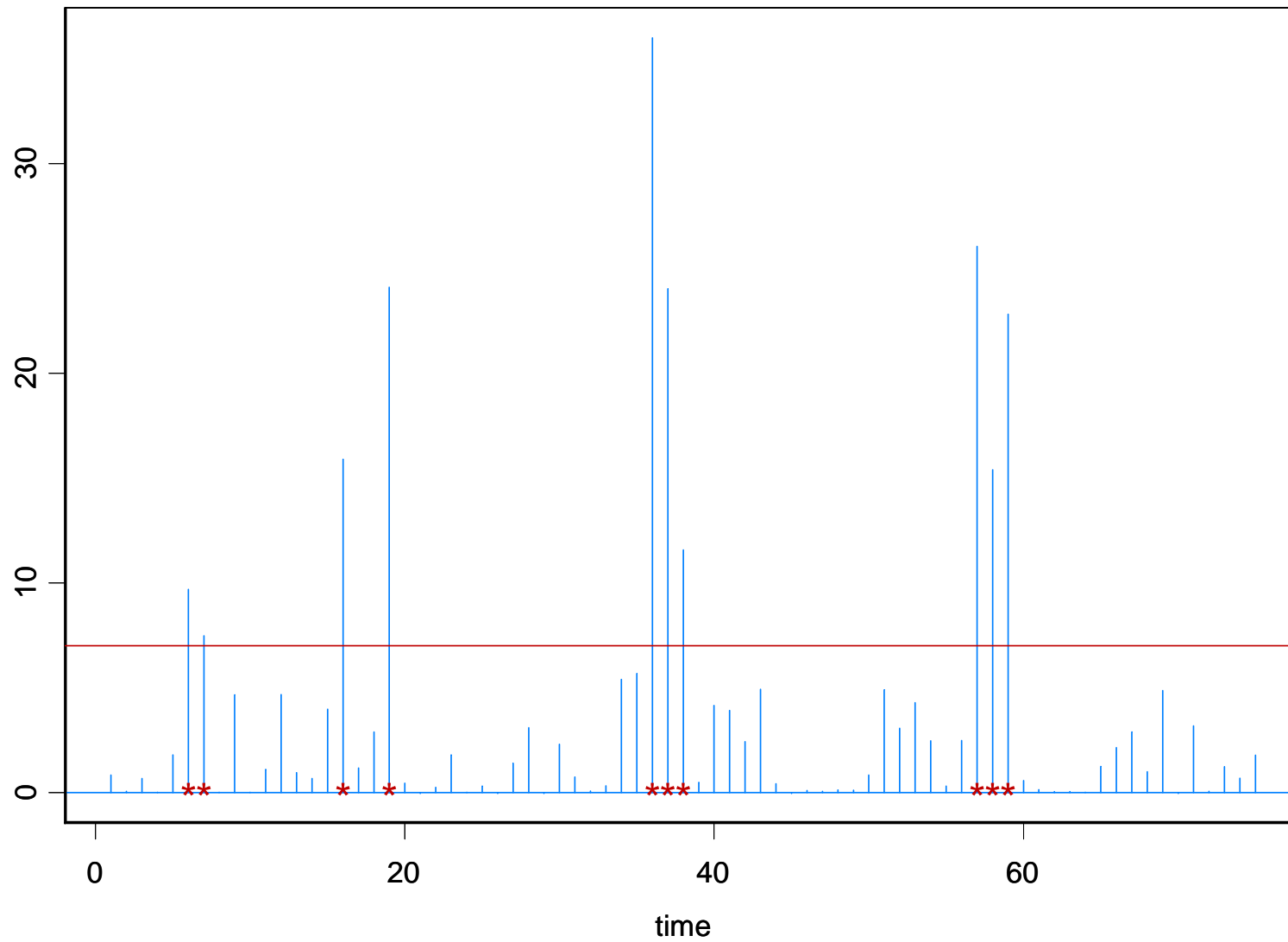
Then, if  $\gamma = .5$ ,  $P(b_n^{-1}M_n \leq c) \sim (.95)^{.5} = .975$

(iii)  $1/\gamma$  is the *mean cluster size* of exceedances.

(iv) Use  $\gamma$  to *discriminate* between GARCH and SV models.

(v) Even for the light-tailed SV model (i.e.,  $\{Z_t\} \sim \text{IID } N(0,1)$ ), the *extremal index* is 1 (see Breidt and Davis '98 )

## Extremes for GARCH and SV processes (cont)



## Summary of results for ACF of GARCH(p,q) and SV models

### GARCH(p,q)

$\alpha \in (0,2)$ :

$$(\hat{\rho}_X(h))_{h=1,\dots,m} \xrightarrow{d} (V_h / V_0)_{h=1,\dots,m},$$

$\alpha \in (2,4)$ :

$$(n^{1-2/\alpha} \hat{\rho}_X(h))_{h=1,\dots,m} \xrightarrow{d} \gamma_X^{-1}(0)(V_h)_{h=1,\dots,m}.$$

$\alpha \in (4,\infty)$ :

$$(n^{1/2} \hat{\rho}_X(h))_{h=1,\dots,m} \xrightarrow{d} \gamma_X^{-1}(0)(G_h)_{h=1,\dots,m}.$$

**Remark:** Similar results hold for the sample ACF based on  $|X_t|$  and  $X_t^2$ .

## Summary of results for ACF of GARCH(p,q) and SV models (cont)

### SV Model

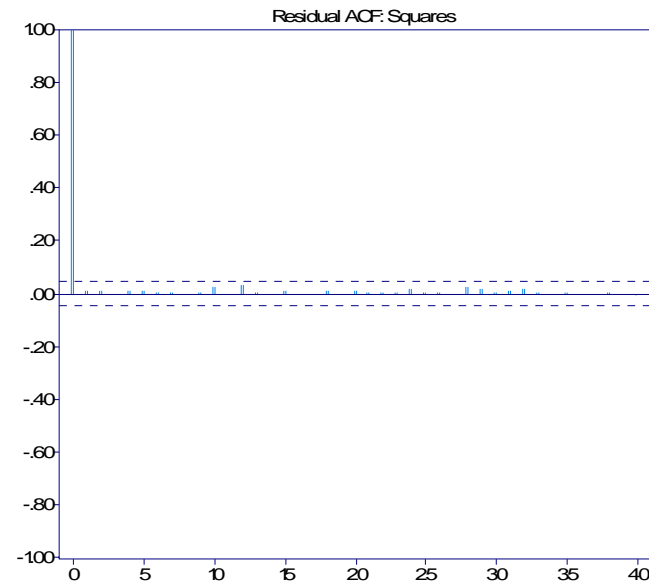
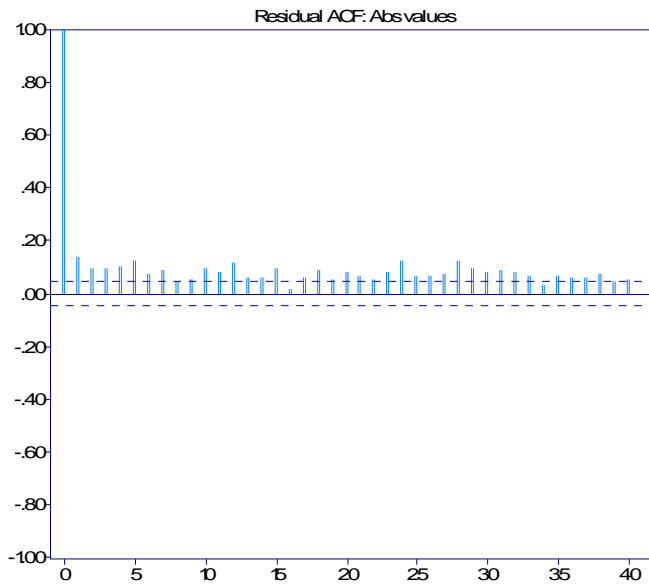
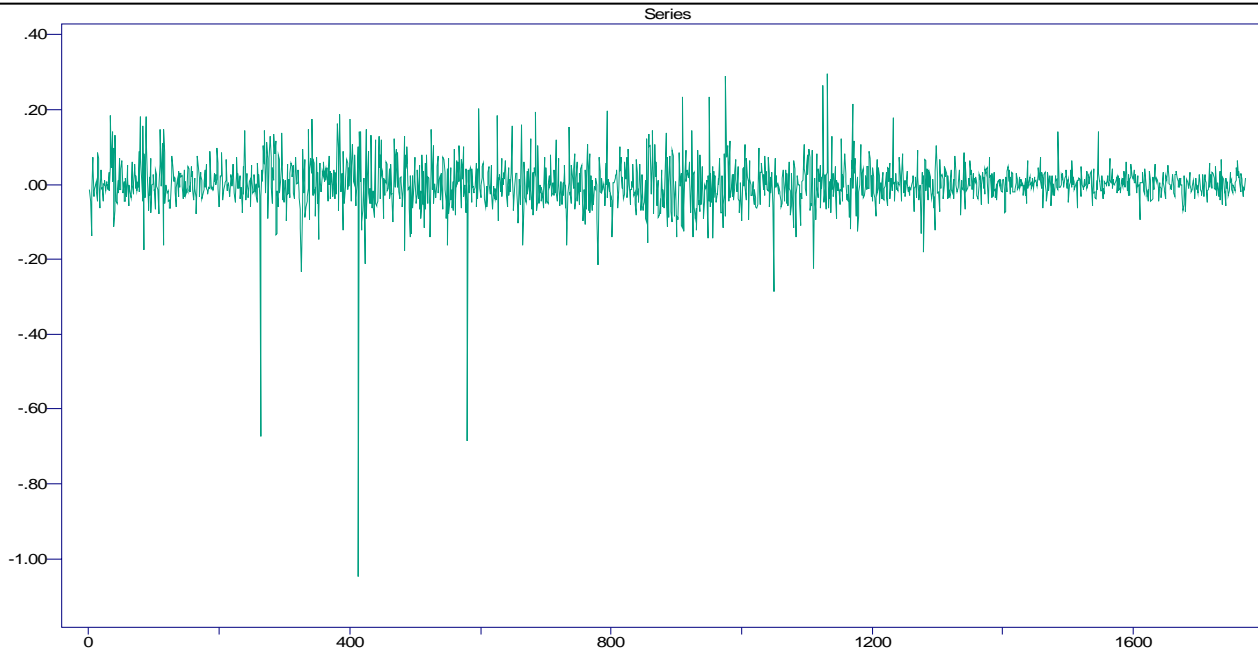
$\alpha \in (0, 2)$ :

$$(n / \ln n)^{1/\alpha} \hat{\rho}_X(h) \xrightarrow{d} \frac{\|\sigma_1 \sigma_{h+1}\|_\alpha}{\|\sigma_1\|_\alpha^2} \frac{S_h}{S_0}.$$

$\alpha \in (2, \infty)$ :

$$\left( n^{1/2} \hat{\rho}_X(h) \right)_{h=1, \dots, m} \xrightarrow{d} \gamma_X^{-1}(0) (G_h)_{h=1, \dots, m}.$$

# Amazon returns May 16, 1997 to June 16, 2004.



## Wrap-up

- *Regular variation* is a flexible tool for modeling both *dependence* and *tail heaviness*.
- Useful for establishing *point process convergence* of heavy-tailed time series.
- *Extremal index*  $\gamma < 1$  for GARCH and  $\gamma = 1$  for SV.

### Unresolved issues related to $RV \Leftrightarrow (LC)$

- $\alpha = 2n$ ?
- there is an example for which  $\mathbf{X}_1, \mathbf{X}_2 > 0$ , and  $(\mathbf{c}, \mathbf{X}_1)$  and  $(\mathbf{c}, \mathbf{X}_2)$  have the same limits for all  $\mathbf{c} > \mathbf{0}$ .
- $\alpha = 2n-1$  and  $\mathbf{X} \not\geq 0$  (not true in general).