

Learning about social and political polarization using “How many X’s do you know” surveys

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Overview

- ▶ Social and political polarization
- ▶ “How many X’s do you know” surveys
- ▶ 3 models and Bayesian inference
- ▶ Our research plan
- ▶ collaborators:
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 - ▶ Matt Salganik, Dept of Sociology, Columbia University
 - ▶ Tom DiPrete, Dept of Sociology, Columbia University
 - ▶ Julien Teitler, School of Social Work, Columbia University
 - ▶ Jouni Kerman, Dept of Statistics, Columbia University
 - ▶ Peter Killworth and Chris McCarty shared their survey data

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Increasing social/economic heterogeneity in U.S. since 1950s?

► Social polarization:

- More variety in domestic arrangements
- Greater income inequality
- We tend to know people of similar social class to ourselves
- Counter-trend: more interracial marriages

► Decline in social capital:

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 - ▶ Later marriage, fewer children
 - ▶ "Bowling alone" (Putnam)
 - ▶ Less involvement in community organizations

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Past work studying polarization using surveys

- ▶ Lots and lots has been done; this is an incomplete review
- ▶ Social polarization, social capital:
 - ▶ Census data on family characteristics (Cherlin, Mayer, Held, ...)
 - ▶ GSS, NES questions on values (White, Brooks, ...)
 - ▶ General Social Survey questions about how close communities are (DiPrete, ...)
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 - ▶ NES and commercial polls (Page and Shapiro, Bauman, ...)

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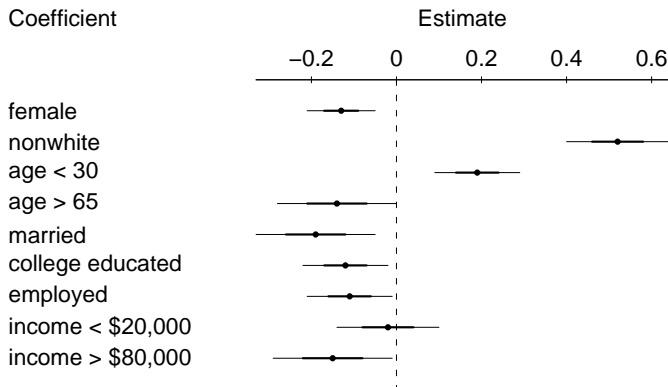
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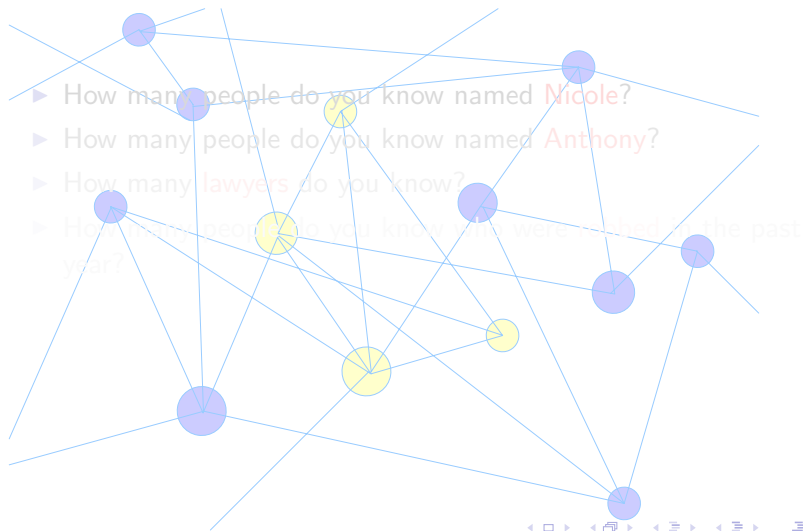
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Example analysis: regression of residuals for "How many prisoners do you know?"



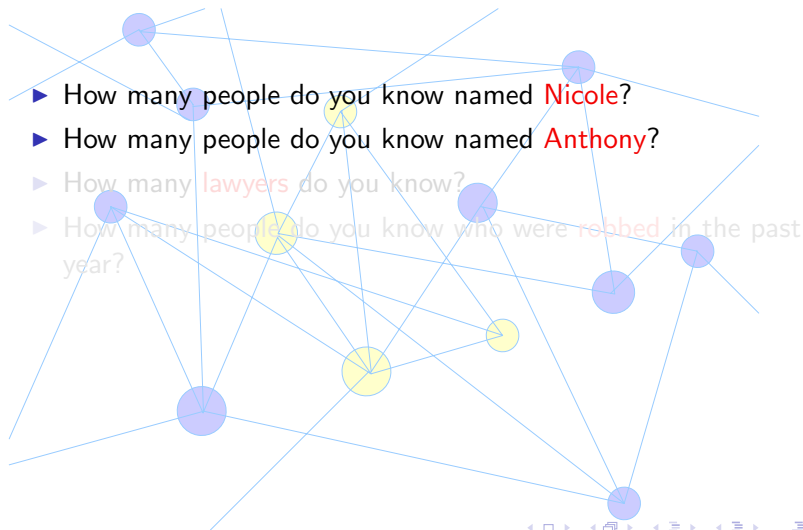
How many people do you know? Demonstration



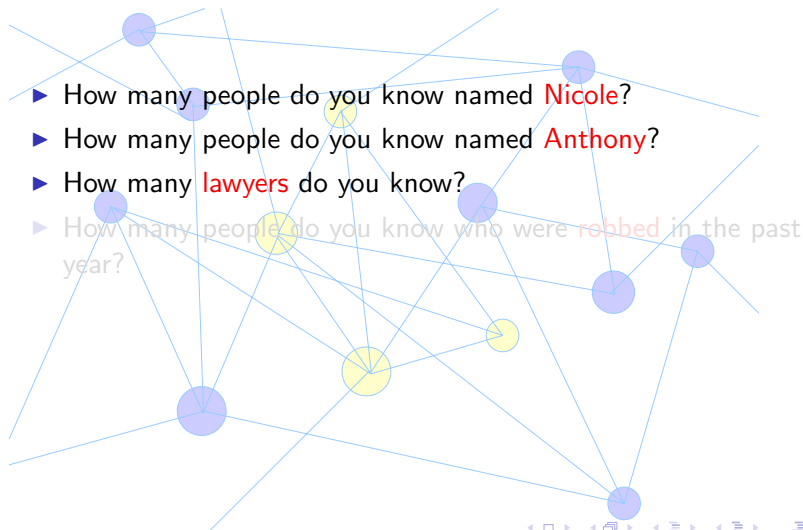
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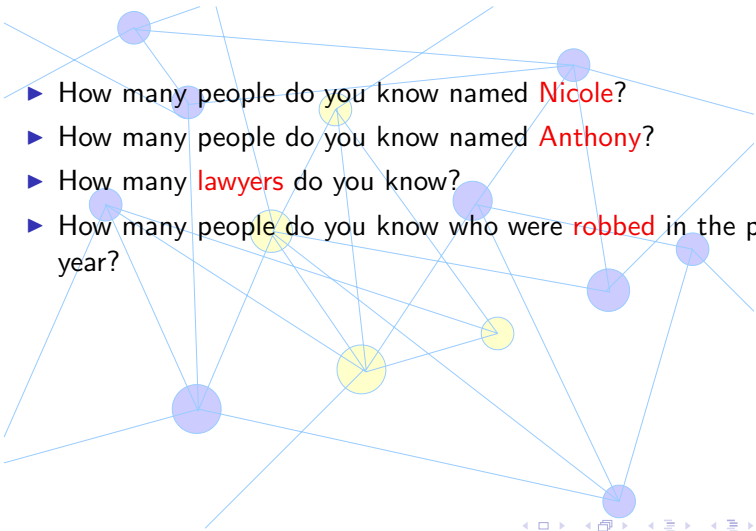
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How many people do you know? Demonstration

- 
- ▶ How many people do you know named **Nicole**?
 - ▶ How many people do you know named **Anthony**?
 - ▶ How many **lawyers** do you know?
 - ▶ How many people do you know who were **robbed** in the past year?

Scale-up method: demonstration

- ▶ On average, you knew 0.6 Nicoles
- ▶ 0.13% of Americans are named Nicole
- ▶ Assume 0.13% of **your** acquaintances are Nicoles
- ▶ Estimate: on average, you know $0.6/0.0013 = 450$ people
- ▶ On average, you know 0.8 Anthonys
- ▶ 0.31% of Americans are named Anthony
- ▶ Estimate: on average, you know $1.6/0.0031 = 260$ people
- ▶ Why do these differ?

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Estimating group sizes: demonstration

- ▶ On average, you know 2.6 lawyers
 - ▶ Assume average network size is 450 people
 - ▶ Estimate: lawyers represent $2.6/450 = 0.58\%$ of the network
 - ▶ Estimate: $0.0058 \cdot 290$ million = 1.7 million lawyers in the U.S.
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- ▶ On average, you know 0.25 people who were robbed last year
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Killworth, McCarty et al. surveys

- ▶ How many X's do you know?
- ▶ Stephanie, Jacqueline, Kimberly, Nicole, Christina, Jennifer
- ▶ Christopher, David, Anthony, Robert, James, Michael
- ▶ Twin, woman adopted kid in past year, gave birth in past year, widow(er) under 65
- ▶ Commercial pilot, gun dealer, postal worker, member of Jaycees, opened business in past year, American Indian
- ▶ Suicide in past year, died in auto accident, diabetic, kidney dialysis, AIDS, HIV-positive, rape victim, homicide victim, male in prison, homeless

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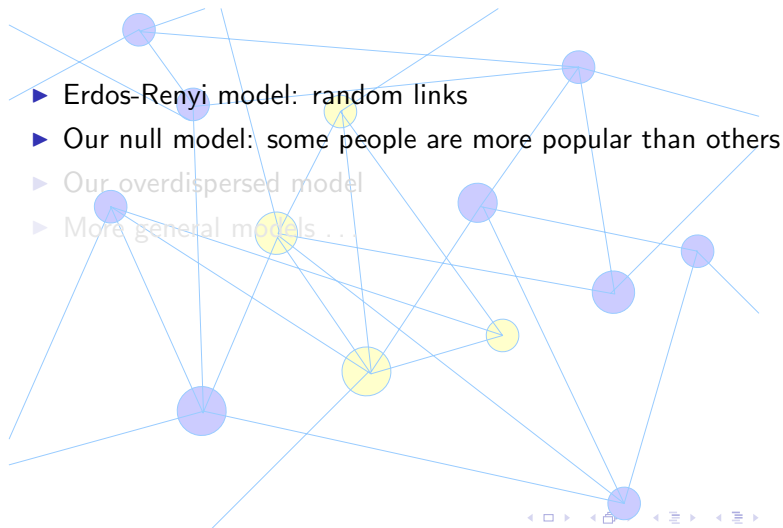
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Models of social network data

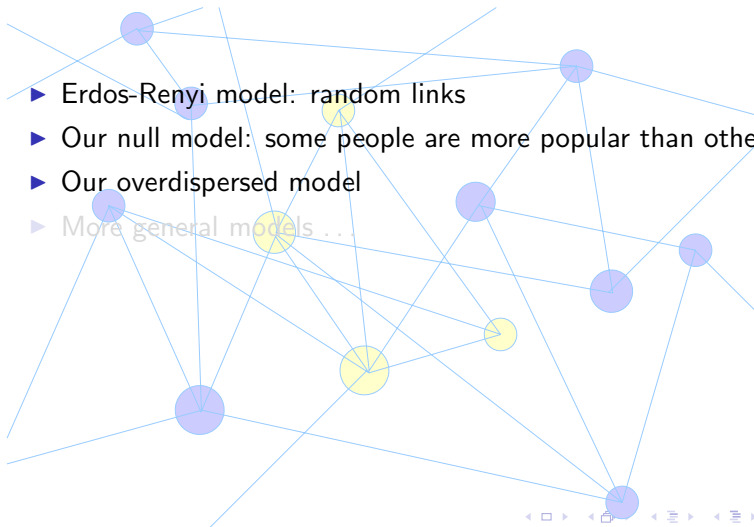
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 - ▶ Our null model: some people are more popular than others
 - ▶ Our overdispersed model
 - ▶ More general models ...

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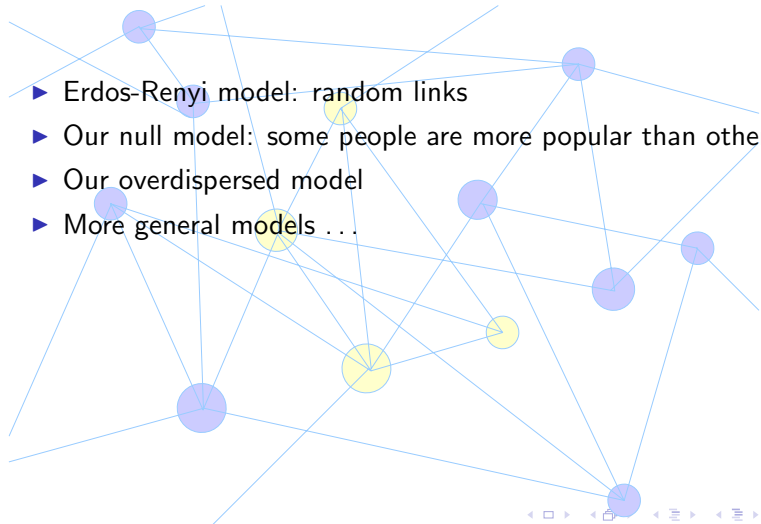


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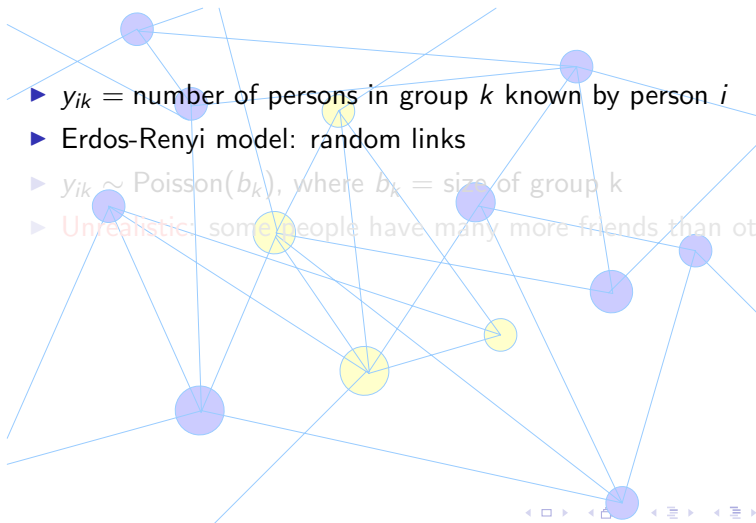


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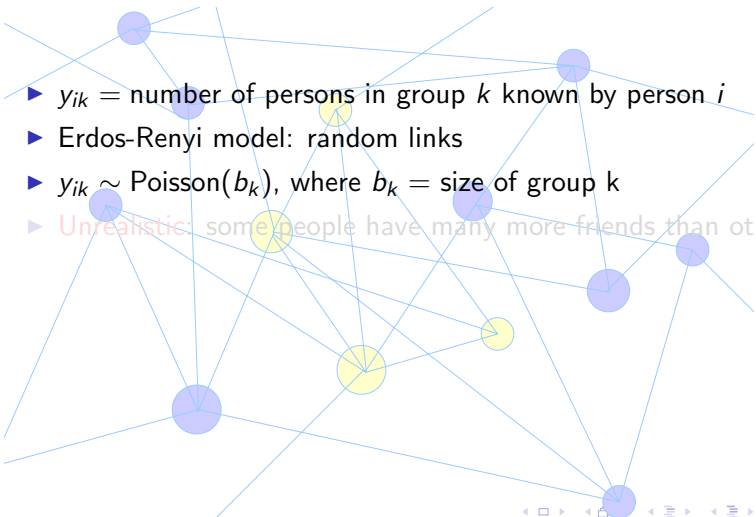
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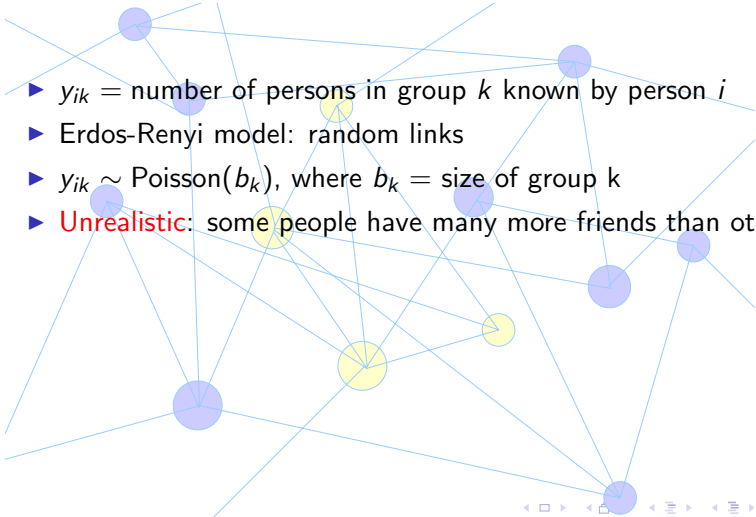
- ▶ y_{ik} = number of persons in group k known by person i
- ▶ Erdos-Renyi model: random links
- ▶ $y_{ik} \sim \text{Poisson}(b_k)$, where b_k = size of group k
- ▶ **Unrealistic:** some people have many more friends than others



Erdos-Renyi model

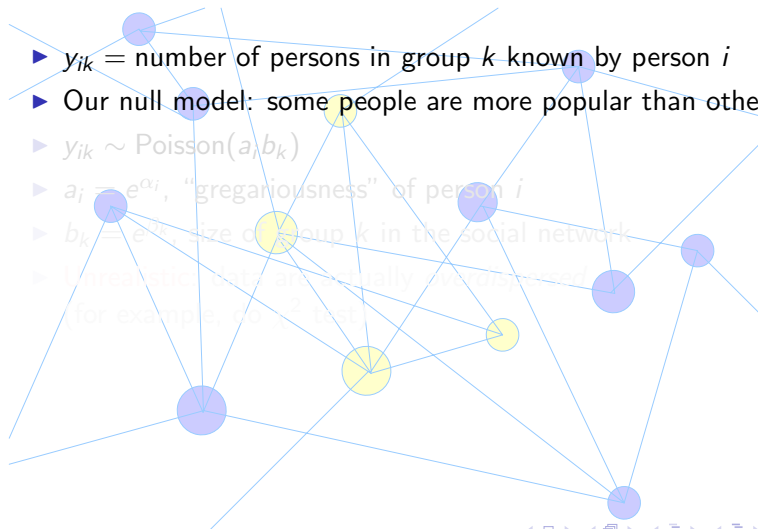
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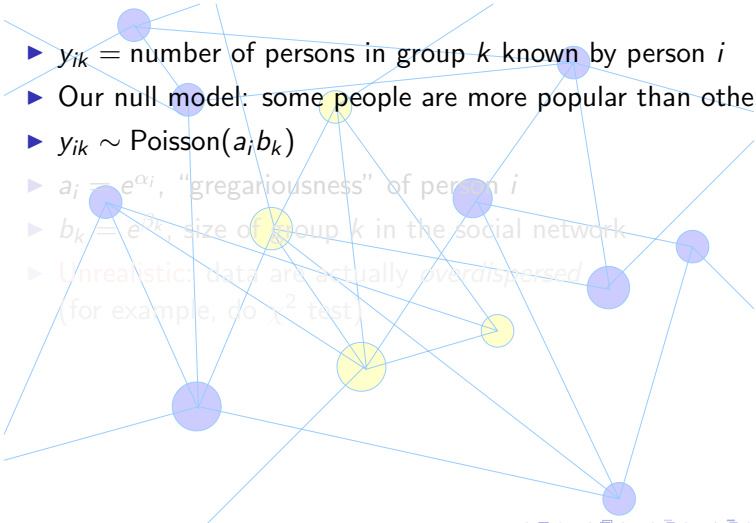
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Our null model

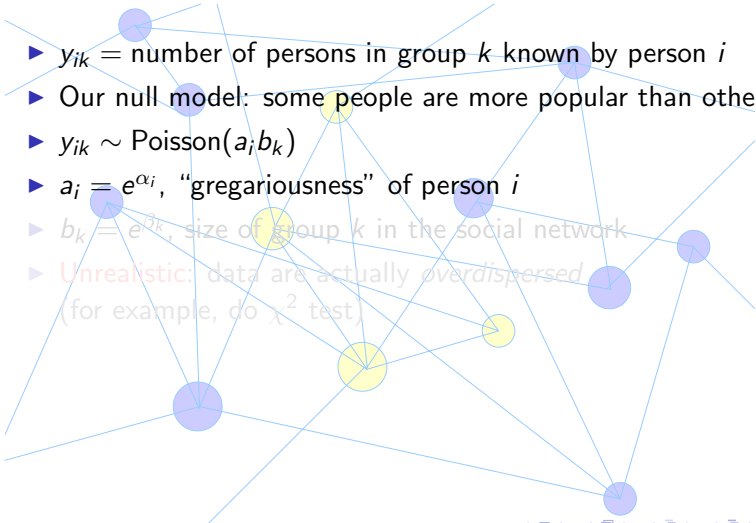
- ▶ y_{ik} = number of persons in group k known by person i
- ▶ Our null model: some people are more popular than others
- ▶ $y_{ik} \sim \text{Poisson}(a_i b_k)$
- ▶ $a_i = e^{\alpha_i}$, "gregariousness" of person i
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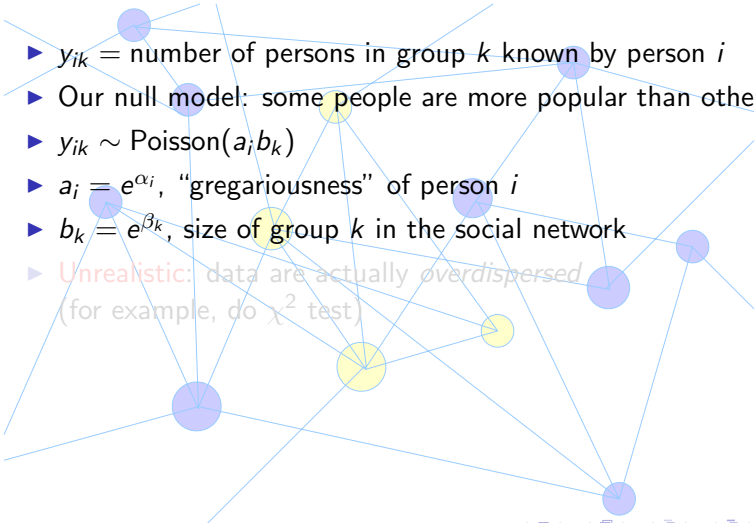
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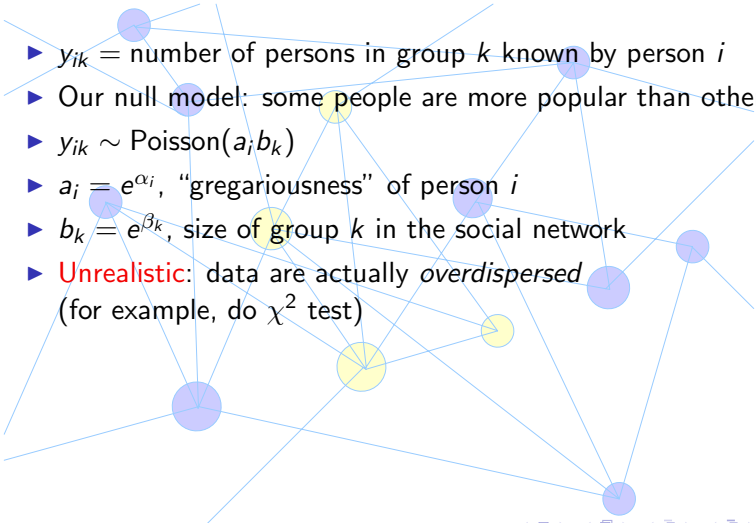
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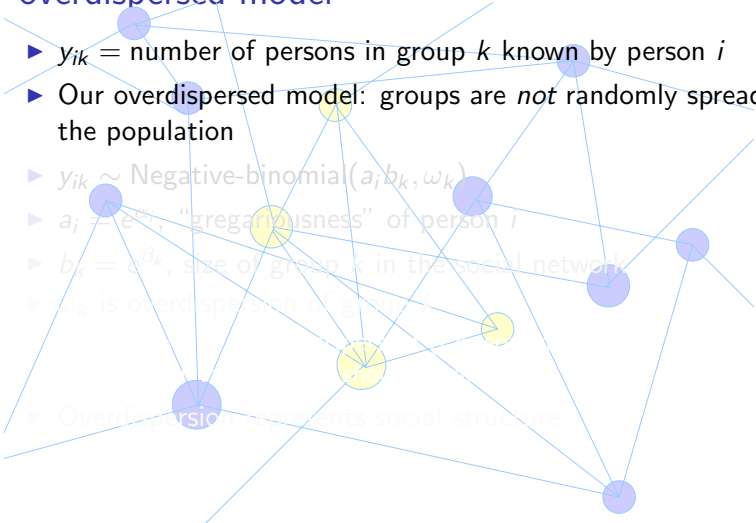
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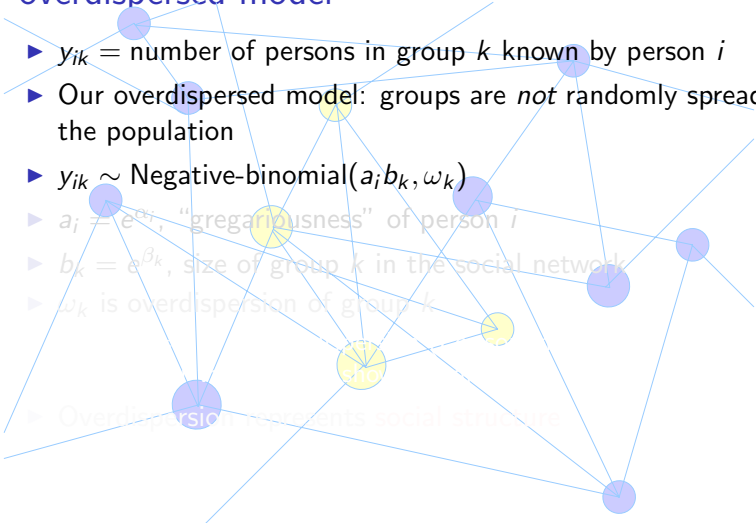
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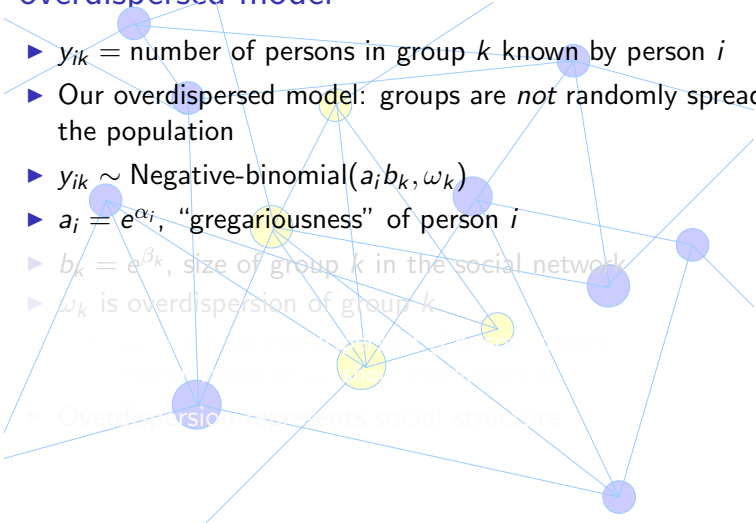
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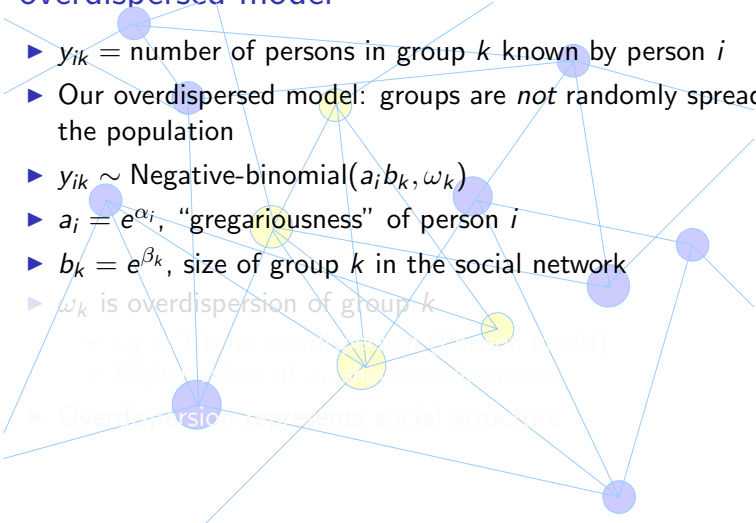
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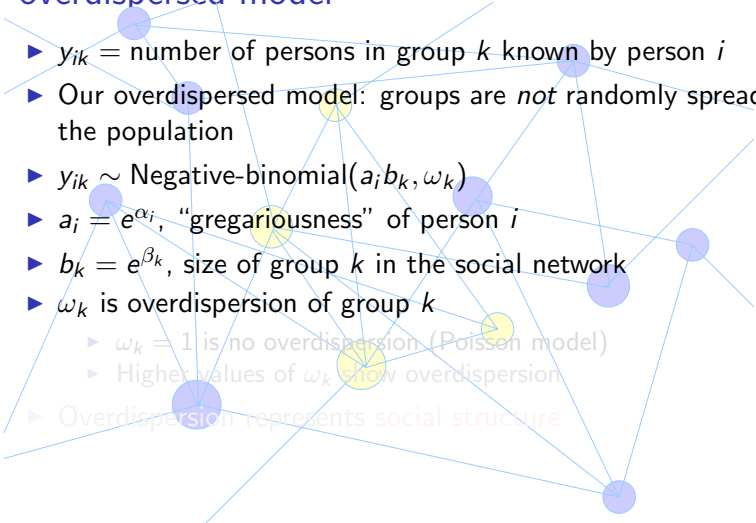
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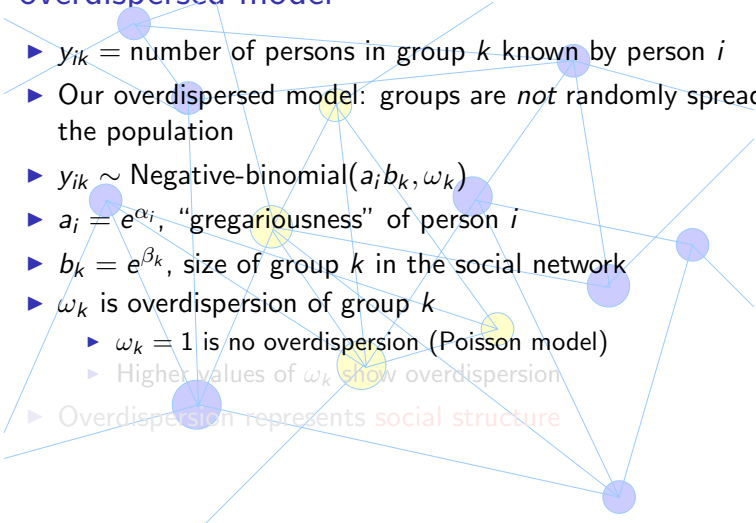
▶ Higher values of ω_k show overdispersion

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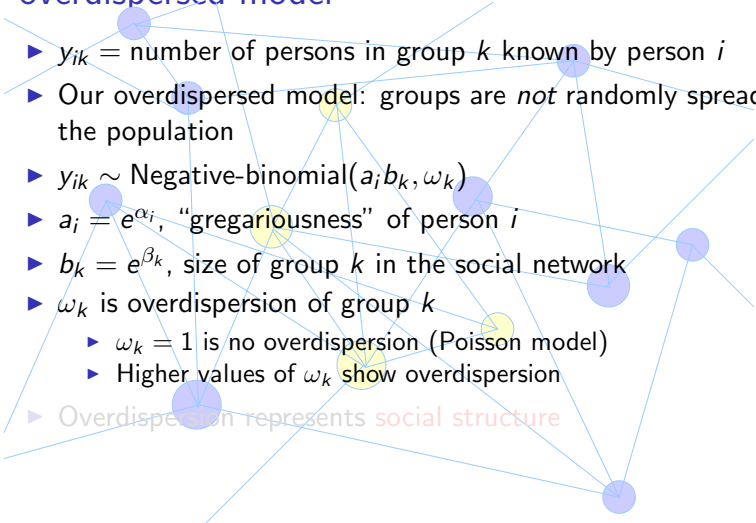
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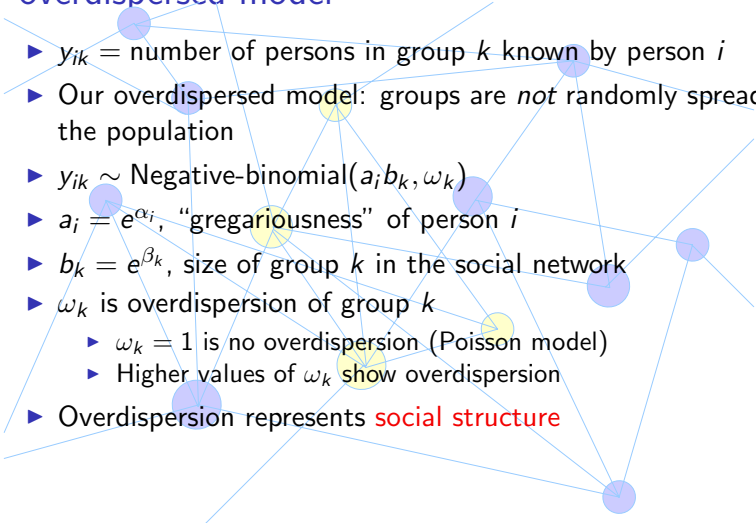
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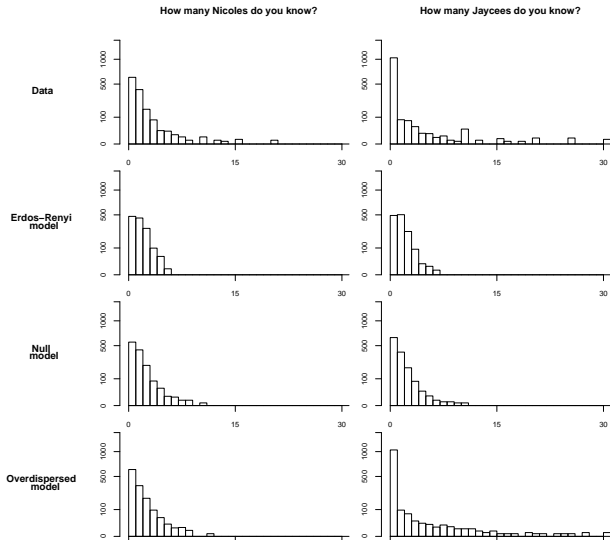
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Data, compared to simulations from 3 models



Bayesian inference

- ▶ Negative-binomial data model allowing overdispersion
- ▶ Hierarchical models for gregariousness, group-size, and overdispersion parameters
- ▶ $1370 + 32 + 32 + 4$ parameters to estimate
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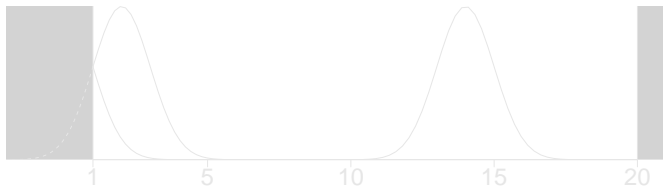
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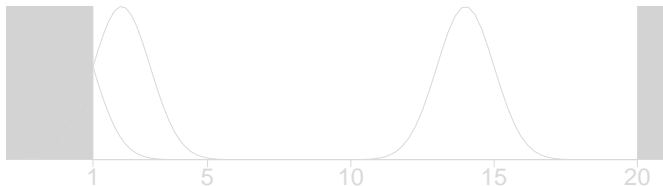
Gibbs-Metropolis algorithm: updating α, β, ω

- ▶ For each i , update α_i using Metropolis with jumping dist.
 $\alpha_i^* \sim N(\alpha_i^{(t-1)}, (\text{jumping scale of } \alpha_i)^2)$.
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- Reflect jumps off the edges:



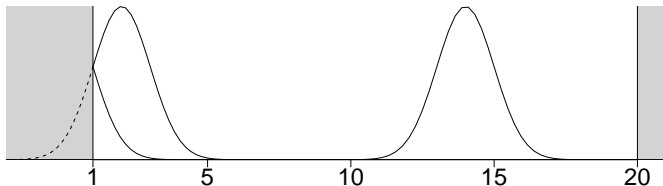
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Gibbs-Metropolis algorithm: updating hyperparameters

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Renormalizing the α_i 's and β_k 's

- ▶ Problem: α_i 's and β_k 's are not separately identified in the model, $y_{ik} \sim \text{Negative-binomial}(e^{\alpha_i + \beta_k}, \omega_k)$
- ▶ Possible solutions:
 - ▶ Choose a “baseline” value: set $\alpha_1 = 0$ (for example)
 - ▶ Renormalize a group of parameters: set $\sum_{i=1}^n \alpha_i = 0$
 - ▶ Apply a group of constraints: e.g., $\alpha_i = \beta_i$
- ▶ Our solution: rescale so that the b_k 's for the names (Nicole, Anthony, etc.) equal their proportion in the population:

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 - ▶ Subtract C from all the β_k ’s and μ_β

Renormalizing the α_i 's and β_k 's

- ▶ Problem: α_i 's and β_k 's are not separately identified in the model, $y_{ik} \sim \text{Negative-binomial}(e^{\alpha_i + \beta_k}, \omega_k)$
- ▶ Possible solutions:
 - ▶ Choose a "baseline" value: set $\alpha_1 = 0$ (for example)
 - ▶ Renormalize a group of parameters: set $\sum_{i=1}^n \alpha_i = 0$
 - ▶ Anchor the prior distribution: set $\mu_\alpha = 0$
- ▶ Our solution: rescale so that the b_k 's for the names (Nicole, Anthony, etc.) equal their proportion in the population:
 - ▶ Compute $C = \log \left(\sum_{k=1}^{12} e^{\beta_k} / 0.069 \right)$
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Adaptive Metropolis jumping

- ▶ Parallel scalar updating of the components of α, β, ω
- ▶ Adapt each of $1370 + 32 + 32$ jumping scales to have $E(p_{\text{jump}}) \approx 0.44$
- ▶ Save p_{jump} from each Metropolis step, then average them and rescale every 50 iterations:
 - ▶ Where avg $p_{\text{jump}} > 0.44$, **increase** the jump scale
 - ▶ Where avg $p_{\text{jump}} < 0.44$, **decrease** the jump scale
- ▶ After burn-in, stop adapting
- ▶ If we had vector jumps, we would adapt the scale so that $E(p_{\text{jump}}) \approx 0.23$
- ▶ More effective adaptation uses avg. squared jumped distance

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Computation in R

- ▶ BUGS was too slow (over 1400 parameters)
- ▶ Programming from scratch in R is awkward, buggy
- ▶ Instead, we use our general Gibbs/Metropolis programming environment
- ▶ Set up MCMC object
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Setting up the MCMC object

```
network.1 <- mcmcEngine (network.data, network.init,  
  update=network.update, n.iter=1000, n.chains=3)  
network.update <- list(  
  alpha = Metropolis (f.logpost.alpha),  
  beta = Metropolis (f.logpost.beta),  
  omega = Metropolis (f.logpost.omega,  
    jump=Jump("omega.jump", lower=1.01, upper=20)),  
  mu.alpha = Gibbs (mu.alpha.update),  
  mu.beta = Gibbs (mu.beta.update),  
  sigma.alpha = Gibbs (sigma.alpha.update),  
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  renorm.network)
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Data and initial values

```
y <- as.matrix (read.dta ("social.dta"))  
y <- y[1:50,]  
network.data <- list (y=y, data.n=nrow(y),  
  data.j=ncol(y))  
network.init <- function(){  
  alpha <- rnorm(data.n)  
  beta <- rnorm(data.j)  
  omega <- runif(data.j,1.01,20)  
  mu.alpha <- rnorm(1)  
  mu.beta <- rnorm(1)  
  sigma.alpha <- runif(1)  
  sigma.beta <- runif(1)}
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```

Gibbs samplers for the hyperparameters

```
mu.alpha.update <- function()  
  rnorm (1, mean(alpha), sigma.alpha/sqrt(data.n))  
mu.beta.update <- function()  
  rnorm (1, mean(beta), sigma.beta/sqrt(data.j))  
sigma.alpha.update <- function()  
  sqrt (sum((alpha-mu.alpha)^2)/rchisq(1, data.n-1))  
sigma.beta.update <- function()  
  sqrt (sum((beta-mu.beta)^2)/rchisq(1, data.j-1))
```

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sigma.beta.update <- function()  
  sqrt (sum((beta-mu.beta)^2)/rchisq(1, data.j-1))
```

Log-likelihood for each data point

```
f.loglik <- function (y, alpha, beta, omega, data.n){  
  theta.mat <- exp(outer(alpha, beta, "+"))  
  omega.mat <- outer(rep(0, data.n), omega, "+")  
  dnbinom (y, theta.mat/(omega.mat-1), 1/omega.mat,  
    log=T)}
```


Log-posterior density for each vector parameter

```
f.logpost.alpha <- function() {  
  loglik <- f.loglik (y, alpha, beta, omega, data.n)  
  rowSums (loglik, na.rm=TRUE) +  
    dnorm (alpha, mu.alpha, sigma.alpha, log=TRUE)}  
f.logpost.beta <- function() {  
  loglik <- f.loglik (y, alpha, beta, omega, data.n)  
  colSums (loglik, na.rm=TRUE) +  
    dnorm (beta, mu.beta, sigma.beta, log=TRUE)}  
f.logpost.omega <- function() {  
  loglik <- f.loglik (y, alpha, beta, omega, data.n)  
  colSums (loglik, na.rm=T)}
```

Log-posterior density for each vector parameter

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f.logpost.alpha <- function() {  
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f.logpost.omega <- function() {  
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  colSums (loglik, na.rm=TRUE) +  
    dnorm (beta, mu.beta, sigma.beta, log=TRUE)}  
f.logpost.omega <- function() {  
  loglik <- f.loglik (y, alpha, beta, omega, data.n)  
  colSums (loglik, na.rm=T)}
```

Bounded jumping for the ω_k 's

Customized Metropolis jumping rule for the components of ω :

```
omega.jump <- function (omega, sigma) {  
  reflect (rnorm (length(omega), omega, sigma),  
    .lower, .upper)}
```

Renormalization of the α_i 's and β_k 's

```
renorm.network <- function() {  
  const <- log (sum(exp(beta[1:12]))/0.069)  
  alpha <- alpha + const  
  mu.alpha <- mu.alpha + const  
  beta <- beta - const  
  mu.beta <- mu.beta - const}
```

Running MCMC and looking at the output

```
net <- run(network.1)
attach (as.rv (net))
```

Some output:

name	mean	sd	25%	50%	75%	Rhat
beta[1]	-5.1	0.1	(-5.4	-5.2	-5.1)	1.0
beta[2]	-6.4	0.1	(-6.9	-6.7	-6.5)	1.2
beta[3]	-6.1	0.1	(-6.5	-6.3	-6.2)	1.1
beta[4]	-7.0	0.2	(-7.6	-7.4	-7.1)	1.0
beta[5]	-5.1	0.1	(-5.4	-5.3	-5.2)	1.2
beta[6]	-5.6	0.2	(-6.1	-5.9	-5.8)	1.0

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Running MCMC and looking at the output

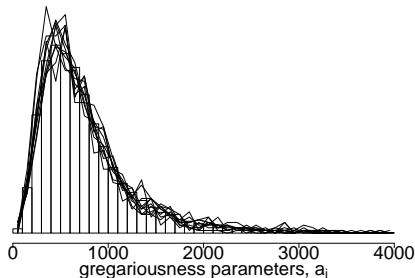
```
net <- run(network.1)
attach (as.rv (net))
```

Some output:

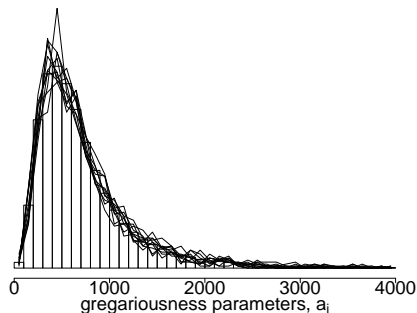
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Estimated distributions of network sizes for men and women

men



women



Regression of $\log(\text{gregariousness})$

Coefficient

Estimate

female

nonwhite

age < 30

age > 65

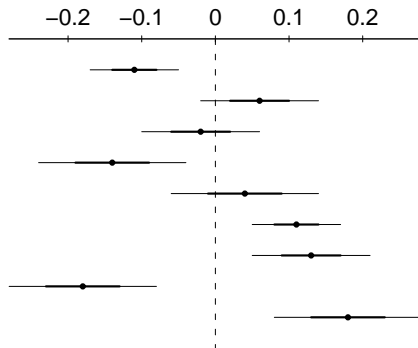
married

college educated

employed

income < \$20,000

income > \$80,000



Parameter estimates for the 32 subpopulations

► Subpopulations

- Names (Stephanie, Michael, etc.)
- Other groups (pilots, diabetics, etc.)

► Parameters

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- ▶ Subpopulations
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- Other groups (pilots, diabetics, etc.)

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- Proportion of the social network, e^{β_k}
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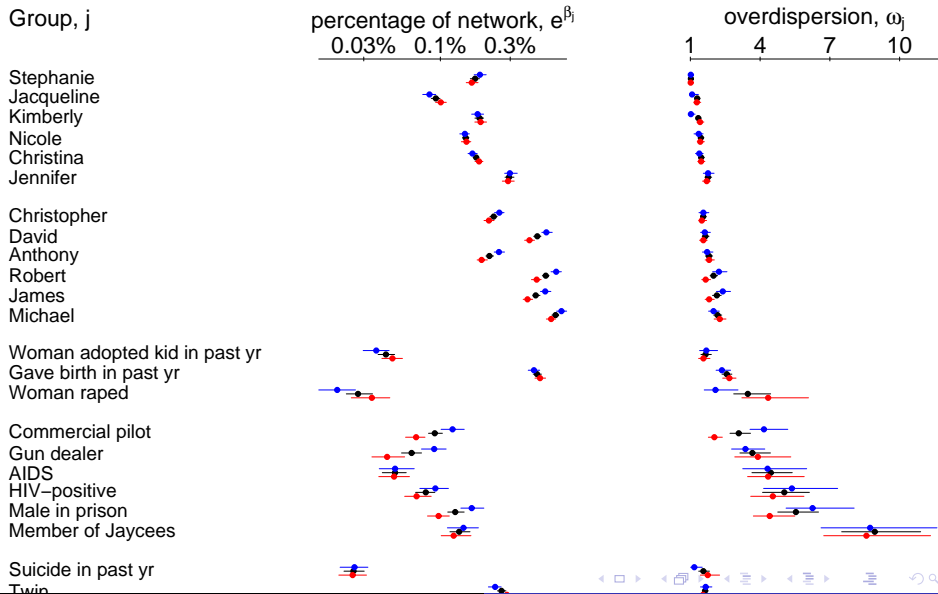
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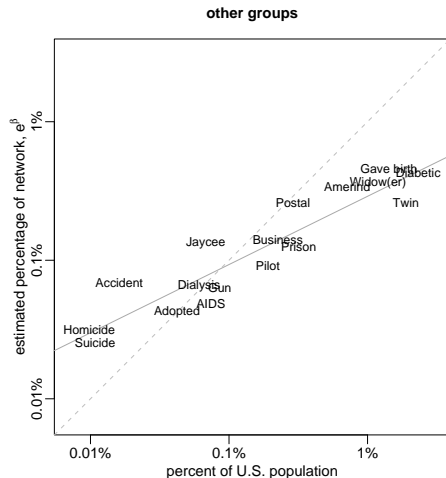
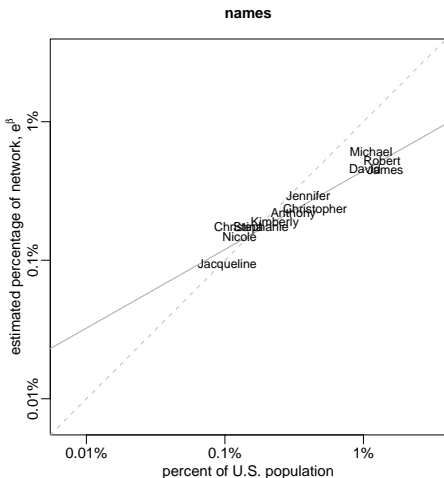
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Common groups (e.g., “I know a lot of people who work at Google”) are under-represented

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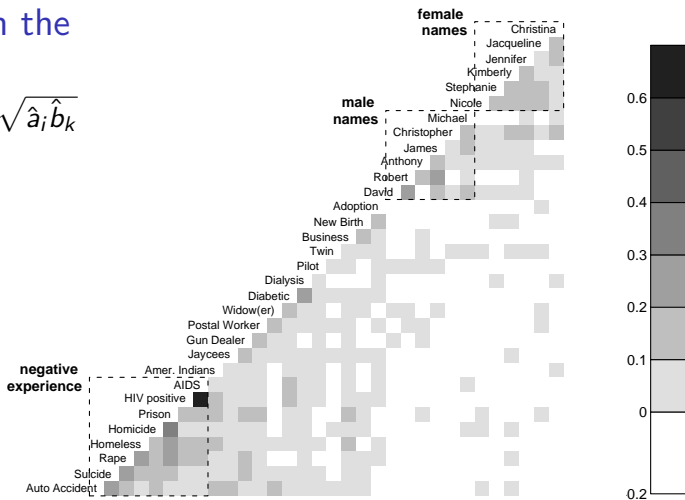
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Correlations in the residuals

$$r_{ik} = \sqrt{y_{ik}} - \sqrt{\hat{a}_i \hat{b}_k}$$



Confidence building

- ▶ Posterior predictive checking: compare data to simulated replications from the model
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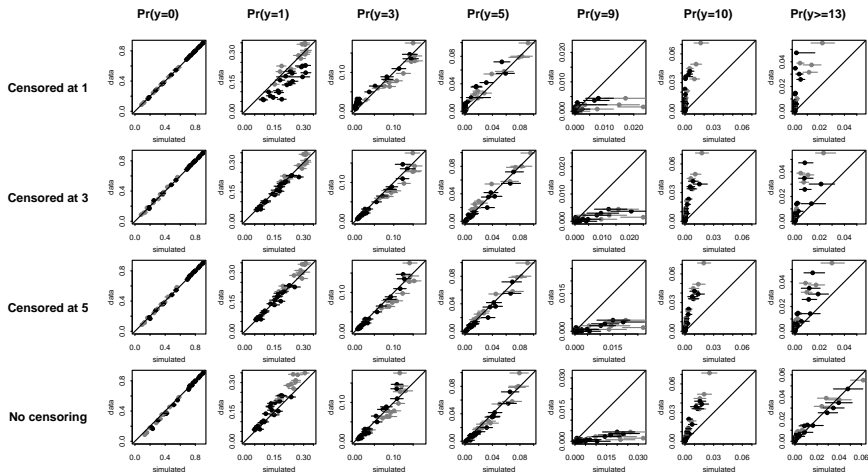
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Actual vs. simulated proportions of $y = 0, 1, \dots$



Do you know 0, 1, 2, or 3 or more Nicoles?

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- ▶ $y_{ik} = 0, 1, 2, \text{ or } \geq 3$

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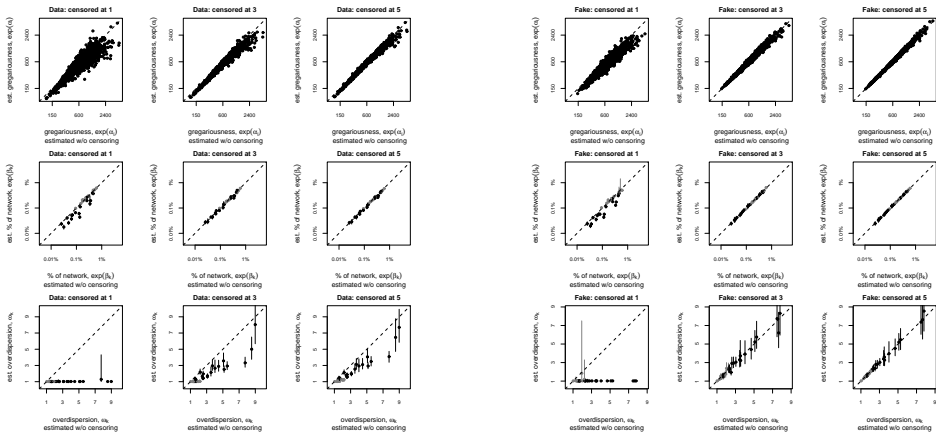
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Evaluation of inferences using fake data



Running the demo

- ▶ How many Nicoles, Anthonys, lawyers, people robbed?
- ▶ Real-time data analysis
 - ▶ Entering in the data: 20 minutes
 - ▶ Running the program: 500 iterations (40 seconds), 1000 iterations (80 seconds)
 - ▶ Real-time debugging: 5 minutes
 - ▶ Altering the presentation: 15 minutes
- ▶ Results for social network sizes, α
- ▶ Results for group sizes, β
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Regression of $\log(\text{gregariousness})$: as a table

Coefficient	Estimate (s.e.)
female	-0.11 (0.03)
nonwhite	0.06 (0.04)
age < 30	-0.02 (0.04)
age > 65	-0.14 (0.05)
married	0.04 (0.05)
college educated	0.11 (0.03)
employed	0.13 (0.04)
income < \$20,000	-0.18 (0.05)
income > \$80,000	0.18 (0.05)

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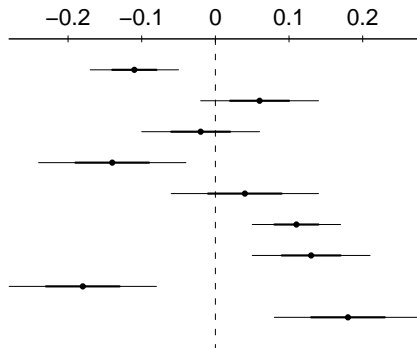
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