# Challenges and opportunities in statistical neuroscience

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### A golden age of statistical neuroscience

Some notable recent developments:

- machine learning / statistics / optimization methods for extracting information from high-dimensional data in a computationally-tractable, systematic fashion
- computing (Moore's law, massive parallel computing)
- optical and optogenetic methods for recording from and perturbing neuronal populations, at multiple scales
- large-scale, high-density multielectrode recordings
- growing acceptance that many fundamental neuroscience questions are in fact statistics questions in disguise

#### A few grand challenges

- Optimal decoding and dimensionality reduction of large-scale multineuronal point process / count data
- Network inference from multineuronal spike train data
- Optimal control of large networks
- Hierarchical nonlinear models for input-output relationships in neuronal networks
- Robust, expressive brain-machine interfaces; brain reading and writing
- Understanding dendritic computation and location-dependent synaptic plasticity via optical imaging (statistical spatiotemporal signal processing on trees)

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#### Circuit inference via optical methods

Record large-scale calcium movie







Aim 3: Check results via photostimulation





#### Aim 1: Model-based estimation of spike rates



#### equations

$$F_t = \alpha C_t + \beta + \sigma \varepsilon_t, \qquad \varepsilon_t \stackrel{iid}{\sim} \mathcal{N}(0,1)$$
$$C_t = -(1 - \Delta/\tau)C_{t-1} + n_t$$
$$n_t \sim \text{poisson}(\lambda \Delta)$$

Note: each component here can be generalized easily.

#### Fast maximum a posteriori (MAP) estimation

Recipe: biophysical model, then likelihood, then computation.

Start by writing out the posterior:

$$\log p(C|F) = \log p(C) + \log p(F|C) + const.$$
  
= 
$$\sum_{t} \log p(C_{t+1}|C_t) + \sum_{t} \log p(F_t|C_t) + const.$$

Three basic observations:

- If  $\log p(C_{t+1}|C_t)$  and  $\log p(F_t|C_t)$  are concave, then so is  $\log p(C|F)$ .
- Hessian H of  $\log p(C|F)$  is tridiagonal:  $\log p(F_t|C_t)$  contributes a diag term, and  $\log p(C_{t+1}|C_t)$  contributes a tridiag term (Paninski et al., 2010).
- C is a linear function of n.

Newton's method: iteratively solve  $HC_{dir} = \nabla$ . Tridiagonal solver requires O(T) time. Can include nonneg constraint  $n_t \ge 0$  via log-barrier (Koyama and Paninski, 2010) — real-time deconvolution (Vogelstein et al., 2010).

More recently: constrained formulation that eliminates the need to estimate the firing rate hyperparameter (Pnevmatikakis et al 2013).

#### Markov chain Monte Carlo sampling





Spike Histogram



— Pnevmatikakis et al(2014)

#### Multineuronal case: spatiotemporal demixing





$$Y = C + \epsilon$$
  

$$C(x,t) = \sum_{i=1}^{r} s_i(x) f_i(t)$$
  

$$f_i(t+dt) = \left(1 - \frac{dt}{\tau_i}\right) f_i(t) + n_i(t)$$

Goal: infer low-rank matrix C from noisy Y. Rank r = number of visible neurons Additional structure: jumps in  $f_i(t)$  are non-negative

Locally rank-penalized convex optimization with nonnegativity constraints to infer C, followed by iterative matrix factorization under nonnegativity constraints to infer  $s_i(x), f_i(t)$  (Pnevmatikakis et al, 2013). Examples: Machado, Lacefield, Kira, Yuanjun

Idea: instead of raster scans, take randomized projections in each frame.



<sup>(</sup>from Studer et al, 2011)

Estimating C given randomized projections Y can be cast as a similar convex optimization.

Spikes



2 measurements per timestep (30x undersampling); Pnevmatikakis et al (2013)

Spikes



4 measurements per timestep (15x undersampling); Pnevmatikakis et al (2013)

Spikes



8 measurements per timestep (7.5x undersampling); Pnevmatikakis et al (2013)

#### Compressed real data

True (noisy) traces



 $\sim 2x$  undersampling; Pnevmatikakis et al (2013). superposition movie

#### Phase transitions in decoding accuracy



New tool in compressed sensing theory: "statistical dimension" (Amelunxen, Lotz, McCoy, Tropp '13).

Interesting feature of this problem: phase transition depends on pattern of spikes, not just sparsity (as in standard LASSO problem).

#### Aim 2: estimating network connectivity



Model:

$$n_{i,t} \sim Poiss(\lambda_{i,t}), \ \lambda_{i,t} = \exp(b_i + W_i n_{t-1} + stim.)$$

Coupled generalized linear model structure; concave loglikelihoods, optimization is straightforward (Paninski, 2004; Pillow et al., 2008). Easy to incorporate prior information about sparsity of connections, cell types, etc.

## The dreaded common input problem

How to distinguish direct connectivity from common input?



<sup>(</sup>from Nykamp '07)

Previous work (e.g., Vidne et al, 2012) modeled common input terms explicitly as latent variables; works well given enough a priori information, but not a general solution.

## A "shotgun sampling" approach

We can only observe K cells at a time.

Idea: don't observe the same subset of K cells throughout the experiment.

Instead, observe as many different K-subsets as possible.

Hard with multi-electrode arrays; easy with imaging approaches.

Statistics problem: how to patch together all of the estimated subnetworks?

Want to integrate over  $\{n_i(t)\}$ , but scaling to large networks is a big challenge.

## Approximate sufficient statistics in large Poisson regression network models Model:

$$n_{i,t} \sim Poiss(\lambda_{i,t}), \quad \lambda_{i,t} = \exp(b_i + W_i n_{t-1})$$

$$LL_{i} = \sum_{t} n_{i,t}(b_{i} + W_{i}n_{t-1}) - \sum_{t} \exp(b_{i} + W_{i}n_{t-1})$$

Idea: central limit theorem approximation for second term:  $W_i n_{t-1}$  is a big sum.

Dramatic simplification: approximate log-likelihood is quadratic! (Ramirez and Paninski '13)

Approximate sufficient statistics:  $E(n_t), E(n_t n_t^T), E(n_t n_{t-1}^T)$ . Can be estimated from just the observed data, or can be augmented with imputed unobserved  $\{n_{i,t}\}$ . Can further regularize  $E(n_t n_t^T)$  via sparse-inverse penalty.

#### Filling in missing spikes



Gibbs



#### Simulated "shotgun" results



K = 20% of network size; spike-and-slab priors (Keshri et al, 2013)

## Incorporating cell type structure



— simple spectral clustering methods work well to infer stochastic block model structure

#### Aim 3: Optimal control of spike timing

To test our results, we want to perturb the network at will. How can we make a neuron fire exactly when we want it to? Assume bounded inputs; otherwise problem is trivial.

Start with a simple integrate-and-soft-threshold model:

$$\lambda_t = f(V_t + h_t)$$
  
$$V_{t+dt} = V_t + dt \left(-gV_t + aI_t\right) + \sqrt{dt}\sigma\epsilon_t, \quad \epsilon_t \sim \mathcal{N}(0, 1).$$

Now we can just optimize the likelihood of the desired spike train, as a function of the input  $I_t$ , with  $I_t$  bounded.

Concave objective function over convex set of possible inputs  $I_t$ + Hessian is tridiagonal  $\implies O(T)$  optimization.

— again, can be done in real time (Ahmadian et al., 2011)... though some open challenges when  $I_t$  is high-d, spatiotemporal

#### Applications

- sensory prosthetics, e.g. retinal prosthetics
- fine-grained behavioral control

- online adaptive experimental design: choose stimuli which provide as much information about network as possible. Major problem here: updating sparse posteriors. Can speed inference significantly (Shababo, Paige et al, '13)



#### Robust point-process dimensionality reduction

Low-dimensional latent variable  $z_t$ 

Fixed matrix B mapping  $z_t$  up to the higher-dimensional neural rate space

Simplest firing rate model:

 $n_i(t) \sim Poiss(\lambda_i(t))$ 

 $\log \lambda_i(t) = B_i z_t = r_i(t)$ 

(can be generalized easily to include stimulus terms, spike history effects, etc.)

How to estimate low-d structure B without making a lot of assumptions about the dynamics (e.g., linear, Gaussian) of  $z_t$ ? Generalization of PCA: max. likelihood of spike data  $n = \{n_i(t)\}$  as function of  $R = \{r_i(t)\}$ , while minimizing rank of R:  $rank(R) \leq dim(z)$ . Nuclear norm  $\implies$  convex problem.

## Dimensionality reduction of non stationary spiking data





Method can infer B well even in highly non stationary settings — see Pfau, Pnevmatikakis et al (2013) for details.

#### Smooth, clustered activity in spinal cord

data



— Machado, Buesing et al. 2014. data; spikes and rates

#### Measuring phase tuning in single neurons



#### Mapping phase tuning across the population



<sup>—</sup> Machado, Buesing et al. 2014

#### Extension: Connectivity at the dendritic scale



## Signal processing on trees

Spatiotemporal imaging data opens an exciting window on the computations performed by single neurons, but we have to deal with noise and intermittent observations.



#### Basic paradigm: compartmental models



- write neuronal dynamics in terms of equivalent nonlinear, time-varying RC circuits
- leads to a coupled system of stochastic differential equations

#### Simplest case: Kalman filter

Dynamics and observation equations:

$$d\vec{V}/dt = A\vec{V} + \vec{\epsilon}_t$$
$$\vec{y}_t = B_t\vec{V} + \vec{\eta}_t$$

 $V_i(t) =$ voltage at compartment i

A = cable dynamics matrix: includes leak terms  $(A_{ii} = -g_l)$  and intercompartmental terms  $(A_{ij} = 0$  unless compartments are adjacent)  $B_t =$  observation matrix: point-spread function of microscope

Even this case is challenging, since  $d = \dim(\vec{V})$  is very large Standard Kalman filter:  $O(d^3)$  computation per timestep (matrix inversion)

#### Low-rank approximations

Key fact: current experimental methods provide just a few low-SNR observations per time step.

Basic idea: if dynamics are approximately linear and time-invariant, we can approximate Kalman covariance  $C_t = cov(q_t|Y_{1:t})$  as a perturbation of the marginal covariance  $C_0 + U_t D_t U_t^T$ , with  $C_0 = \lim_{t \to \infty} cov(q_t)$ .

 $C_0$  is the solution to a Lyapunov equation. It turns out that we can solve linear equations involving  $C_0$  in  $O(\dim(q))$  time via Gaussian belief propagation, using the fact that the dendrite is a tree.

The necessary recursions — i.e., updating  $U_t, D_t$  and the Kalman mean  $E(q_t|Y_{1:t})$  — involve linear manipulations of  $C_0$ , using

$$C_t = [(AC_{t-1}A^T + Q)^{-1} + B_t]^{-1}$$
  

$$C_0 + U_t D_t U_t^T = ([A(C_0 + U_{t-1}D_{t-1}U_{t-1}^T)A^T + Q]^{-1} + B_t)^{-1},$$

and can be done in  $O(\dim(q))$  time (Paninski, 2010). Generalizable to many other state-space models (Pnevmatikakis and Paninski, 2011). Examples: speckle, vertical

#### Application: synaptic locations/weights



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Including known terms:

$$d\vec{V}/dt = A\vec{V}(t) + W\vec{U}(t) + \vec{\epsilon}(t);$$

U(t) are known presynaptic spike times, and we want to detect which compartments are connected (i.e., infer the weight matrix W).

Loglikelihood is quadratic; W is a sparse vector.  $L_1$ -penalized loglikelihood can be optimized efficiently with homotopy (LARS) approach.

Total computation time: O(dTk); d = # compartments, T = # timesteps, k = # nonzero weights.

#### Example: real neural geometry



700 timesteps observed; 40 random compartments (of >2000) observed per timestep. Zecevic data

Compressed sensing measurements improve accuracy further (Pakman et al 2013).

#### Conclusions

- Modern statistical approaches provide flexible, powerful methods for answering key questions in neuroscience many of these problems are statistics problems in disguise
- Close relationships between biophysics, statistical modeling, and experimental design
- Modern optimization methods make computations very tractable; suitable for closed-loop experiments
- Experimental methods progressing rapidly; many new challenges and opportunities for breakthroughs based on statistical ideas

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