

# Structural Break Detection for a Class of Nonlinear Time Series Models

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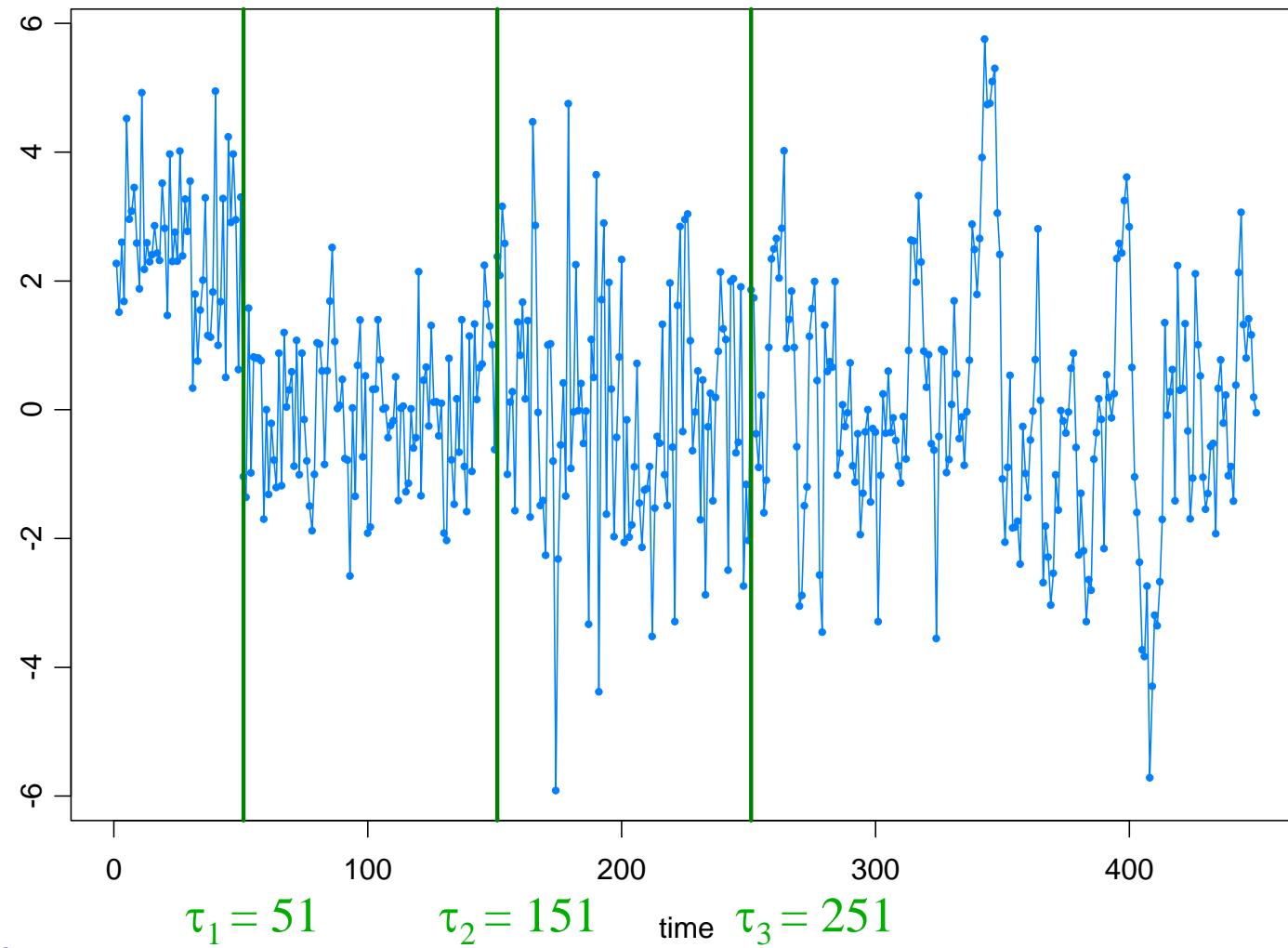
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(<http://www.stat.colostate.edu/~rdavis/lectures>)

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## Illustrative Example

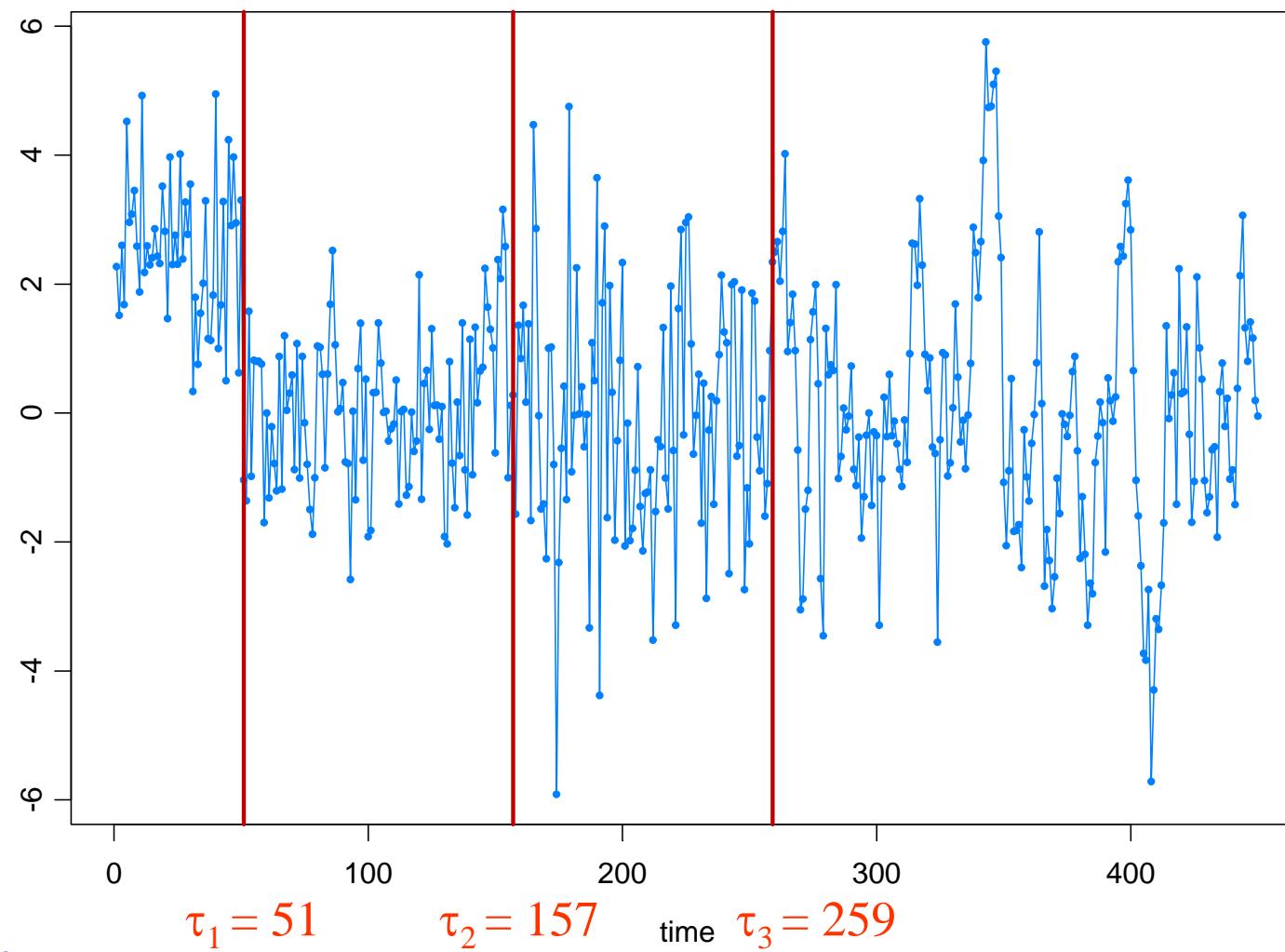
How many segments do you see?



## Illustrative Example

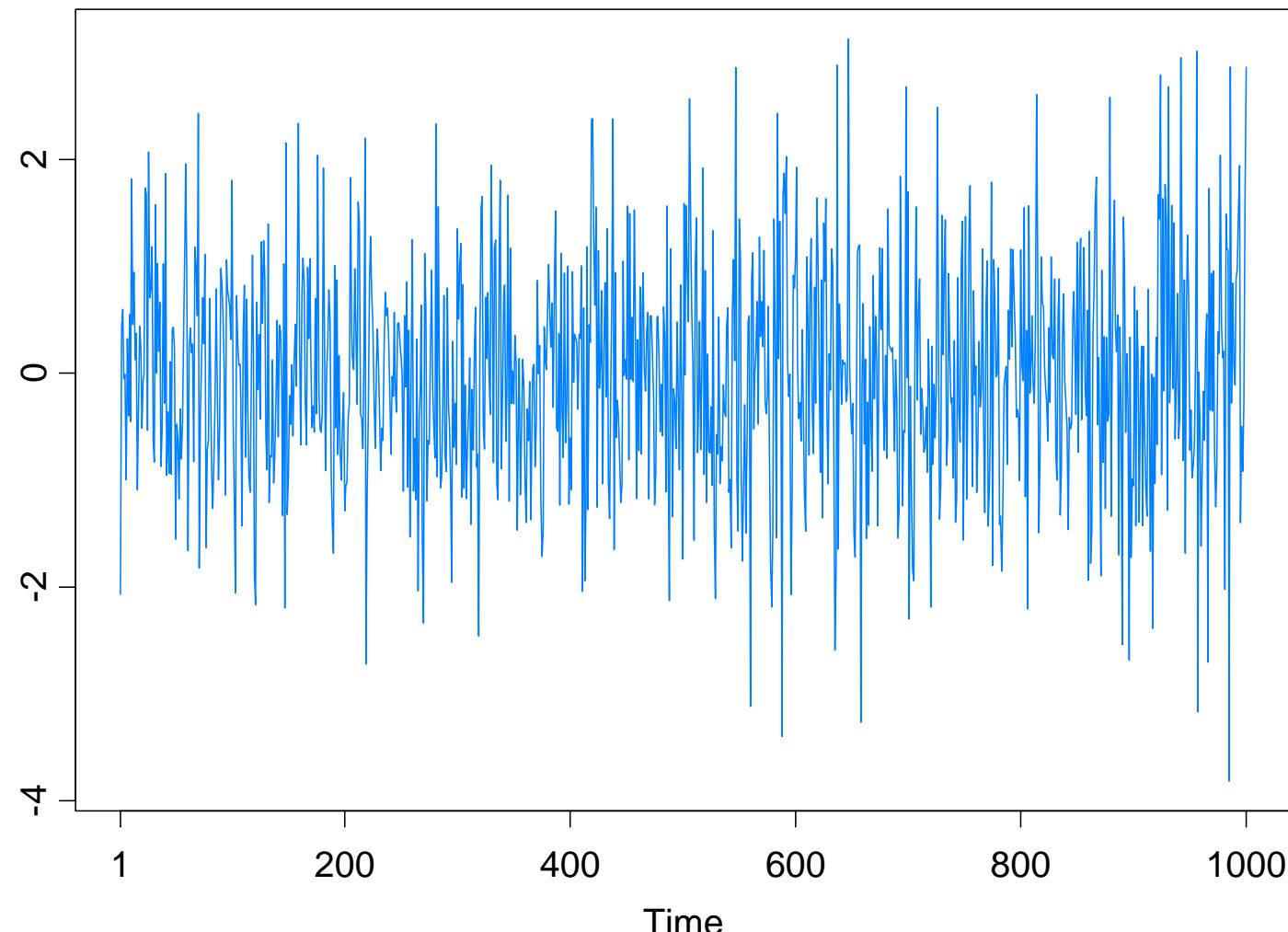
Auto-PARM=Auto-Piecewise AutoRegressive Modeling

4 pieces, 2.58 seconds.



## A Second Example

Any breaks in this series?



- Introduction
  - Examples
    - AR
    - GARCH
    - Stochastic volatility
    - State space models
- Model selection using Minimum Description Length (MDL)
  - General principles
  - Application to AR models with breaks
- Optimization using a Genetic Algorithm
  - Basics
  - Implementation for structural break estimation
- Simulation results
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- Simulation results for GARCH and SV models

## Examples

### 1. Piecewise AR model:

$$Y_t = \gamma_j + \phi_{j1} Y_{t-1} + \dots + \phi_{jp_j} Y_{t-p_j} + \sigma_j \varepsilon_t, \quad \text{if } \tau_{j-1} \leq t < \tau_j,$$

where  $\tau_0 = 1 < \tau_1 < \dots < \tau_{m-1} < \tau_m = n + 1$ , and  $\{\varepsilon_t\}$  is IID(0,1).

Goal: Estimate

$m$  = number of segments

$\tau_j$  = location of  $j^{\text{th}}$  break point

$\gamma_j$  = level in  $j^{\text{th}}$  epoch

$p_j$  = order of AR process in  $j^{\text{th}}$  epoch

$(\phi_{j1}, \dots, \phi_{jp_j})$  = AR coefficients in  $j^{\text{th}}$  epoch

$\sigma_j$  = scale in  $j^{\text{th}}$  epoch

## Examples (cont)

### 2. Segmented GARCH model:

$$Y_t = \sigma_t \varepsilon_t,$$

$$\sigma_t^2 = \omega_j + \alpha_{j1} Y_{t-1}^2 + \dots + \alpha_{jp_j} Y_{t-p_j}^2 + \beta_{j1} \sigma_{t-1}^2 + \dots + \beta_{jq_j} \sigma_{t-q_j}^2, \quad \text{if } \tau_{j-1} \leq t < \tau_j,$$

where  $\tau_0 = 1 < \tau_1 < \dots < \tau_{m-1} < \tau_m = n + 1$ , and  $\{\varepsilon_t\}$  is IID(0,1).

### 3. Segmented stochastic volatility model:

$$Y_t = \sigma_t \varepsilon_t,$$

$$\log \sigma_t^2 = \gamma_j + \phi_{j1} \log \sigma_{t-1}^2 + \dots + \phi_{jp_j} \log \sigma_{t-p_j}^2 + v_j \eta_t, \quad \text{if } \tau_{j-1} \leq t < \tau_j.$$

### 4. Segmented state-space model (SVM a special case):

$$p(y_t | \alpha_t, \dots, \alpha_1, y_{t-1}, \dots, y_1) = p(y_t | \alpha_t) \text{ is specified}$$

$$\alpha_t = \gamma_j + \phi_{j1} \alpha_{t-1} + \dots + \phi_{jp_j} \alpha_{t-p_j} + \sigma_j \eta_t, \quad \text{if } \tau_{j-1} \leq t < \tau_j.$$

## Model Selection Using Minimum Description Length

Basics of MDL:

Choose the model which *maximizes the compression* of the data or, equivalently, select the model that *minimizes the code length* of the data (i.e., amount of memory required to encode the data).

$\mathcal{M}$  = class of operating models for  $y = (y_1, \dots, y_n)$

$L_F(y)$  = code length of  $y$  relative to  $F \in \mathcal{M}$

Typically, this term can be decomposed into two pieces (two-part code),

$$L_F(y) = L(\hat{F}/y) + L(\hat{e} | \hat{F}),$$

where

$L(\hat{F}/y)$  = code length of the fitted model for  $F$

$L(\hat{e}/\hat{F})$  = code length of the residuals based on the fitted model

## Model Selection Using Minimum Description Length (cont)

Applied to the segmented AR model:

$$Y_t = \gamma_j + \phi_{j1} Y_{t-1} + \dots + \phi_{jp_j} Y_{t-p_j} + \sigma_j \varepsilon_t, \quad \text{if } \tau_{j-1} \leq t < \tau_j,$$

First term  $L(\hat{\mathbf{F}}/y)$  :

$$\begin{aligned} L(\hat{\mathbf{F}}/y) &= L(m) + L(\tau_1, \dots, \tau_m) + L(p_1, \dots, p_m) + L(\hat{\psi}_1 | y) + \dots + L(\hat{\psi}_m | y) \\ &= \log_2 m + m \log_2 n + \sum_{j=1}^m \log_2 p_j + \sum_{j=1}^m \frac{p_j + 2}{2} \log_2 n_j \end{aligned}$$

Second term  $L(\hat{e} | \hat{\mathbf{F}})$  :

$$L(\hat{e} | \hat{\mathbf{F}}) \approx - \sum_{j=1}^m \log_2 L(\hat{\psi}_j | y)$$

$$MDL(m, (\tau_1, p_1), \dots, (\tau_m, p_m))$$

$$= \log_2 m + m \log_2 n + \sum_{j=1}^m \log_2 p_j + \sum_{j=1}^m \frac{p_j + 2}{2} \log_2 n_j + \sum_{j=1}^m (\log_2(2\pi\hat{\sigma}_j^2) + n_j)$$

## Optimization Using Genetic Algorithm

**Genetic Algorithm:** Chromosome consists of  $n$  genes, each taking the value of  $-1$  (no break) or  $p$  (order of AR process). Use natural selection to find a *near* optimal solution.

Map the break points with a chromosome  $c$  via

$$(m, (\tau_1, p_1), \dots, (\tau_m, p_m)) \longleftrightarrow c = (\delta_1, \dots, \delta_n),$$

where

$$\delta_t = \begin{cases} -1, & \text{if no break point at } t, \\ p_j, & \text{if break point at time } t = \tau_{j-1} \text{ and AR order is } p_j. \end{cases}$$

For example,

$$c = (2, -1, -1, -1, -1, 0, -1, -1, -1, -1, 0, -1, -1, -1, -1, 3, -1, -1, -1, -1, -1)$$
$$\begin{matrix} t: 1 & & 6 & & 11 & & 15 \end{matrix}$$

would correspond to a process as follows:

$$\text{AR}(2), t=1:5; \text{AR}(0), t=6:10; \text{AR}(0), t=11:14; \text{AR}(3), t=15:20$$

## Implementation of Genetic Algorithm—(cont)

**Generation 0:** Start with  $L$  (200) randomly generated chromosomes,  $c_1, \dots, c_L$  with associated MDL values,  $MDL(c_1), \dots, MDL(c_L)$ .

**Generation 1:** A new child in the next generation is formed from the chromosomes  $c_1, \dots, c_L$  of the previous generation as follows:

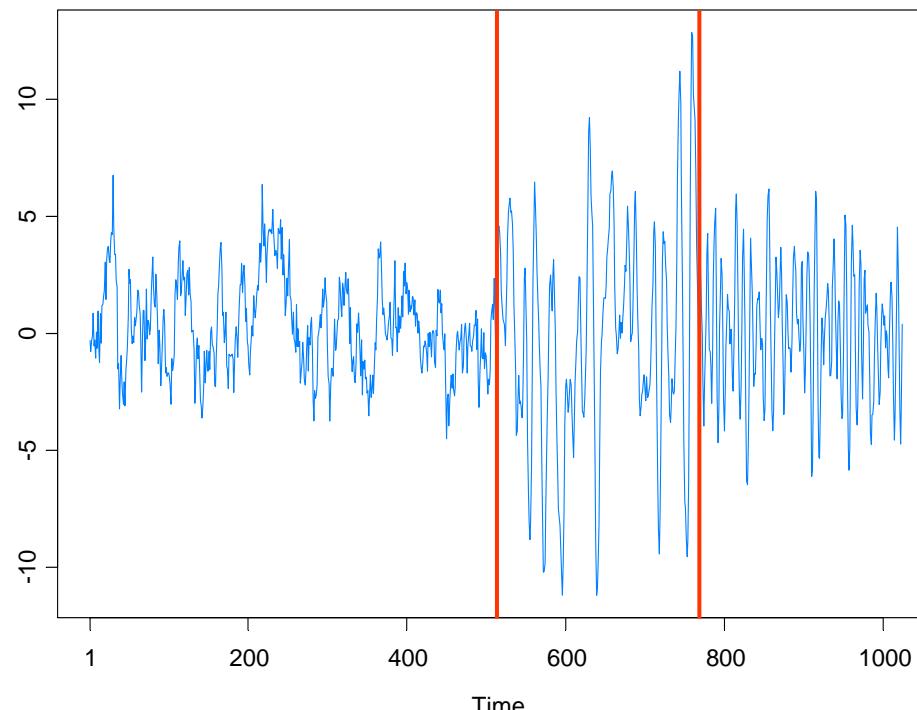
- with probability  $\pi_c$ , *crossover* occurs.
  - two parent chromosomes  $c_i$  and  $c_j$  are selected at random with probabilities proportional to the ranks of  $MDL(c_i)$ .
  - $k^{th}$  gene of child is  $\delta_k = \delta_{i,k}$  w.p.  $\frac{1}{2}$  and  $\delta_{j,k}$  w.p.  $\frac{1}{2}$
- with probability  $1 - \pi_c$ , *mutation* occurs.
  - a parent chromosome  $c_i$  is selected
  - $k^{th}$  gene of child is  $\delta_k = \delta_{i,k}$  w.p.  $\pi_1$ ;  $-1$  w.p.  $\pi_2$ ; and  $p$  w.p.  $1 - \pi_1 - \pi_2$ .

## Simulation Examples-based on Ombao et al. (2001) test cases

1. Piecewise stationary with dyadic structure: Consider a time series following the model,

$$Y_t = \begin{cases} .9Y_{t-1} + \varepsilon_t, & \text{if } 1 \leq t < 513, \\ 1.69Y_{t-1} - .81Y_{t-2} + \varepsilon_t, & \text{if } 513 \leq t < 769, \\ 1.32Y_{t-1} - .81Y_{t-2} + \varepsilon_t, & \text{if } 769 \leq t \leq 1024, \end{cases}$$

where  $\{\varepsilon_t\} \sim \text{IID } N(0,1)$ .



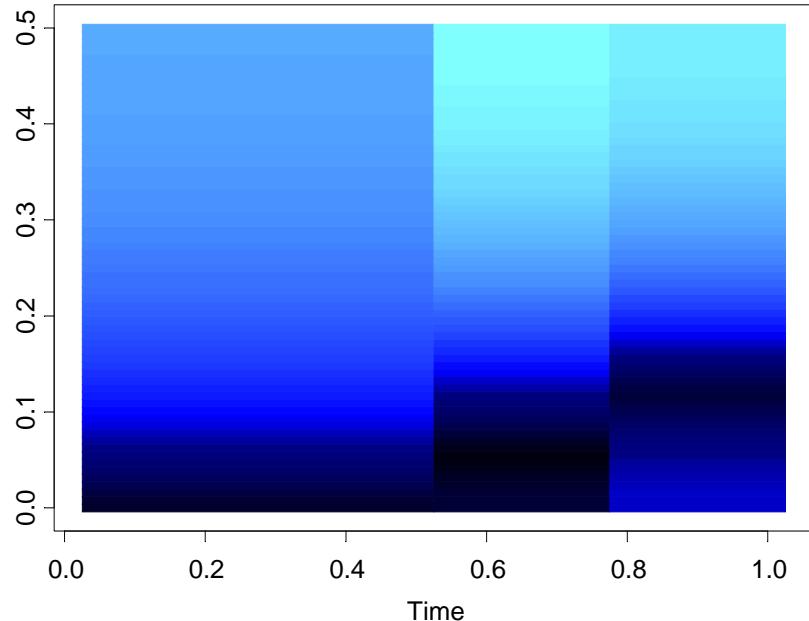
## 1. Piecewise stat (cont)

GA results: 3 pieces breaks at  $\tau_1=513$ ;  $\tau_2=769$ . Total run time 16.31 secs

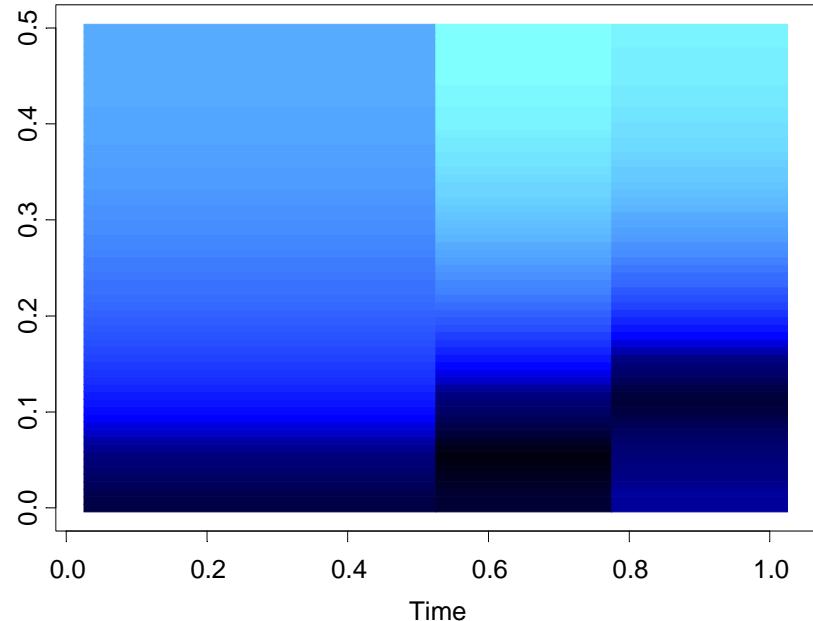
Fitted model:

	$\phi_1$	$\phi_2$	$\sigma^2$
1- 512:	.857		.9945
513- 768:	1.68	-0.801	1.1134
769-1024:	1.36	-0.801	1.1300

True Model



Fitted Model

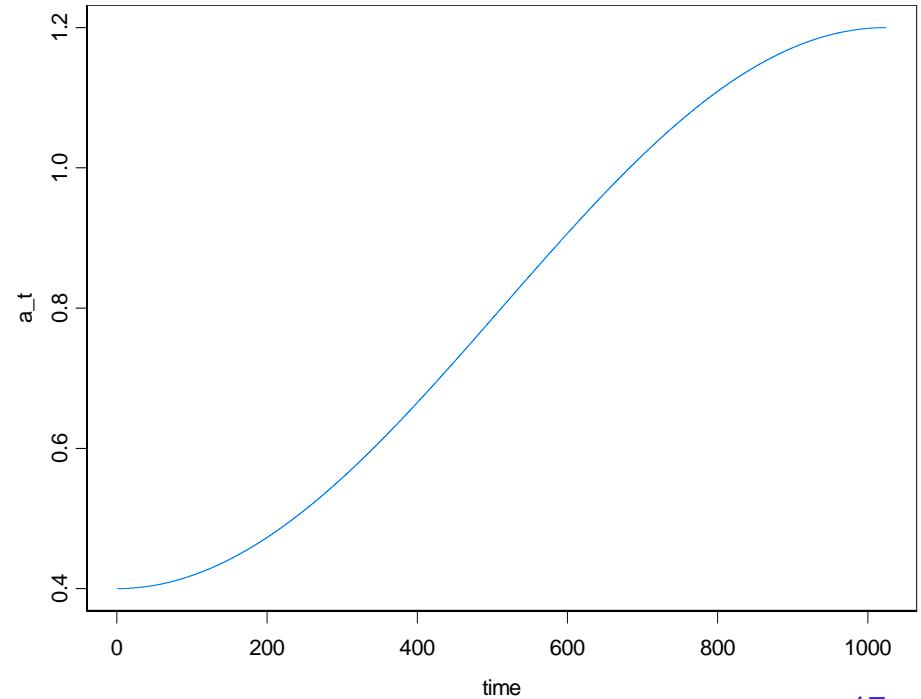
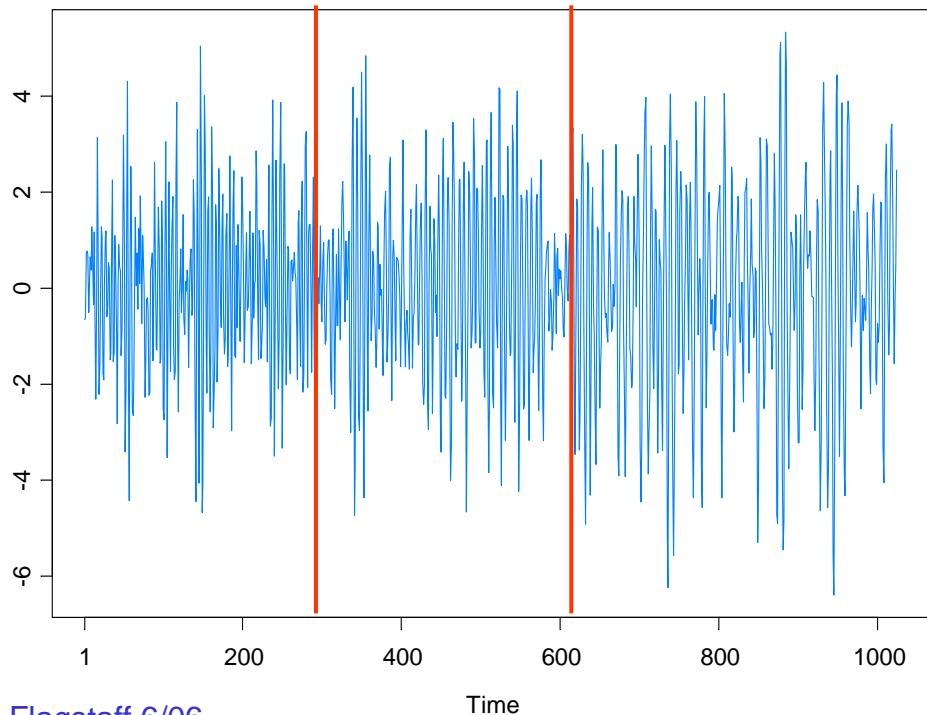


## Simulation Examples (cont)

### 2. Slowly varying AR(2) model:

$$Y_t = a_t Y_{t-1} - .81 Y_{t-2} + \varepsilon_t \quad \text{if } 1 \leq t \leq 1024$$

where  $a_t = .8[1 - 0.5\cos(\pi t/1024)]$ , and  $\{\varepsilon_t\} \sim \text{IID } N(0, 1)$ .



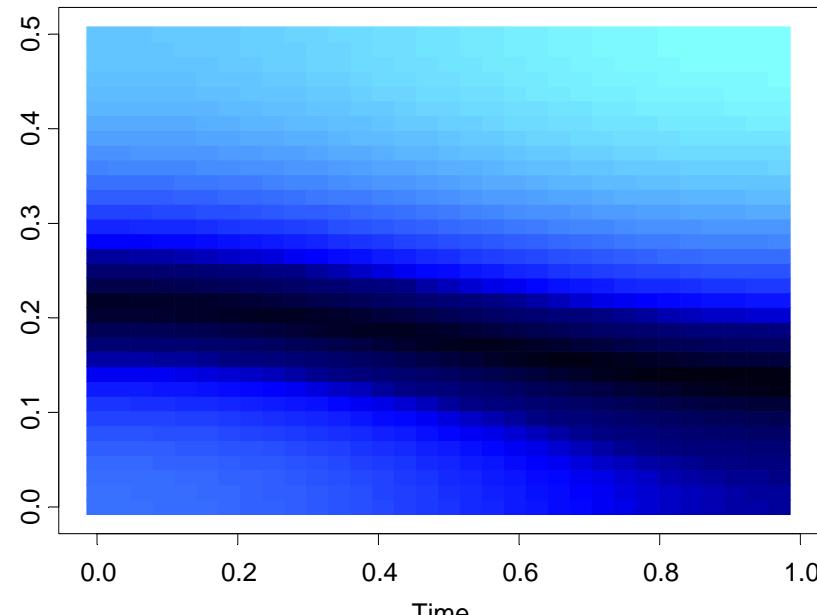
## 2. Slowly varying AR(2) (cont)

GA results: 3 pieces, breaks at  $\tau_1=293$ ,  $\tau_2=615$ . Total run time 27.45 secs

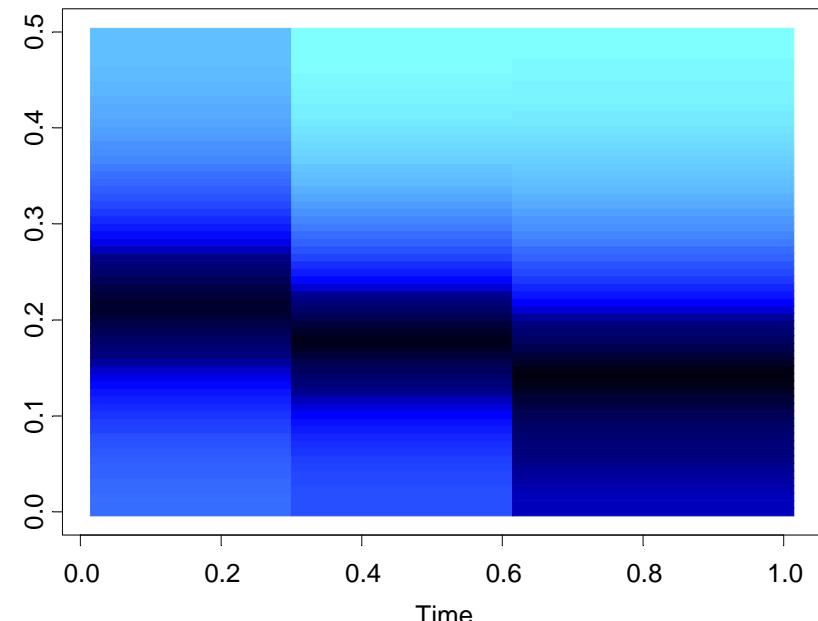
Fitted model:

	$\phi_1$	$\phi_2$	$\sigma^2$
1- 292:	.365	-0.753	1.149
293- 614:	.821	-0.790	1.176
615-1024:	1.084	-0.760	0.960

True Model

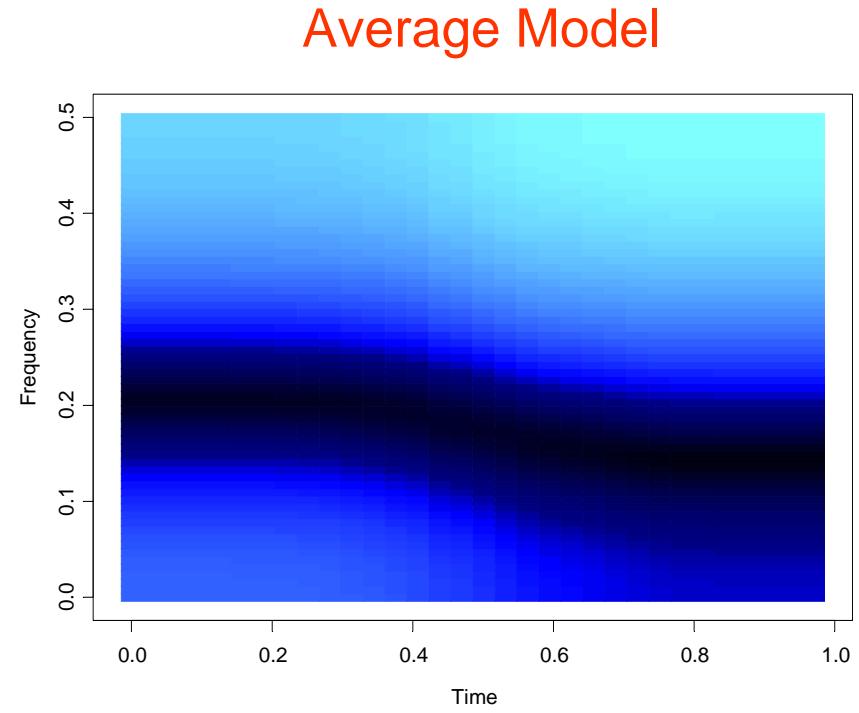
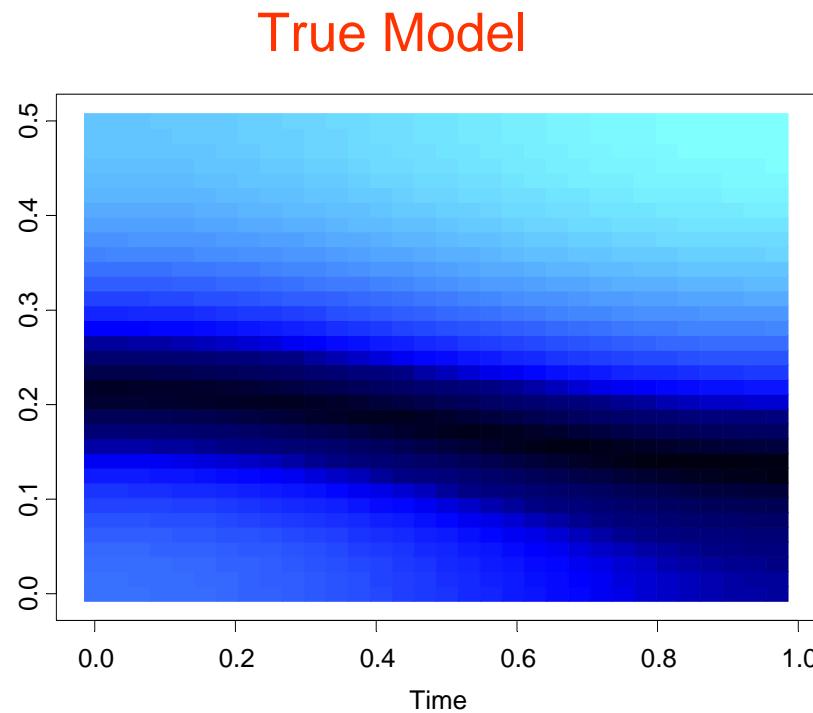


Fitted Model



## 2. Slowly varying AR(2) (cont)

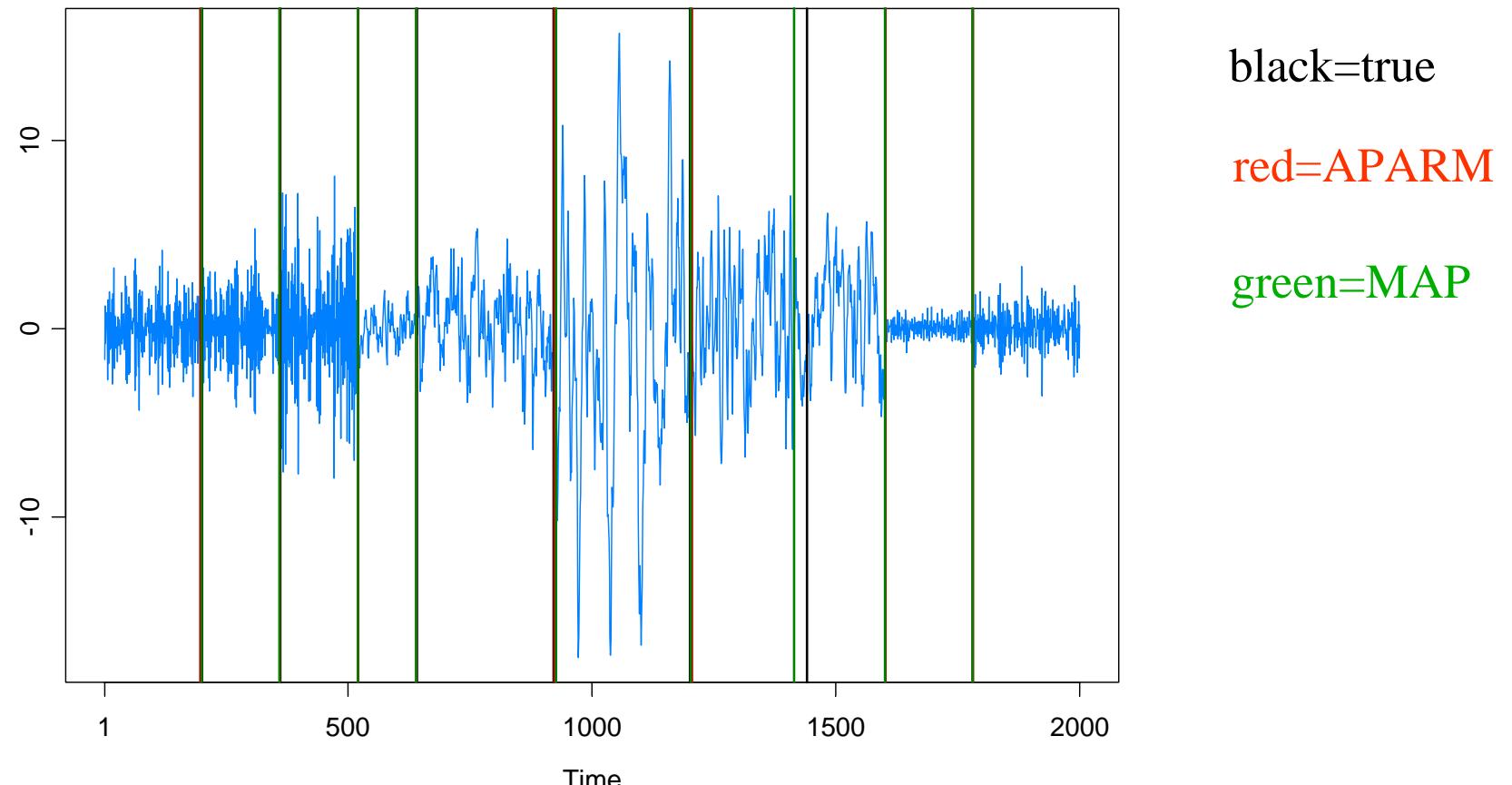
In the graph below right, we average the spectrogram over the **GA fitted models** generated from each of the 200 simulated realizations.



## Simulation Examples (cont)

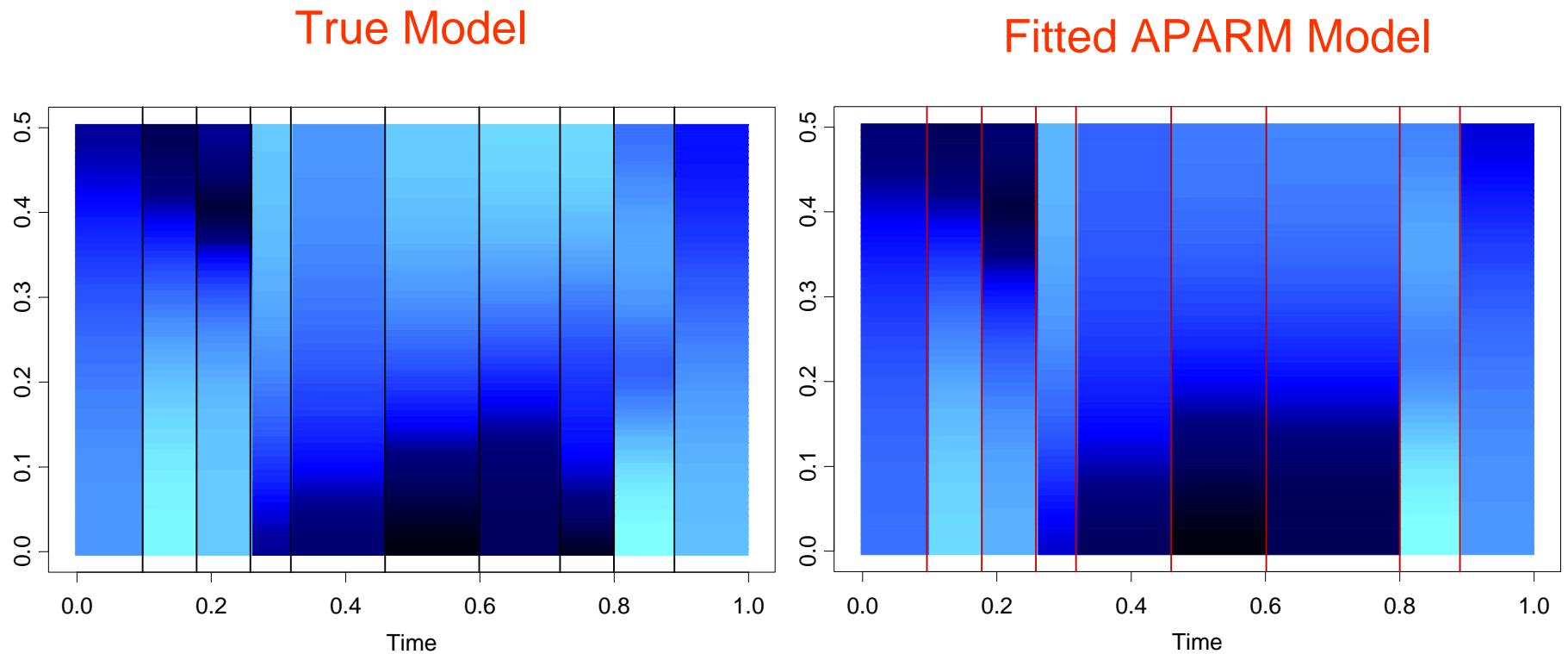
### 3. Simulated data from Fearnhead (2005):

True model has 9 changepoints



MAP est of  $m=9$  while MAP of  $m$  and changepoint locations gives  $m=8$  changepts. Plot is conditional on 9 changepoints.

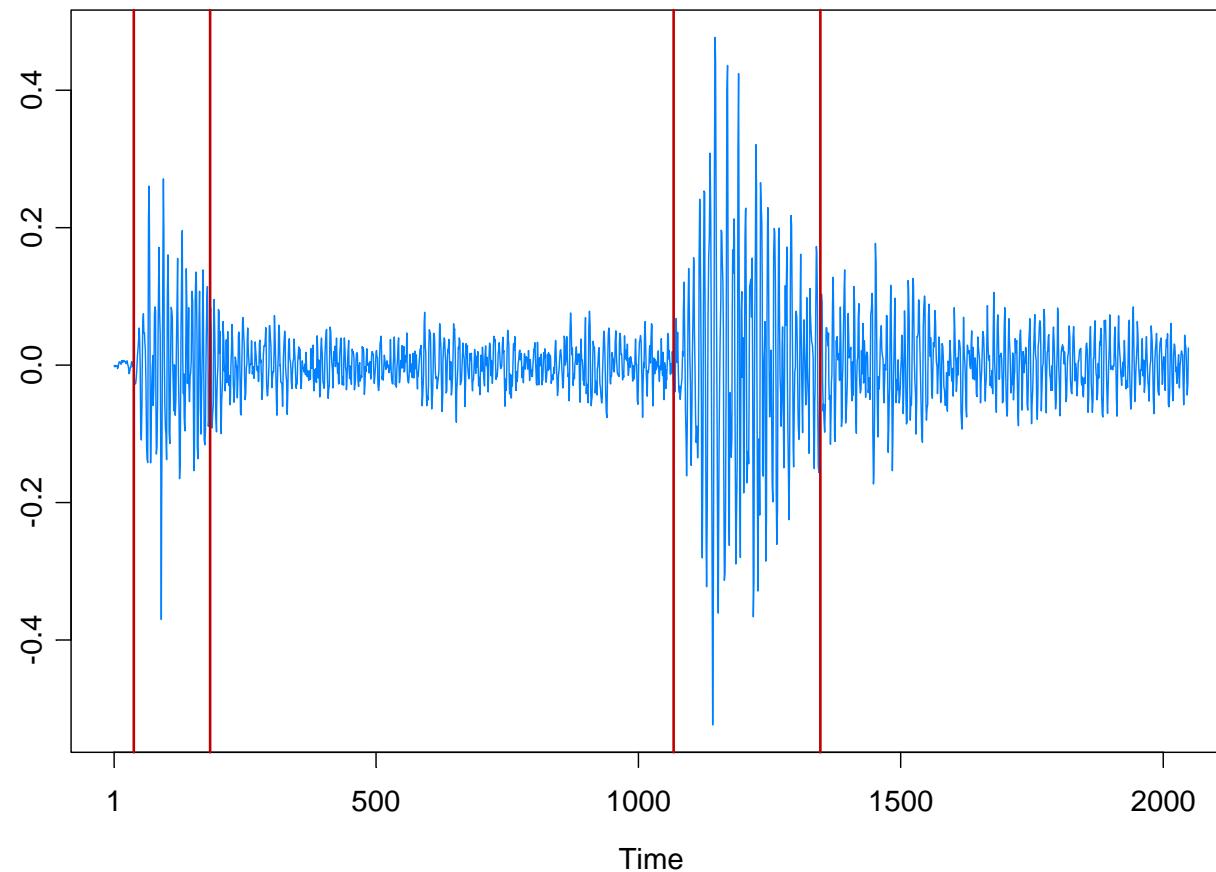
## 4. Fearnhead example



## Examples

Mine explosion seismic trace in Scandinavia: (Shumway and Stoffer 2000, Stoffer et al. 2005)

Two waves: P (primary) compression wave and S (shear) wave



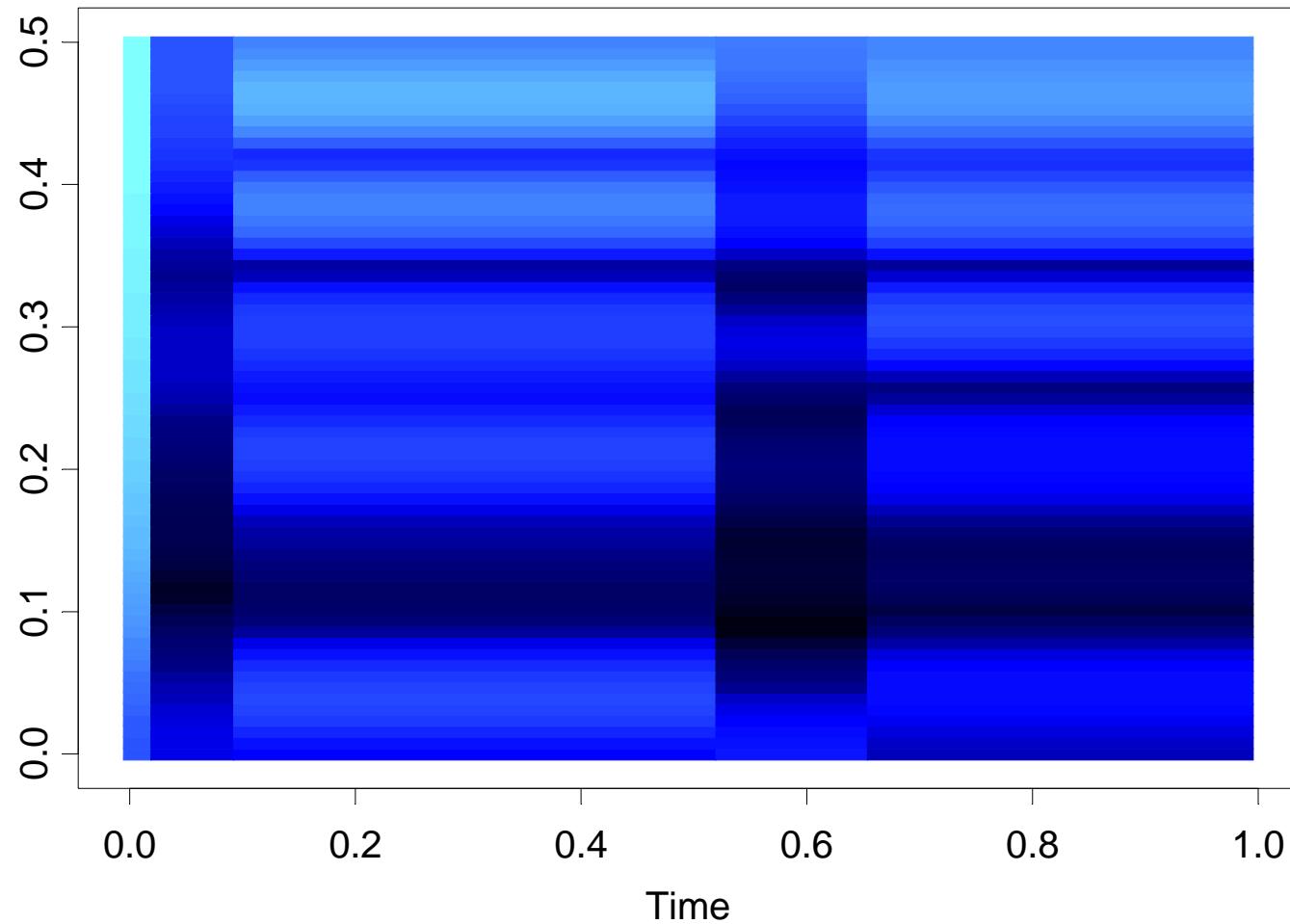
## Examples

AR orders: 1 7

17

13

15

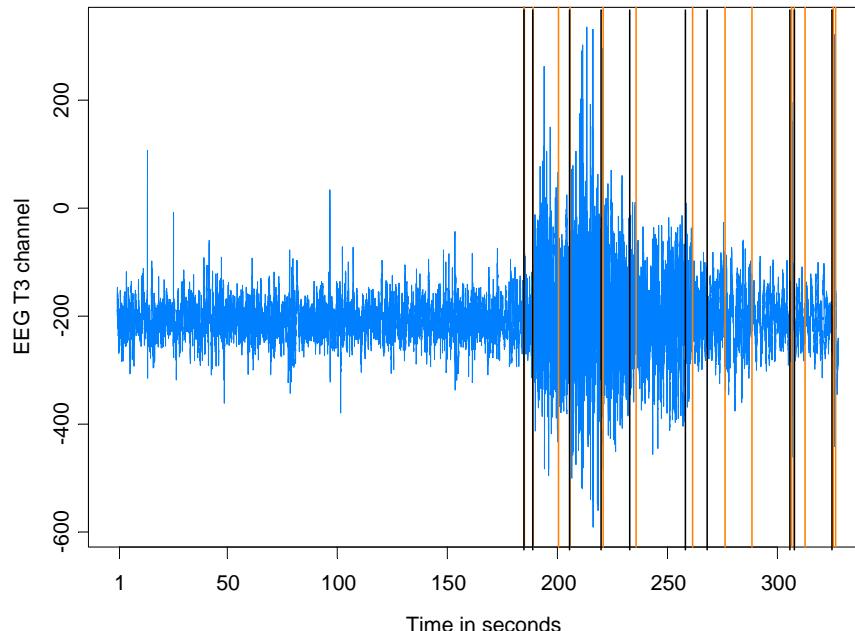


## Example: EEG Time series

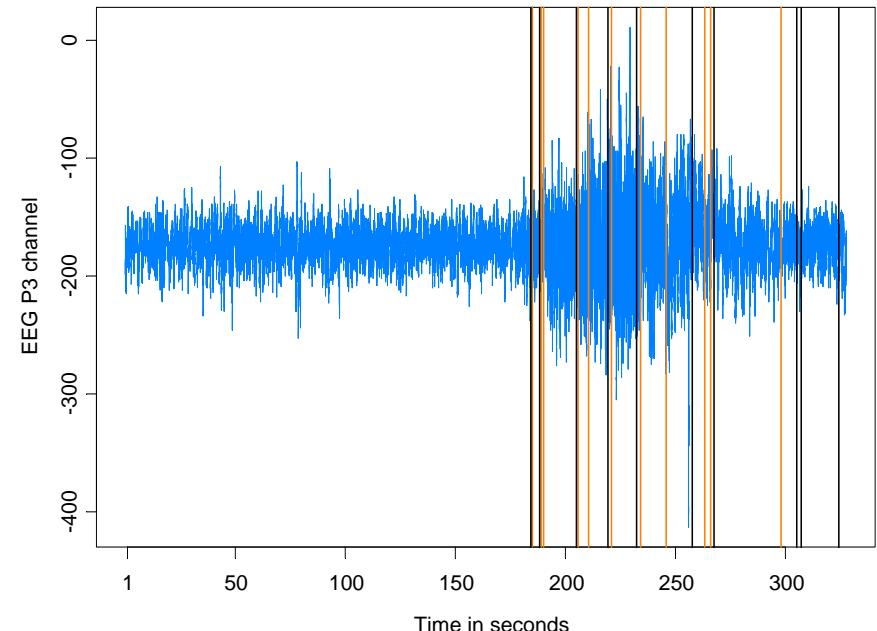
**Data:** Bivariate EEG time series at channels T3 (left temporal) and P3 (left parietal). Female subject was diagnosed with left temporal lobe epilepsy. Data collected by Dr. Beth Malow and analyzed in Ombao et al (2001). ( $n=32,768$ ; sampling rate of 100Hz). Seizure started at about 1.85 seconds.

GAGAvaniatarintatultsple4cbreakpointsforT3; 17, 2breakpoints, forP3, 4, 1

T3 Channel



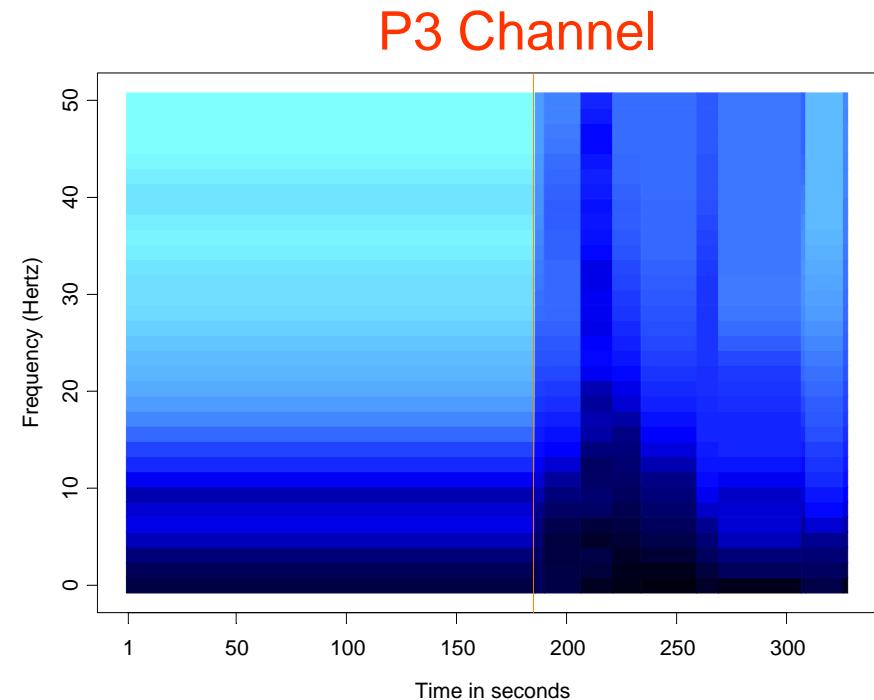
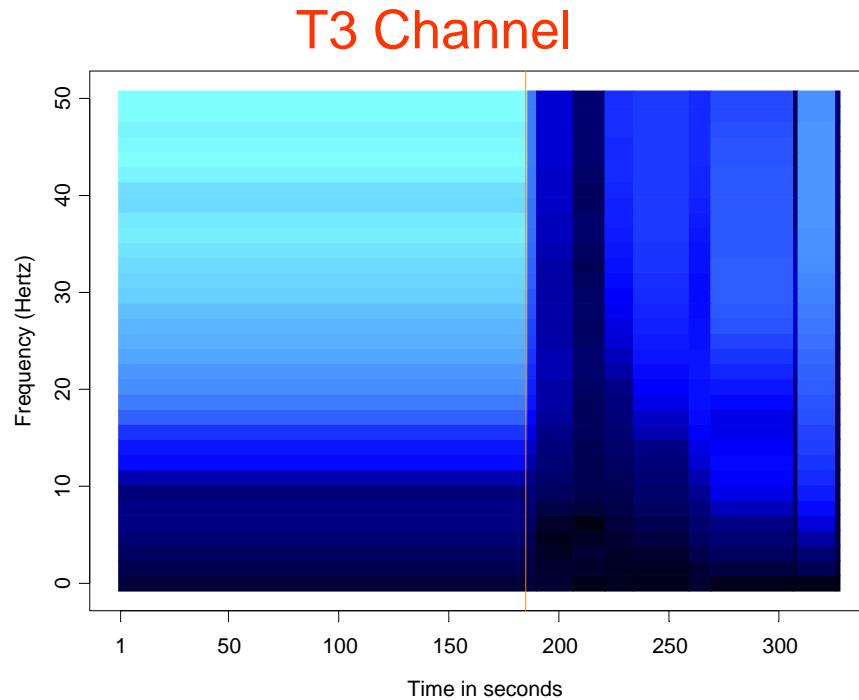
P3 Channel



## Example: EEG Time series (cont)

Remarks:

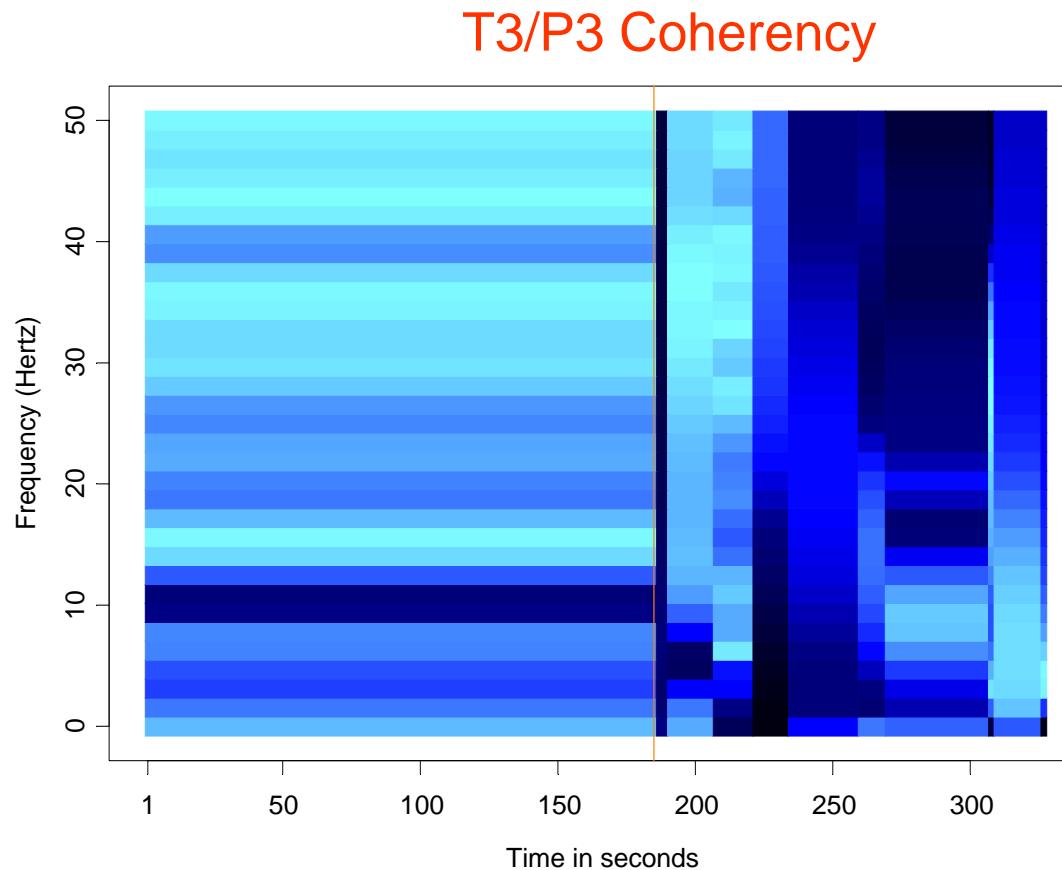
- the general conclusions of this analysis are similar to those reached in Ombao et al.
- prior to seizure, power concentrated at lower frequencies and then spread to high frequencies.
- power returned to the lower frequencies at conclusion of seizure.



## Example: EEG Time series (cont)

Remarks (cont):

- T3 and P3 strongly coherent at 9-12 Hz prior to seizure.
- strong coherence at low frequencies just after onset of seizure.
- strong coherence shifted to high frequencies during the seizure.



## Application to Parameter-Driven SS Models

### State Space Model Setup:

Observation equation:

$$p(y_t | \alpha_t) = \exp\{\alpha_t y_t - b(\alpha_t) + c(y_t)\}.$$

State equation:  $\{\alpha_t\}$  follows the piecewise AR(1) model given by

$$\alpha_t = \gamma_k + \phi_k \alpha_{t-1} + \sigma_k \varepsilon_t, \quad \text{if } \tau_{k-1} \leq t < \tau_k,$$

where  $1 = \tau_0 < \tau_1 < \dots < \tau_m < n$ , and  $\{\varepsilon_t\} \sim \text{IID } N(0,1)$ .

Parameters:

$m$  = number of break points

$\tau_k$  = location of break points

$\gamma_k$  = level in  $k^{\text{th}}$  epoch

$\phi_k$  = AR coefficients  $k^{\text{th}}$  epoch

$\sigma_k$  = scale in  $k^{\text{th}}$  epoch

## Application to Structural Breaks—(cont)

Estimation: For  $(m, \tau_1, \dots, \tau_m)$  fixed, calculate the approximate likelihood evaluated at the “MLE”, i.e.,

$$L_a(\hat{\psi}; y_n) = \frac{|G_n|^{1/2}}{(K + G_n)^{1/2}} \exp\{y_n^T \alpha^* - 1^T \{b(\alpha^*) - c(y_n)\} - (\alpha^* - \mu)^T G_n (\alpha^* - \mu)/2\},$$

where  $\hat{\psi} = (\hat{\gamma}_1, \dots, \hat{\gamma}_m, \hat{\phi}_1, \dots, \hat{\phi}_m, \hat{\sigma}_1^2, \dots, \hat{\sigma}_m^2)$  is the MLE.

Remark: The exact likelihood is given by the following formula

$$L(\psi; y_n) = L_a(\psi; y_n) Er_a(\psi),$$

where

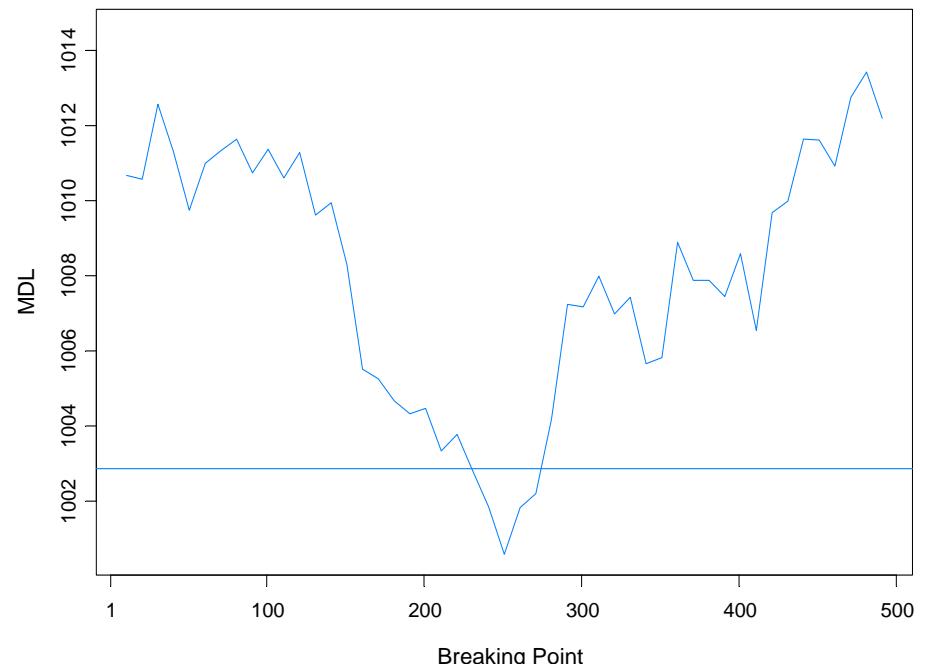
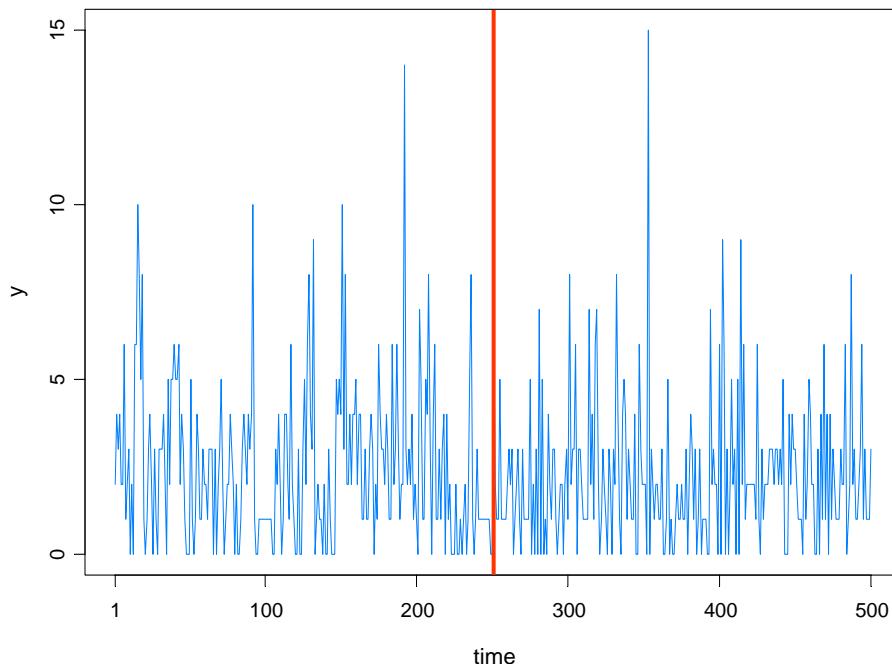
$$Er_a(\psi) = \int \exp\{R(\alpha_n; \alpha^*)\} p_a(\alpha_n | y_n; \psi) d\alpha_n.$$

It turns out that  $\log(Er_a(\psi))$  is nearly linear and can be approximated by a linear function via importance sampling,

$$e(\psi) \sim e(\hat{\psi}_{AL}) + \dot{e}(\hat{\psi}_{AL})(\psi - \hat{\psi}_{AL})$$

## Count Data Example

Model:  $Y_t | \alpha_t \sim Pois(\exp\{\beta + \alpha_t\})$ ,  $\alpha_t = \phi\alpha_{t-1} + \varepsilon_t$ ,  $\{\varepsilon_t\} \sim \text{IID } N(0, \sigma^2)$

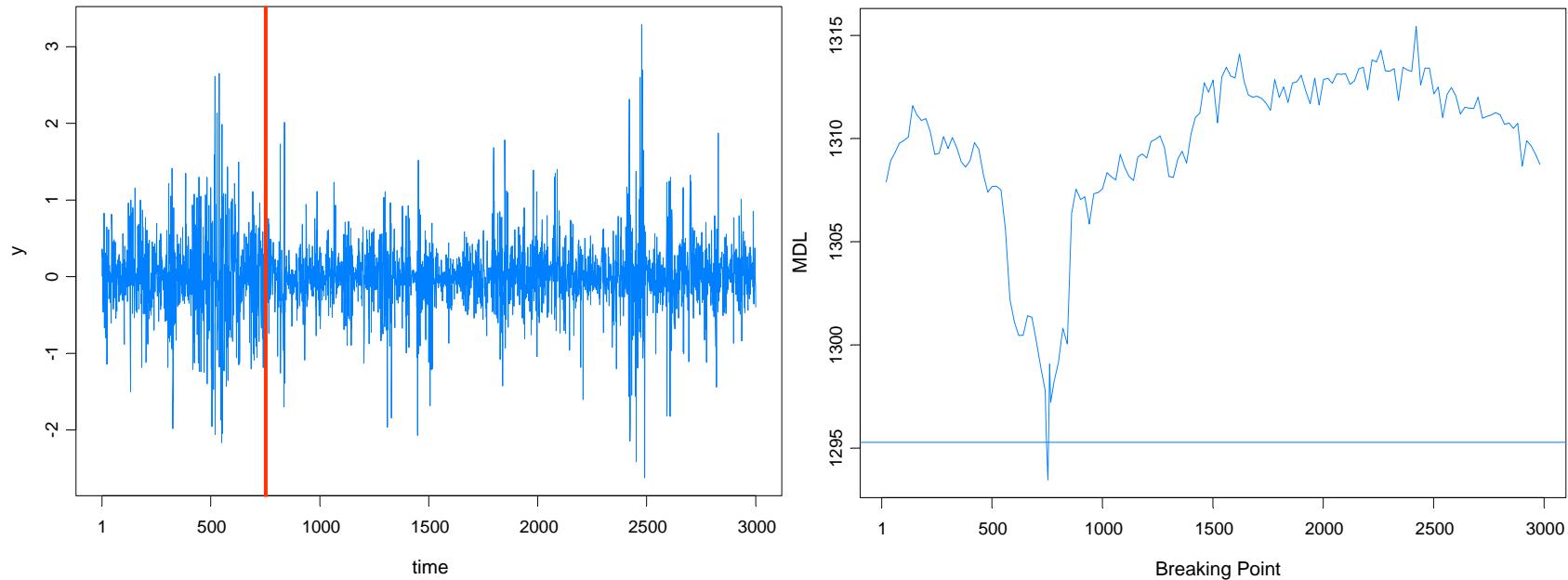


True model:

- $Y_t | \alpha_t \sim Pois(\exp\{.7 + \alpha_t\})$ ,  $\alpha_t = .5\alpha_{t-1} + \varepsilon_t$ ,  $\{\varepsilon_t\} \sim \text{IID } N(0, .3)$ ,  $t < 250$
- $Y_t | \alpha_t \sim Pois(\exp\{.7 + \alpha_t\})$ ,  $\alpha_t = -.5\alpha_{t-1} + \varepsilon_t$ ,  $\{\varepsilon_t\} \sim \text{IID } N(0, .3)$ ,  $t > 250$ .
- GA estimate 251, time 267secs

## SV Process Example

**Model:**  $Y_t | \alpha_t \sim N(0, \exp\{\alpha_t\})$ ,  $\alpha_t = \gamma + \phi \alpha_{t-1} + \varepsilon_t$ ,  $\{\varepsilon_t\} \sim \text{IID } N(0, \sigma^2)$

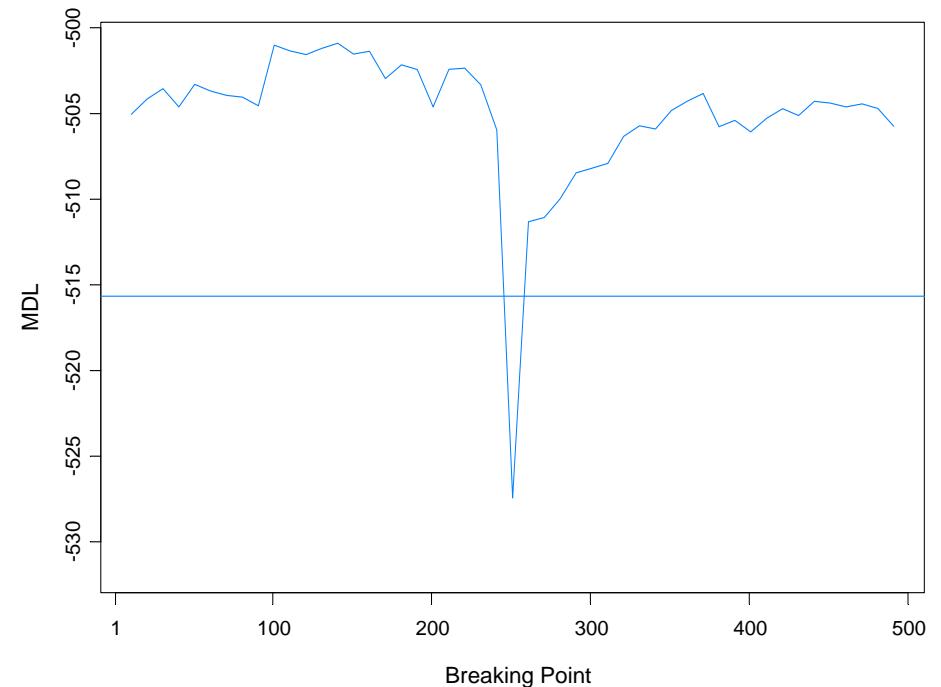
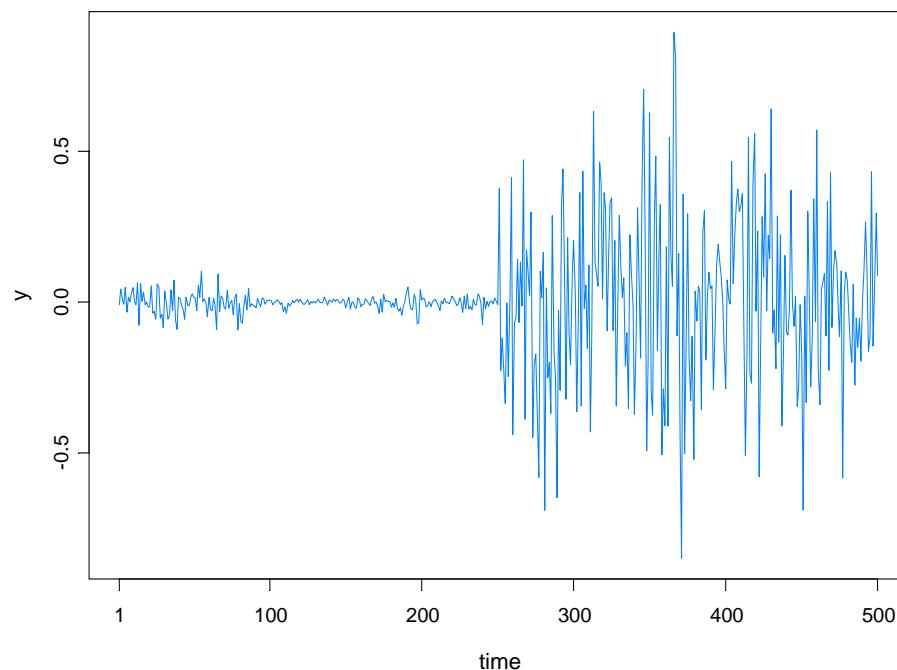


**True model:**

- $Y_t | \alpha_t \sim N(0, \exp\{\alpha_t\})$ ,  $\alpha_t = -.05 + .975\alpha_{t-1} + \varepsilon_t$ ,  $\{\varepsilon_t\} \sim \text{IID } N(0, .05)$ ,  $t \leq 750$
- $Y_t | \alpha_t \sim N(0, \exp\{\alpha_t\})$ ,  $\alpha_t = -.25 + .900\alpha_{t-1} + \varepsilon_t$ ,  $\{\varepsilon_t\} \sim \text{IID } N(0, .25)$ ,  $t > 750$ .
- GA estimate 754, time 1053 secs

## SV Process Example

**Model:**  $Y_t | \alpha_t \sim N(0, \exp\{\alpha_t\}), \quad \alpha_t = \gamma + \phi \alpha_{t-1} + \varepsilon_t, \quad \{\varepsilon_t\} \sim \text{IID } N(0, \sigma^2)$



**True model:**

- $Y_t | \alpha_t \sim N(0, \exp\{\alpha_t\}), \quad \alpha_t = -.175 + .977\alpha_{t-1} + \varepsilon_t, \quad \{\varepsilon_t\} \sim \text{IID } N(0, .1810), \quad t \leq 250$
- $Y_t | \alpha_t \sim N(0, \exp\{\alpha_t\}), \quad \alpha_t = -.010 + .996\alpha_{t-1} + \varepsilon_t, \quad \{\varepsilon_t\} \sim \text{IID } N(0, .0089), \quad t > 250$ .
- **GA estimate 251, time 269s**

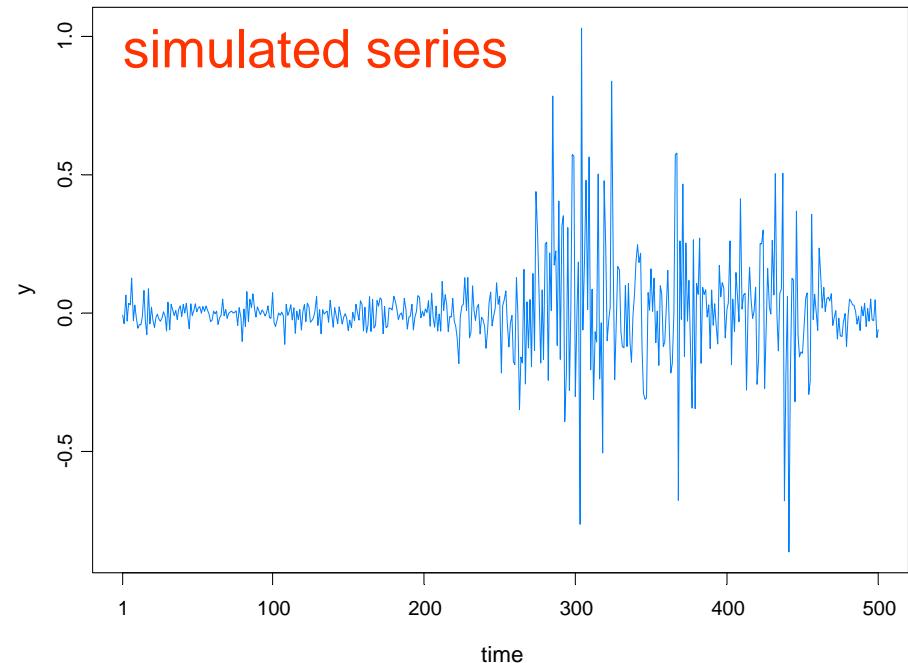
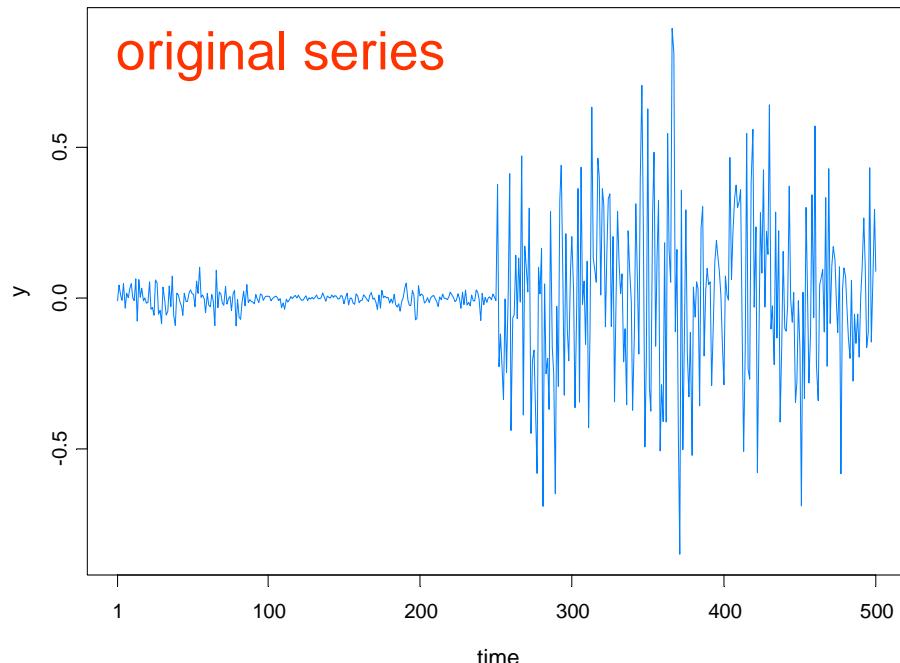
## SV Process Example-(cont)

True model:

- $Y_t | \alpha_t \sim N(0, \exp\{\alpha_t\})$ ,  $\alpha_t = -.175 + .977\alpha_{t-1} + e_t$ ,  $\{e_t\} \sim \text{IID } N(0, .1810)$ ,  $t \leq 250$
- $Y_t | \alpha_t \sim N(0, \exp\{\alpha_t\})$ ,  $\alpha_t = -.010 + .996\alpha_{t-1} + \varepsilon_t$ ,  $\{\varepsilon_t\} \sim \text{IID } N(0, .0089)$ ,  $t > 250$ .

Fitted model based on no structural break:

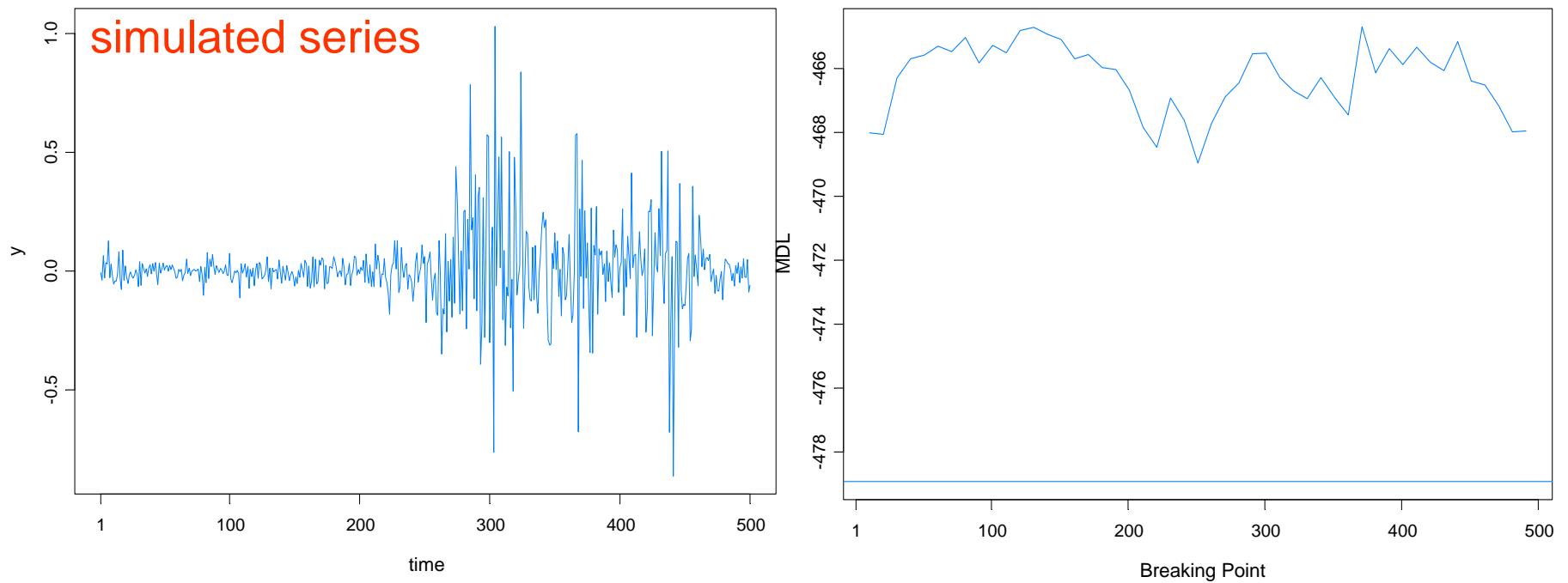
- $Y_t | \alpha_t \sim N(0, \exp\{\alpha_t\})$ ,  $\alpha_t = -.0645 + .9889\alpha_{t-1} + \varepsilon_t$ ,  $\{\varepsilon_t\} \sim \text{IID } N(0, .0935)$



## SV Process Example-(cont)

Fitted model based on no structural break:

- $Y_t | \alpha_t \sim N(0, \exp\{\alpha_t\})$ ,  $\alpha_t = -.0645 + .9889\alpha_{t-1} + \varepsilon_t$ ,  $\{\varepsilon_t\} \sim \text{IID } N(0, .0935)$



## Summary Remarks

1. *MDL* appears to be a good criterion for detecting structural breaks.
2. Optimization using a *genetic algorithm* is well suited to find a near optimal value of MDL.
3. This procedure extends easily to *multivariate* problems.
4. While estimating structural breaks for nonlinear time series models is *more challenging*, this paradigm of using *MDL together GA* holds promise for break detection in *parameter-driven* models and other nonlinear models.