

Fitting and understanding multilevel (hierarchical) models

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Making more use of existing information

- ▶ **The problem:** not enough data to estimate effects with confidence
- ▶ **The solution:** make your studies *broader* and *deeper*
 - Broader: extend to other countries, other years, other outcomes, ...
 - Deeper: inferences for individual states, demographic subgroups, components of outcomes, ...
- ▶ **The solution:** multilevel modeling

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 - ▶ No such thing as "too many predictors"

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Fitting and understanding multilevel models

- ▶ The effectiveness of multilevel models
- ▶ Multilevel models in unexpected places
- ▶ Multilevel models as a way of life
- ▶ collaborators:
 - ▶ Iain Pardoe, Dept of Decision Sciences, University of Oregon
 - ▶ David Park, Dept of Political Science, Washington University
 - ▶ Joe Bafumi, Dept of Political Science, Columbia University
 - ▶ Boris Shor, Dept of Political Science, Columbia University
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- ▶ The effectiveness of multilevel models
 - ▶ State-level opinions from national polls
(crossed multilevel modeling and poststratification)
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National opinion trends



State-level opinion trends

- ▶ Goal: estimating time series within each state
- ▶ One poll at a time: small-area estimation
- ▶ It works! Validated for pre-election polls
- ▶ Combining surveys: model for parallel time series
- ▶ Multilevel modeling + poststratification
- ▶ Poststratification cells: sex \times ethnicity \times age \times education \times state

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Multilevel modeling of opinions

- ▶ Logistic regression: $\Pr(y_i = 1) = \text{logit}^{-1}((X\beta)_i)$
- ▶ X includes demographic and geographic predictors
- ▶ Group-level model for the 16 age \times education predictors
- ▶ Group-level model for the 50 state predictors
- ▶ Bayesian inference, summarize by posterior simulations of β :

Simulation	θ_1	\dots	θ_{75}
1	**	\dots	**
\vdots	\vdots	\dots	\vdots
1000	**	\dots	**

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- ▶ Several overlapping “hierarchies”

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Poststratification to estimate state opinions

- ▶ Implied inference for $\theta_j = \text{logit}^{-1}(X\beta)$ in each of 3264 cells j (e.g., black female, age 18–29, college graduate, Massachusetts)
- ▶ Poststratification

Weight each cell by $W_j = N_j / \sum_j N_j$ (where N_j is the number of people in cell j)
 $\hat{\theta} = \sum_j W_j \theta_j$ (weighted average of cell means)
1000 simulation draws propagate to uncertainty for each θ_j

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 - ▶ Within each state s , average over 64 cells:
$$\sum_{j \in s} N_j \theta_j / \sum_{j \in s} N_j$$
 - ▶ N_j = population in cell j (from Census)
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CBS/New York Times pre-election polls from 1988

- ▶ Validation study: fit model on poll data and compare to election results
- ▶ Competing estimates:
 - ▶ Nonpooling: separate estimates for each poll
 - ▶ Complete pooling: no state predictors
 - ▶ Hierarchical model and poststratify
- ▶ Mean absolute state errors:

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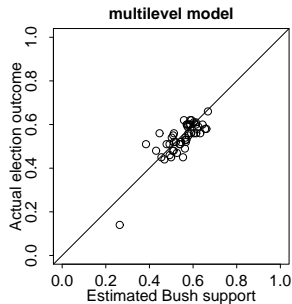
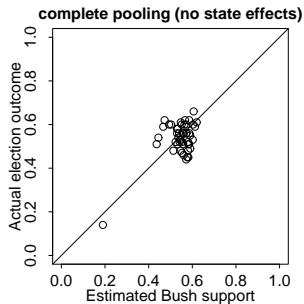
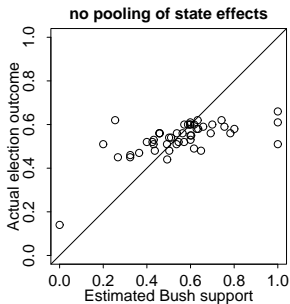
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Validation study: comparison of state errors

1988 election outcome vs. poll estimate



Multilevel models always

- ▶ Anything worth doing is worth doing repeatedly
- ▶ A “method” is any procedure applied more than once
- ▶ City planning

Outward expansion, estimating a model for other categories, other years, other outcomes, ...
Inference for individual states, demographic subgroups, components of data, ...

- ▶ “Frequentist” statistical theory of repeated inferences

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Incumbency advantage in U.S. House elections

- ▶ Regression approach (Gelman and King, 1990):
 - ▶ For any year, compare districts with and without incs running
 - ▶ Control for vote in previous election
 - ▶ Control for incumbent *party*
 - ▶ $v_{it} = \beta_0 + \beta_1 v_{i,t-1} + \beta_2 P_{it} + \psi l_{it} + \epsilon_{it}$
- ▶ Other estimates (sophomore surge, etc.) have selection bias

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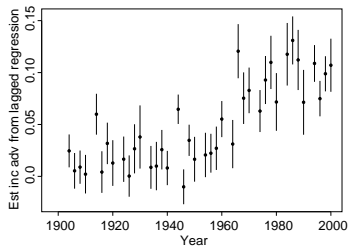
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Estimated incumbency advantage from lagged regressions



Can we do better?

- ▶ Regression estimate: $v_{it} = \beta_0 + \beta_1 v_{i,t-1} + \beta_2 P_{it} + \psi l_{it} + \epsilon_{it}$
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- ▶ “Statistics” problem: the model doesn’t fit the data
- ▶ We’ll show pictures of the model not fitting
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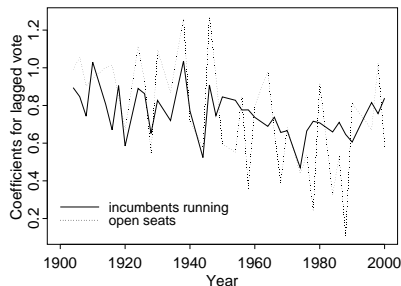
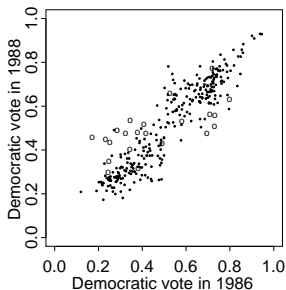
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Model misfit

Under the model, parallel lines are fitted to the circles (open seats) and dots (incs running for reelection)



Multilevel model

- ▶ for $t = 1, 2$: $v_{it} = 0.5 + \delta_t + \alpha_i + \phi_{it}l_{it} + \epsilon_{it}$
 - ▶ $\delta_2 - \delta_1$ is the national vote swing
 - ▶ α_i is the “normal vote” for district i : mean 0, sd σ_α .
 - ▶ ϕ_{it} is the inc advantage in district i at time t : mean ψ , sd σ_ϕ
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- ▶ Candidate-level incumbency effects:

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Fitting the multilevel model

- ▶ Bayesian inference
- ▶ Linear parameters: national vote swings, district effects, incumbency effects
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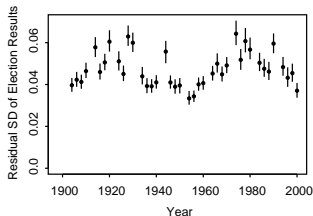
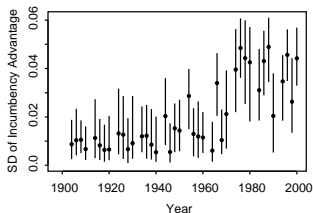
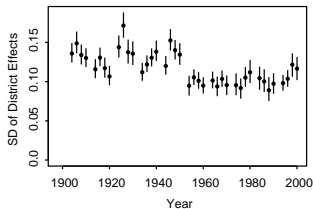
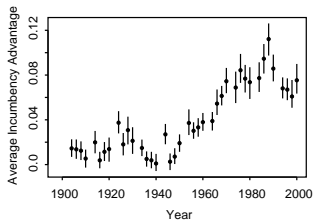
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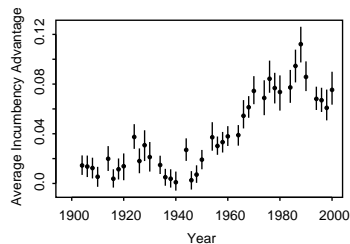
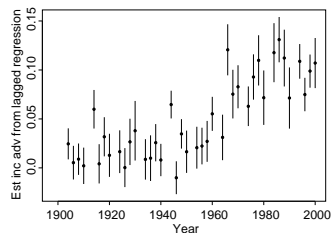
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Estimated incumbency advantage and its variation

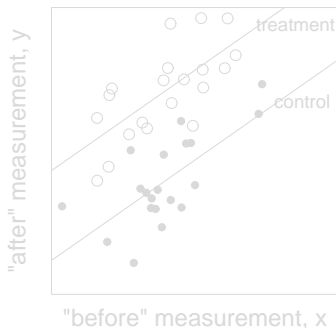


Compare old and new estimates



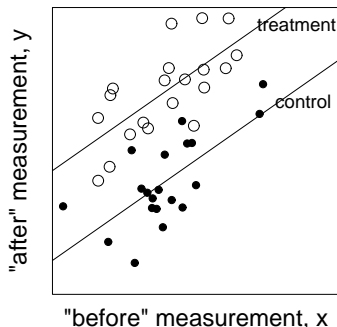
No-interaction model

- ▶ Before-after data with treatment and control groups
- ▶ Default model: constant treatment effects
 - ▶ Fisher's classical null hyp: effect is zero for all cases
 - ▶ Regression model: $y_i = T_i\theta + X_i\beta + \epsilon_i$



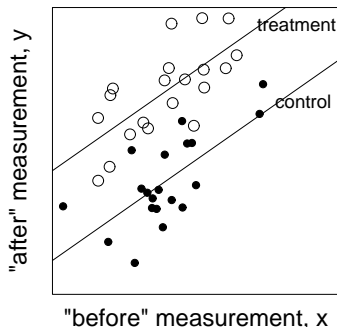
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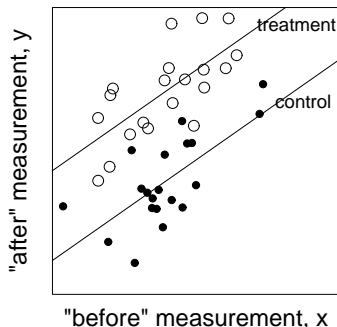
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Actual data show interactions

- ▶ Treatment interacts with “before” measurement
- ▶ Before-after correlation is higher for *controls* than for *treated* units
- ▶ Examples

• [The 2008 US Presidential election](#) (Gelman, Su, & Yajima, 2010)
• [An experiment with gender, post-test, and time](#) (Gelman, Su, & Yajima, 2010)
• [Longitudinal data on the effect of a health intervention](#) (Gelman, Su, & Yajima, 2010)

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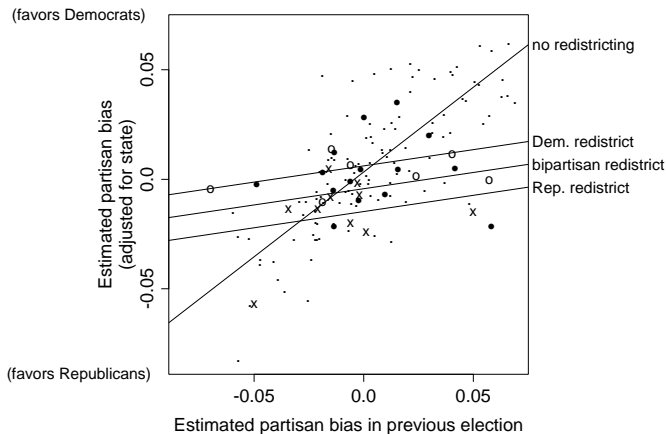
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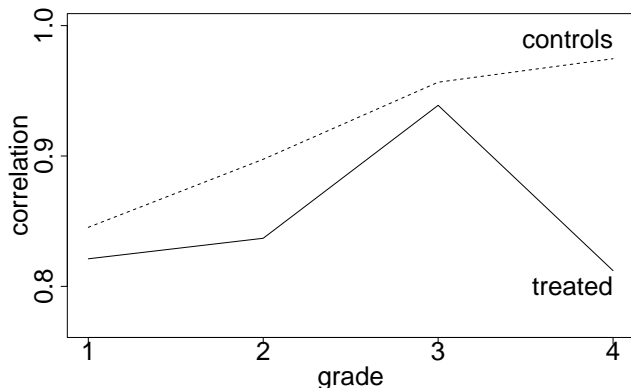
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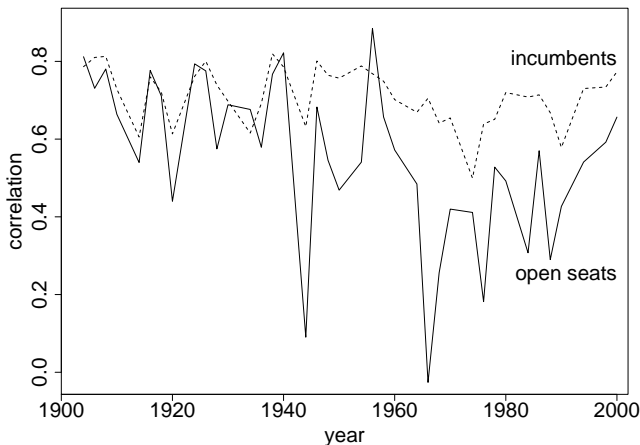
Observational study of legislative redistricting before-after data



Experiment: correlation between pre-test and post-test data for controls and for treated units



Correlation between two successive Congressional elections for incumbents running (controls) and open seats (treated)



Interactions as variance components

Unit-level “error term” η_i

- ▶ For control units, η_i persists from time 1 to time 2
- ▶ For treatment units, η_i changes:
 - ▶ Subtractive treatment error (negative incidence 1)
 - ▶ Additive treatment error (positive incidence 2)
 - ▶ Replacement treatment error
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- ▶ Building and fitting multilevel models
- ▶ Displaying and summarizing inferences

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Redundant parameterization

- ▶ Data model: $\Pr(y_i = 1) = \text{logit}^{-1} \left(\beta^0 + \beta_{\text{age}(i)}^{\text{age}} + \beta_{\text{state}(i)}^{\text{state}} \right)$
- ▶ Usual model for the coefficients:

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$$\begin{aligned}\beta_j^{\text{age}} &\sim N(\mu_{\text{age}}, \sigma_{\text{age}}^2), \quad \text{for } j = 1, \dots, 4 \\ \beta_j^{\text{state}} &\sim N(\mu_{\text{state}}, \sigma_{\text{state}}^2), \quad \text{for } j = 1, \dots, 50\end{aligned}$$

- ▶ Why add the redundant $\mu_{\text{age}}, \mu_{\text{state}}$?

▶ Iterative algorithm moves more smoothly

Redundant parameterization

- ▶ Data model: $\Pr(y_i = 1) = \text{logit}^{-1} \left(\beta^0 + \beta_{\text{age}(i)}^{\text{age}} + \beta_{\text{state}(i)}^{\text{state}} \right)$
- ▶ Usual model for the coefficients:

$$\begin{aligned}\beta_j^{\text{age}} &\sim N(0, \sigma_{\text{age}}^2), \quad \text{for } j = 1, \dots, 4 \\ \beta_j^{\text{state}} &\sim N(0, \sigma_{\text{state}}^2), \quad \text{for } j = 1, \dots, 50\end{aligned}$$

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Redundant additive parameterization

► Model

$$\begin{aligned}\Pr(y_i = 1) &= \text{logit}^{-1} \left(\beta^0 + \beta_{\text{age}(i)}^{\text{age}} + \beta_{\text{state}(i)}^{\text{state}} \right) \\ \beta_j^{\text{age}} &\sim \text{N}(\mu_{\text{age}}, \sigma_{\text{age}}^2), \quad \text{for } j = 1, \dots, 4 \\ \beta_j^{\text{state}} &\sim \text{N}(\mu_{\text{state}}, \sigma_{\text{state}}^2), \quad \text{for } j = 1, \dots, 50\end{aligned}$$

► Identify using centered parameters:

$$\begin{aligned}\tilde{\beta}_j^{\text{age}} &= \beta_j^{\text{age}} - \bar{\beta}^{\text{age}}, \quad \text{for } j = 1, \dots, 4 \\ \tilde{\beta}_j^{\text{state}} &= \beta_j^{\text{state}} - \bar{\beta}^{\text{state}}, \quad \text{for } j = 1, \dots, 50\end{aligned}$$

► Redefine the constant term:

$$\tilde{\beta}^0 = \beta^0 + \bar{\beta}^{\text{age}} + \bar{\beta}^{\text{state}}$$

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Redundant multiplicative parameterization

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$$\Pr(y_i = 1) = \text{logit}^{-1} \left(\beta^0 + \xi^{\text{age}} \beta_{\text{age}(i)}^{\text{age}} + \xi^{\text{state}} \beta_{\text{state}(i)}^{\text{state}} \right)$$

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Weakly informative prior distribution for the multilevel variance parameter

- ▶ Redundant multiplicative parameterization:

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- ▶ Separate prior distributions on the ξ and σ parameters:
 - ▶ Normal on ξ
 - ▶ Inverse-gamma on σ^2
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Displaying and summarizing inferences

- ▶ Displaying parameters in groups rather than as a long list
- ▶ Average predictive effects
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Displaying and summarizing inferences

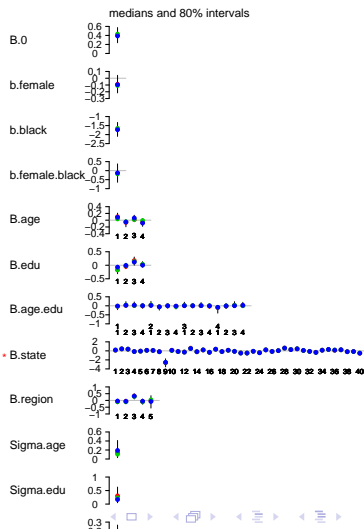
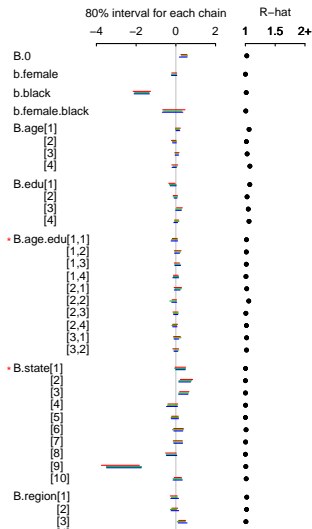
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Raw display of inference

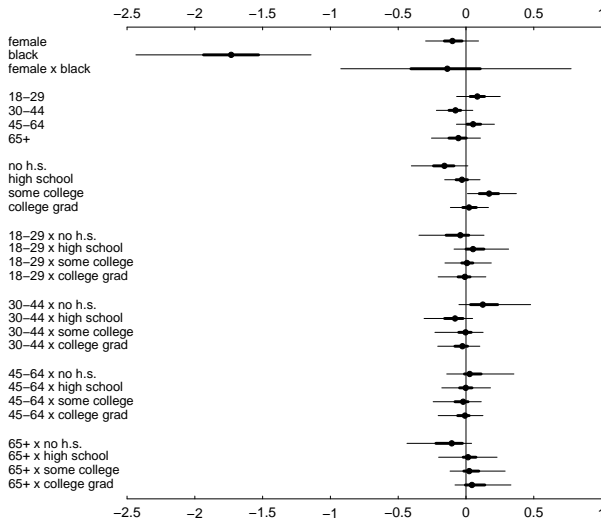
	mean	sd	2.5%	25%	50%	75%	97.5%	Rhat	n.eff
B.0	0.402	0.147	0.044	0.326	0.413	0.499	0.652	1.024	110
b.female	-0.094	0.102	-0.283	-0.162	-0.095	-0.034	0.107	1.001	1000
b.black	-1.701	0.305	-2.323	-1.910	-1.691	-1.486	-1.152	1.014	500
b.female.black	-0.143	0.393	-0.834	-0.383	-0.155	0.104	0.620	1.007	1000
B.age[1]	0.084	0.088	-0.053	0.012	0.075	0.140	0.277	1.062	45
B.age[2]	-0.072	0.087	-0.260	-0.121	-0.054	-0.006	0.052	1.017	190
B.age[3]	0.044	0.077	-0.105	-0.007	0.038	0.095	0.203	1.029	130
B.age[4]	-0.057	0.096	-0.265	-0.115	-0.052	0.001	0.133	1.076	32
B.edu[1]	-0.148	0.131	-0.436	-0.241	-0.137	-0.044	0.053	1.074	31
B.edu[2]	-0.022	0.082	-0.182	-0.069	-0.021	0.025	0.152	1.028	160
B.edu[3]	0.148	0.112	-0.032	0.065	0.142	0.228	0.370	1.049	45
B.edu[4]	0.023	0.090	-0.170	-0.030	0.015	0.074	0.224	1.061	37
B.age.edu[1,1]	-0.044	0.133	-0.363	-0.104	-0.019	0.025	0.170	1.018	1000
B.age.edu[1,2]	0.059	0.123	-0.153	-0.011	0.032	0.118	0.353	1.016	580
B.age.edu[1,3]	0.049	0.124	-0.146	-0.023	0.022	0.104	0.349	1.015	280
B.age.edu[1,4]	0.001	0.116	-0.237	-0.061	0.000	0.052	0.280	1.010	1000
B.age.edu[2,1]	0.066	0.152	-0.208	-0.008	0.032	0.124	0.449	1.022	140
B.age.edu[2,2]	-0.081	0.127	-0.407	-0.137	-0.055	0.001	0.094	1.057	120
B.age.edu[2,3]	-0.004	0.102	-0.226	-0.048	0.000	0.041	0.215	1.008	940
B.age.edu[2,4]	-0.042	0.108	-0.282	-0.100	-0.026	0.014	0.157	1.017	170
B.age.edu[3,1]	0.034	0.135	-0.215	-0.030	0.009	0.091	0.361	1.021	230
B.age.edu[3,2]	0.007	0.102	-0.213	-0.039	0.003	0.052	0.220	1.019	610
B.age.edu[3,3]	0.033	0.130	-0.215	-0.029	0.009	0.076	0.410	1.080	61
B.age.edu[3,4]	-0.009	0.109	-0.236	-0.064	-0.005	0.043	0.214	1.024	150
B.age.edu[4,1]	-0.141	0.190	-0.672	-0.224	-0.086	-0.003	0.100	1.036	270
B.age.edu[4,2]	-0.014	0.119	-0.280	-0.059	-0.008	0.033	0.239	1.017	240
B.age.edu[4,3]	0.046	0.132	-0.192	-0.024	0.019	0.108	0.332	1.010	210
B.age.edu[4,4]	0.042	0.142	-0.193	-0.022	0.016	0.095	0.377	1.015	160
B.state[1]	0.201	0.211	-0.121	0.047	0.172	0.326	0.646	1.002	860

Raw graphical display

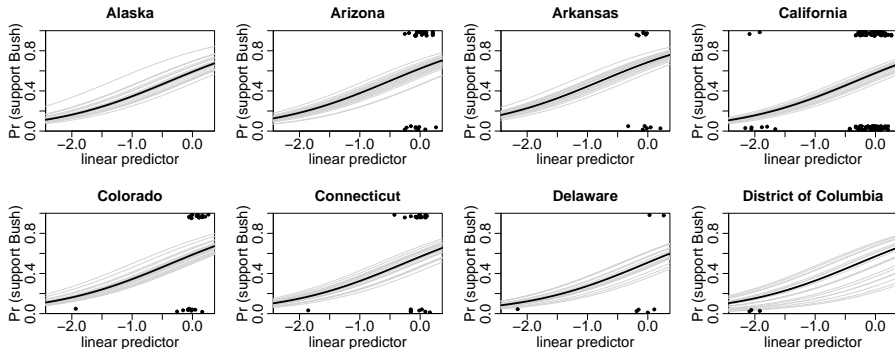
Bugs model at "C:/books/multilevel/election88/model4.bug", 3 chains, each with 2001 iterations



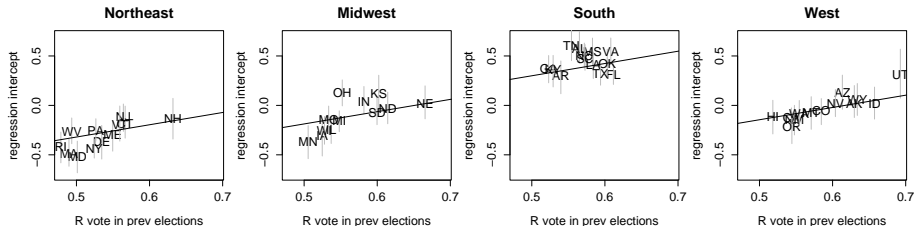
Better graphical display 1: demographics



Better graphical display 2: within states

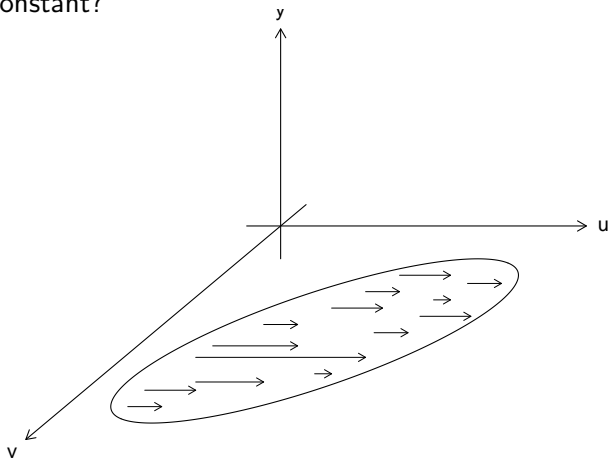


Better graphical display 3: between states



Average predictive effects

- What is $E(y \mid x_1 = \text{high}) - E(y \mid x_1 = \text{low})$, with all other x 's held constant?



Average predictive effects

- ▶ What is $E(y \mid x_1 = \text{high}) - E(y \mid x_1 = \text{low})$, with all other x 's held constant?
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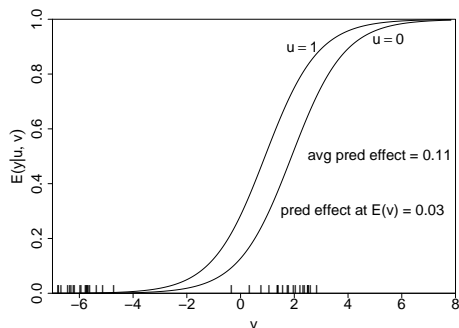
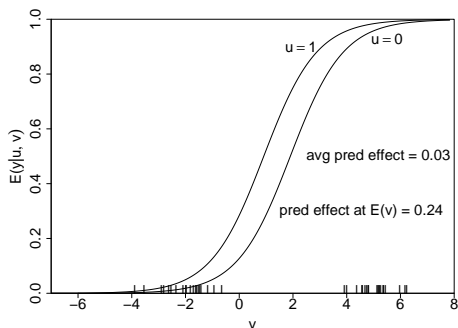
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APE: why you can't just use a central value of x



Understanding sources of variation

- ▶ Generalization of R^2 (explained variance), defined at each level of the model
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- ▶ Analysis of variance
- ▶ Open question: how to construct models with deep interaction structures?

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Conclusions

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